



Analyse de données non stationnaires : représentations, théorie, algorithmes et applications.

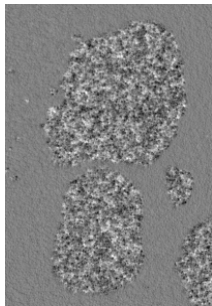
**Barbara Pascal**

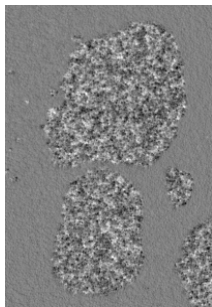
*March 7<sup>th</sup> 2022*

Laboratoire Mathématiques Appliquées À Paris 5 (MAP5)

Groupe de travail Images

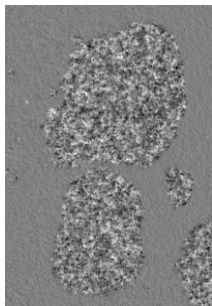
## **Part I:** Texture segmentation based on fractal attributes





**Goal:** obtain a partition of the image into  $K$  homogeneous textures

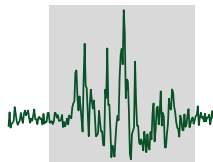
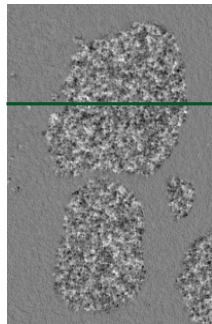
$$\Omega = \Omega_1 \sqcup \dots \sqcup \Omega_K$$



**Goal:** obtain a partition of the image into  $K$  **homogeneous textures**

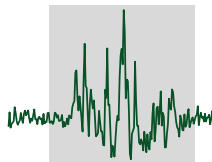
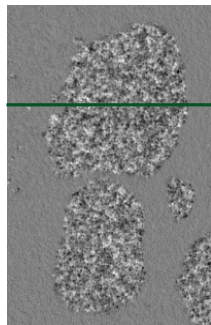
$$\Omega = \Omega_1 \sqcup \dots \sqcup \Omega_K$$

# Piecewise monofractal model



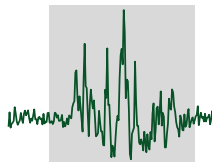
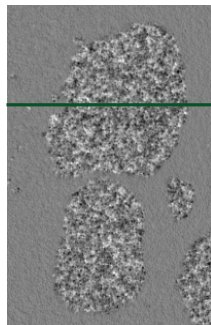
## Fractals attributes

- variance  $\sigma^2$      *amplitude of variations*



## Fractals attributes

- variance  $\sigma^2$       *amplitude of variations*
- local regularity  $h$       *scale invariance*

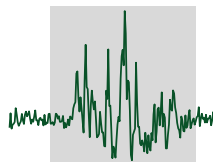
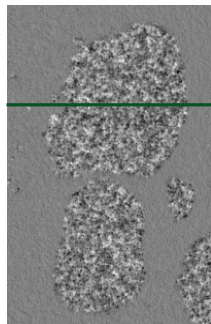




## Fractals attributes

- variance  $\sigma^2$       *amplitude of variations*
- local regularity  $h$       *scale invariance*

$$|f(x) - f(y)| \leq \sigma(x)|x - y|^{h(x)}$$



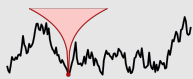
## Fractals attributes

- variance  $\sigma^2$       *amplitude of variations*
- local regularity  $h$       *scale invariance*

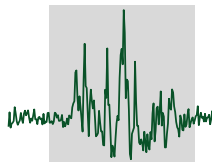
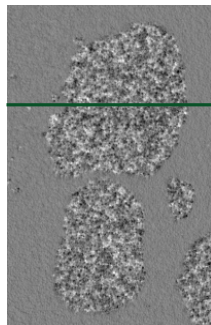
$$|f(x) - f(y)| \leq \sigma(x)|x - y|^{h(x)}$$



$$h(x) \equiv h_1 = 0.9$$



$$h(x) \equiv h_2 = 0.3$$



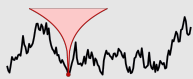
## Fractals attributes

- variance  $\sigma^2$       *amplitude of variations*
- local regularity  $h$       *scale invariance*

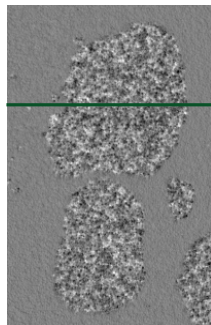
$$|f(x) - f(y)| \leq \sigma(x)|x - y|^{h(x)}$$



$$h(x) \equiv h_1 = 0.9$$

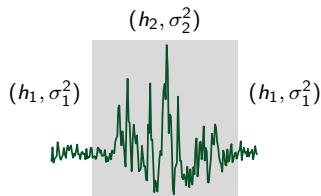


$$h(x) \equiv h_2 = 0.3$$



## Segmentation

- ▶  $\sigma^2$  and  $h$  piecewise constant



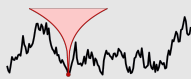
## Fractals attributes

- variance  $\sigma^2$       *amplitude of variations*
- local regularity  $h$       *scale invariance*

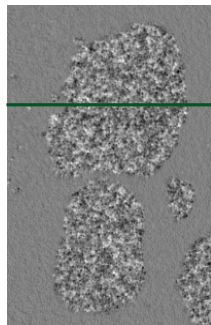
$$|f(x) - f(y)| \leq \sigma(x)|x - y|^{h(x)}$$



$$h(x) \equiv h_1 = 0.9$$

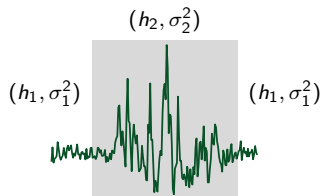


$$h(x) \equiv h_2 = 0.3$$

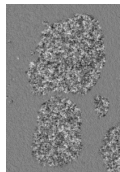


## Segmentation

- ▶  $\sigma^2$  and  $h$  piecewise constant
- ▶ region  $\Omega_k$  characterized by  $(\sigma_k^2, h_k)$



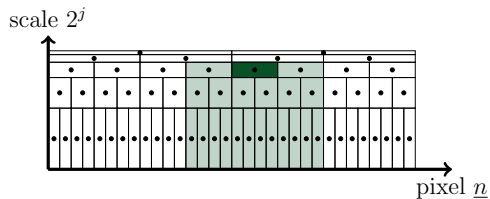
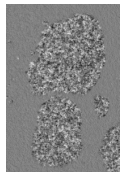
Textured image



# Multiscale analysis

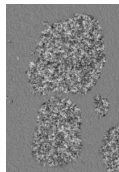
Textured image

Local maximum of wavelet coefficients:  $\mathcal{L}_a$ .



# Multiscale analysis

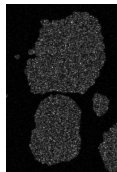
Textured image



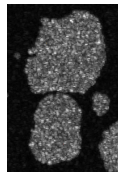
Local maximum of wavelet coefficients:  $\mathcal{L}_a$ .

Scale

$a = 2^1$

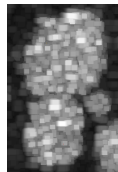


$a = 2^2$

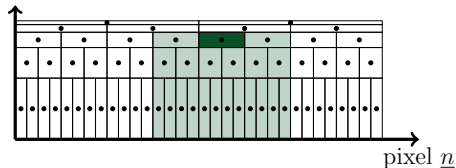


...

$a = 2^5$

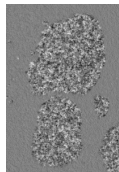


scale  $2^j$



# Multiscale analysis

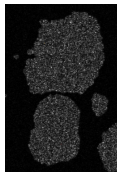
Textured image



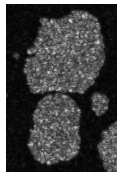
Scale

Local maximum of wavelet coefficients:  $\mathcal{L}_{a,\cdot}$

$a = 2^1$

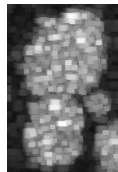


$a = 2^2$



...

$a = 2^5$



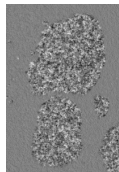
Proposition (Jaffard, 2004), (Wendt, 2008)

$$\log(\mathcal{L}_{a,\cdot}) \underset{a \rightarrow 0}{\simeq} \log(a) \underset{\text{regularity}}{\mathbf{h}} + \underset{\substack{\propto \log(\sigma^2) \\ \text{(variance)}}}{\mathbf{v}}$$



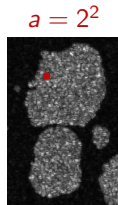
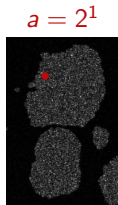
# Multiscale analysis

Textured image

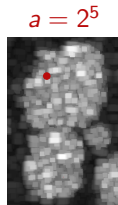


Scale

Local maximum of wavelet coefficients:  $\mathcal{L}_{a,\cdot}$ .

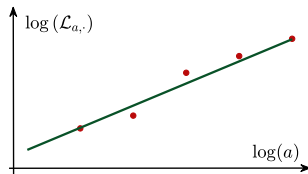


...



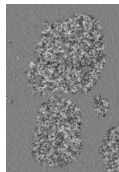
Proposition (Jaffard, 2004), (Wendt, 2008)

$$\log(\mathcal{L}_{a,\cdot}) \underset{a \rightarrow 0}{\simeq} \log(a) \underset{\text{regularity}}{\mathbf{h}} + \underset{\propto \log(\sigma^2) \text{ (variance)}}{\mathbf{v}}$$



# Multiscale analysis

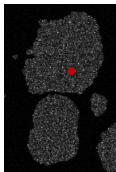
Textured image



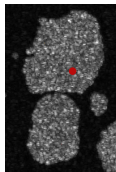
Scale

Local maximum of wavelet coefficients:  $\mathcal{L}_{a,\cdot}$

$a = 2^1$

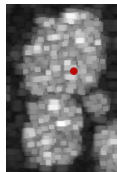


$a = 2^2$



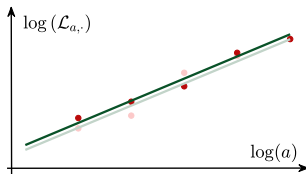
...

$a = 2^5$



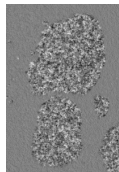
Proposition (Jaffard, 2004), (Wendt, 2008)

$$\log(\mathcal{L}_{a,\cdot}) \underset{a \rightarrow 0}{\simeq} \log(a) \underset{\text{regularity}}{\mathbf{h}} + \underset{\propto \log(\sigma^2) \text{ (variance)}}{\mathbf{v}}$$



# Multiscale analysis

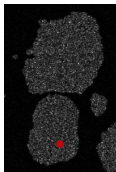
Textured image



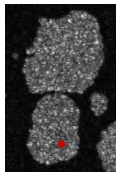
Scale

Local maximum of wavelet coefficients:  $\mathcal{L}_{a,\cdot}$

$a = 2^1$

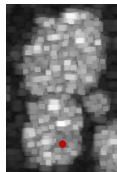


$a = 2^2$



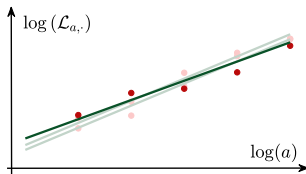
...

$a = 2^5$



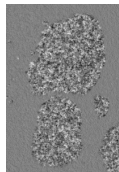
Proposition (Jaffard, 2004), (Wendt, 2008)

$$\log(\mathcal{L}_{a,\cdot}) \underset{a \rightarrow 0}{\simeq} \log(a) \underset{\text{regularity}}{\mathbf{h}} + \underset{\propto \log(\sigma^2) \text{ (variance)}}{\mathbf{v}}$$



# Multiscale analysis

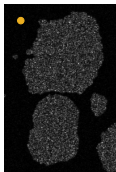
Textured image



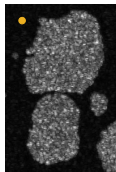
Scale

Local maximum of wavelet coefficients:  $\mathcal{L}_{a,\cdot}$

$a = 2^1$

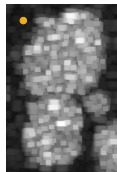


$a = 2^2$



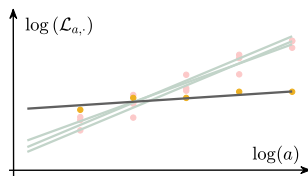
...

$a = 2^5$



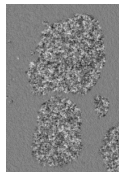
Proposition (Jaffard, 2004), (Wendt, 2008)

$$\log(\mathcal{L}_{a,\cdot}) \underset{a \rightarrow 0}{\simeq} \log(a) \underset{\text{regularity}}{\mathbf{h}} + \underset{\propto \log(\sigma^2) \text{ (variance)}}{\mathbf{v}}$$



# Multiscale analysis

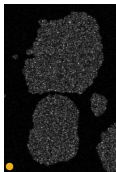
Textured image



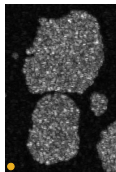
Scale

Local maximum of wavelet coefficients:  $\mathcal{L}_{a,\cdot}$

$a = 2^1$

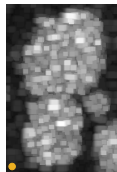


$a = 2^2$



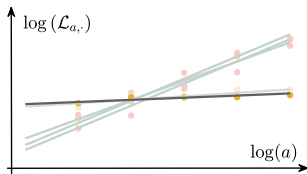
...

$a = 2^5$



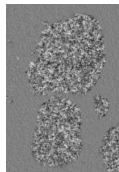
Proposition (Jaffard, 2004), (Wendt, 2008)

$$\log(\mathcal{L}_{a,\cdot}) \underset{a \rightarrow 0}{\simeq} \log(a) \underset{\text{regularity}}{\mathbf{h}} + \underset{\propto \log(\sigma^2) \text{ (variance)}}{\mathbf{v}}$$



# Multiscale analysis

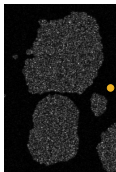
Textured image



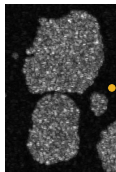
Scale

Local maximum of wavelet coefficients:  $\mathcal{L}_{a,\cdot}$

$a = 2^1$

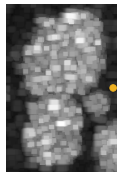


$a = 2^2$



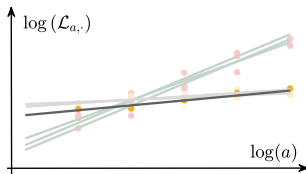
...

$a = 2^5$



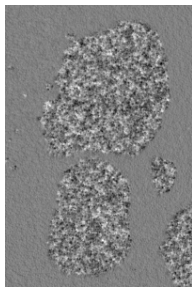
Proposition (Jaffard, 2004), (Wendt, 2008)

$$\log(\mathcal{L}_{a,\cdot}) \underset{a \rightarrow 0}{\simeq} \log(a) \underset{\text{regularity}}{\mathbf{h}} + \underset{\propto \log(\sigma^2) \text{ (variance)}}{\mathbf{v}}$$



**Linear regression**      $\log(\mathcal{L}_{a,\cdot}) \simeq \log(a) \underset{\text{regularity}}{\mathbf{h}} + \underset{\propto \log(\sigma^2)}{\mathbf{v}}$

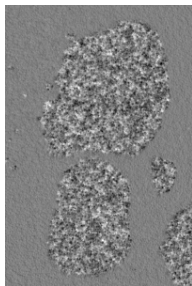
Textured image



**Linear regression**  $\log(\mathcal{L}_{a,\cdot}) \simeq \log(a) \underset{\text{regularity}}{\mathbf{h}} + \underset{\propto \log(\sigma^2)}{\mathbf{v}}$

$$\left(\hat{\mathbf{h}}^{\text{LR}}, \hat{\mathbf{v}}^{\text{LR}}\right) = \underset{\mathbf{h}, \mathbf{v}}{\operatorname{argmin}} \sum_{a=a_{\min}}^{a_{\max}} \|\log(\mathcal{L}_{a,\cdot}) - \log(a)\mathbf{h} - \mathbf{v}\|^2$$

Textured image

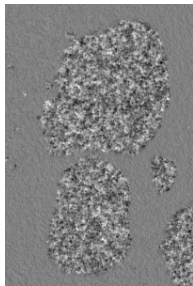




**Linear regression**  $\log(\mathcal{L}_{a,\cdot}) \simeq \log(a) \underset{\text{regularity}}{\mathbf{h}} + \underset{\propto \log(\sigma^2)}{\mathbf{v}}$

$$\left(\widehat{\mathbf{h}}^{\text{LR}}, \widehat{\mathbf{v}}^{\text{LR}}\right) = \underset{\mathbf{h}, \mathbf{v}}{\operatorname{argmin}} \sum_{a=a_{\min}}^{a_{\max}} \|\log(\mathcal{L}_{a,\cdot}) - \log(a)\mathbf{h} - \mathbf{v}\|^2$$

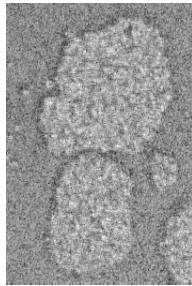
Textured image



Local regularity  $\widehat{\mathbf{h}}^{\text{LR}}$



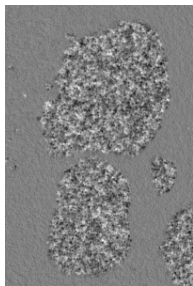
Local power  $\widehat{\mathbf{v}}^{\text{LR}}$



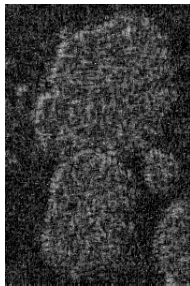
**Linear regression**  $\frac{\mathbb{E} \log(\mathcal{L}_{a,\cdot})}{\text{expected value}} = \log(a) \underset{\text{regularity}}{\bar{\mathbf{h}}} + \underset{\propto \log(\sigma^2)}{\bar{\mathbf{v}}}$

$$\left( \hat{\mathbf{h}}^{\text{LR}}, \hat{\mathbf{v}}^{\text{LR}} \right) = \underset{\mathbf{h}, \mathbf{v}}{\operatorname{argmin}} \sum_{a=a_{\min}}^{a_{\max}} \|\log(\mathcal{L}_{a,\cdot}) - \log(a)\mathbf{h} - \mathbf{v}\|^2$$

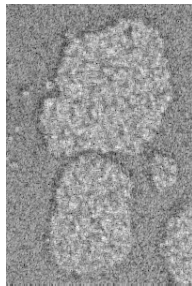
Textured image



Local regularity  $\hat{\mathbf{h}}^{\text{LR}}$



Local power  $\hat{\mathbf{v}}^{\text{LR}}$



→ large estimation variance

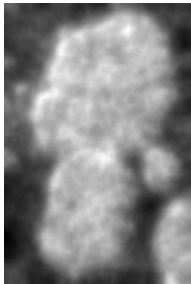
**Filter smoothing** (linear)

$$\left(\mathbf{I} + \lambda \mathbf{D}^\top \mathbf{D}\right)^{-1} \hat{\mathbf{h}}^{\text{LR}}$$

Linear regression  $\hat{\mathbf{h}}^{\text{LR}}$



Lissage



**Filter smoothing** (linear)

$$\left(\mathbf{I} + \lambda \mathbf{D}^\top \mathbf{D}\right)^{-1} \hat{\mathbf{h}}^{\text{LR}}$$

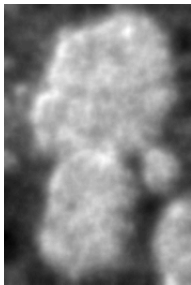
**ROF denoising** (nonlinear)

$$\underset{\mathbf{h}}{\operatorname{argmin}} \|\mathbf{h} - \hat{\mathbf{h}}^{\text{LR}}\|^2 + \lambda \|\mathbf{D}\mathbf{h}\|_{2,1}$$

Linear regression  $\hat{\mathbf{h}}^{\text{LR}}$



Lissage



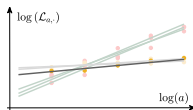
ROF



→ cumulative estimation variance and regularization bias

$$\sum_a \frac{\|\log \mathcal{L}_{a,\cdot} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}}$$

→ fidelity to the log-linear model

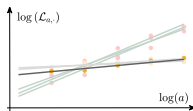


# Functionals with either free or co-localized contours

$$\sum_a \frac{\|\log \mathcal{L}_{a,\cdot} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}} + \lambda \frac{Q(\mathbf{Dh}, \mathbf{Dv}; \alpha)}{\text{Total Variation}}$$

→ fidelity to the log-linear model

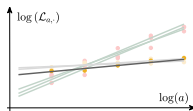
→ favors piecewise constancy



# Functionals with either free or co-localized contours

$$\underset{h, v}{\text{minimize}} \quad \sum_a \frac{\|\log \mathcal{L}_{a,\cdot} - \log(a)h - v\|^2}{\text{Least-Squares}} + \lambda \frac{Q(Dh, Dv; \alpha)}{\text{Total Variation}}$$

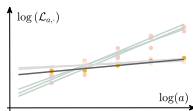
$\rightarrow$  fidelity to the log-linear model  $\rightarrow$  favors piecewise constancy



# Functionals with either free or co-localized contours

$$\underset{h, v}{\text{minimize}} \sum_a \frac{\|\log \mathcal{L}_{a,\cdot} - \log(a)h - v\|^2}{\text{Least-Squares}} + \lambda \frac{Q(\mathbf{D}h, \mathbf{D}v; \alpha)}{\text{Total Variation}}$$

→ fidelity to the log-linear model                      → favors piecewise constancy



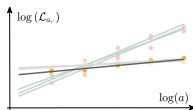
**Finite differences**  $\mathbf{D}_1\mathbf{x}$  (horizontal),  $\mathbf{D}_2\mathbf{x}$  (vertical) in each pixel



# Functionals with either free or co-localized contours

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \frac{\|\log \mathcal{L}_{a,\cdot} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}} + \lambda \frac{Q(\mathbf{Dh}, \mathbf{Dv}; \alpha)}{\text{Total Variation}}$$

$\rightarrow$  fidelity to the log-linear model  $\rightarrow$  favors piecewise constancy



**Finite differences**  $\mathbf{Dx} = [\mathbf{D}_1\mathbf{x}, \mathbf{D}_2\mathbf{x}]$

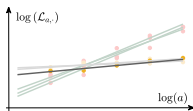
Free:  $\mathbf{h}, \mathbf{v}$  are **independently** piecewise constant

$$Q_F(\mathbf{Dh}, \mathbf{Dv}; \alpha) = \alpha \|\mathbf{Dh}\|_{2,1} + \|\mathbf{Dv}\|_{2,1}$$

# Functionals with either free or co-localized contours

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \frac{\|\log \mathcal{L}_{a,\cdot} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}} + \lambda \frac{Q(\mathbf{Dh}, \mathbf{Dv}; \alpha)}{\text{Total Variation}}$$

$\rightarrow$  fidelity to the log-linear model  $\rightarrow$  favors piecewise constancy



## Finite differences $\mathbf{Dx} = [\mathbf{D}_1\mathbf{x}, \mathbf{D}_2\mathbf{x}]$

Free:  $\mathbf{h}, \mathbf{v}$  are **independently** piecewise constant

$$Q_F(\mathbf{Dh}, \mathbf{Dv}; \alpha) = \alpha \|\mathbf{Dh}\|_{2,1} + \|\mathbf{Dv}\|_{2,1}$$

Co-localized:  $\mathbf{h}, \mathbf{v}$  are **concomitantly** piecewise constant

$$Q_C(\mathbf{Dh}, \mathbf{Dv}; \alpha) = \|[\alpha \mathbf{Dh}, \mathbf{Dv}]\|_{2,1}$$

# Functionals minimization

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \underbrace{\|\log \mathcal{L}_{a,\cdot} - \log(a)\mathbf{h} - \mathbf{v}\|^2}_{\text{Least-Squares}} + \lambda \underbrace{Q(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)}_{\text{Total Variation}}$$



# Functionals minimization

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \frac{\|\log \mathcal{L}_{a,\cdot} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}} + \lambda \frac{Q(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)}{\text{Total Variation}}$$



► gradient descent  $\mathbf{x}^{n+1} = \mathbf{x}^n - \tau \nabla \varphi(\mathbf{x}^n)$

# Functionals minimization

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \underbrace{\|\log \mathcal{L}_{a,\cdot} - \log(a)\mathbf{h} - \mathbf{v}\|^2}_{\text{Least-Squares}} + \lambda \underbrace{Q(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)}_{\text{Total Variation}}$$



nonsmooth



- ▶ gradient descent  $\mathbf{x}^{n+1} = \mathbf{x}^n - \tau \nabla \varphi(\mathbf{x}^n)$
- ▶ implicit subgradient descent: proximal point algorithm  
 $\mathbf{x}^{n+1} = \mathbf{x}^n - \tau \mathbf{u}^n, \mathbf{u}^n \in \partial \varphi(\mathbf{x}^{n+1}) \Leftrightarrow \mathbf{x}^{n+1} = \text{prox}_{\tau \varphi}(\mathbf{x}^n)$

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \frac{\|\log \mathcal{L}_{a,\cdot} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}} + \lambda \frac{Q(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)}{\text{Total Variation}}$$



nonsmooth



► gradient descent  $\mathbf{x}^{n+1} = \mathbf{x}^n - \tau \nabla \varphi(\mathbf{x}^n)$

► implicit subgradient descent: proximal point algorithm

$$\mathbf{x}^{n+1} = \mathbf{x}^n - \tau \mathbf{u}^n, \quad \mathbf{u}^n \in \partial \varphi(\mathbf{x}^{n+1}) \Leftrightarrow \mathbf{x}^{n+1} = \text{prox}_{\tau \varphi}(\mathbf{x}^n)$$

► splitting proximal algorithm

$$\mathbf{y}^{n+1} = \text{prox}_{\sigma(\lambda \mathcal{Q})^*}(\mathbf{y}^n + \sigma \mathbf{D} \bar{\mathbf{x}}^n)$$

$$\mathbf{x}^{n+1} = \text{prox}_{\tau \|\mathcal{L} - \Phi \cdot\|_2^2}(\mathbf{x}^n - \tau \mathbf{D}^\top \mathbf{y}^{n+1}), \quad \Phi : (\mathbf{h}, \mathbf{v}) \mapsto \{\log(a)\mathbf{h} + \mathbf{v}\}_a$$

$$\bar{\mathbf{x}}^{n+1} = 2\mathbf{x}^{n+1} - \mathbf{x}^n$$

# Functionals minimization

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \frac{\|\log \mathcal{L}_{a,\cdot} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}} + \lambda \frac{Q(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)}{\text{Total Variation}}$$



nonsmooth



► gradient descent  $\mathbf{x}^{n+1} = \mathbf{x}^n - \tau \nabla \varphi(\mathbf{x}^n)$

► implicit subgradient descent: proximal point algorithm

$$\mathbf{x}^{n+1} = \mathbf{x}^n - \tau \mathbf{u}^n, \quad \mathbf{u}^n \in \partial \varphi(\mathbf{x}^{n+1}) \Leftrightarrow \mathbf{x}^{n+1} = \text{prox}_{\tau \varphi}(\mathbf{x}^n)$$

► splitting proximal algorithm

$$\mathbf{y}^{n+1} = \text{prox}_{\sigma(\lambda \mathcal{Q})^*}(\mathbf{y}^n + \sigma \mathbf{D} \bar{\mathbf{x}}^n)$$

$$\mathbf{x}^{n+1} = \text{prox}_{\tau \|\mathcal{L} - \Phi \cdot\|_2^2}(\mathbf{x}^n - \tau \mathbf{D}^\top \mathbf{y}^{n+1}), \quad \Phi : (\mathbf{h}, \mathbf{v}) \mapsto \{\log(a)\mathbf{h} + \mathbf{v}\}_a$$

$$\bar{\mathbf{x}}^{n+1} = 2\mathbf{x}^{n+1} - \mathbf{x}^n$$

# Accelerated algorithm based on strong-convexity

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \underbrace{\|\log \mathcal{L}_{a,\cdot} - \log(a)\mathbf{h} - \mathbf{v}\|^2}_{\text{Least-Squares}} + \lambda \underbrace{Q(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)}_{\text{Total Variation}}$$



nonsmooth



Primal-dual algorithm (Chambolle, 2011)

$$\delta: \text{duality gap}, \quad \delta(\mathbf{x}^n, \mathbf{y}^n) \xrightarrow{n \rightarrow \infty} 0$$



# Accelerated algorithm based on strong-convexity

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \underbrace{\|\log \mathcal{L}_{a,\cdot} - \log(a)\mathbf{h} - \mathbf{v}\|^2}_{\text{Least-Squares}} + \lambda \underbrace{Q(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)}_{\text{Total Variation}}$$

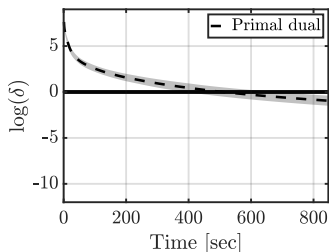


nonsmooth



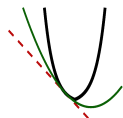
Primal-dual algorithm (Chambolle, 2011)

$\delta$ : duality gap,  $\delta(\mathbf{x}^n, \mathbf{y}^n) \xrightarrow{n \rightarrow \infty} 0$



# Convexity properties

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \frac{\|\log \mathcal{L}_{a,\cdot} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}} + \lambda \frac{Q(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)}{\text{Total Variation}}$$



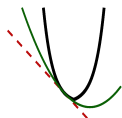
$\mu$ -strongly convex

nonsmooth



# Convexity properties

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \frac{\|\log \mathcal{L}_{a,\cdot} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}} + \lambda \frac{Q(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)}{\text{Total Variation}}$$



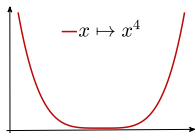
$\mu$ -strongly convex

nonsmooth

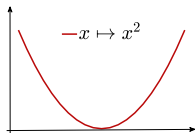


## Strong-convexity

- $\varphi$   $\mu$ -strongly convex iff  $\varphi - \frac{\mu}{2} \|\cdot\|^2$  convex

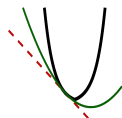


- ✓ strictly convex
- ✗ non strongly convex



- ✓ strictly convex
- ✓ 1-strongly convex

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \frac{\|\log \mathcal{L}_{a,\cdot} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}} + \lambda \frac{Q(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)}{\text{Total Variation}}$$



$\mu$ -strongly convex

nonsmooth

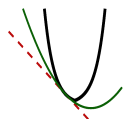


## Strong-convexity

- $\varphi$   $\mu$ -strongly convex iff  $\varphi - \frac{\mu}{2} \|\cdot\|^2$  convex
- $\varphi \in \mathcal{C}^2$  with Hessian matrix  $\mathbf{H}\varphi \succeq 0 \implies \mu = \min \text{Sp}(\mathbf{H}\varphi)$

# Convexity properties

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \underbrace{\|\log \mathcal{L}_{a,\cdot} - \log(a)\mathbf{h} - \mathbf{v}\|^2}_{\text{Least-Squares}} + \lambda \underbrace{Q(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)}_{\text{Total Variation}}$$



$\mu$ -strongly convex

nonsmooth



## Strong-convexity

- $\varphi$   $\mu$ -strongly convex iff  $\varphi - \frac{\mu}{2} \|\cdot\|^2$  convex
- $\varphi \in \mathcal{C}^2$  with Hessian matrix  $\mathbf{H}\varphi \succeq 0 \implies \mu = \min \text{Sp}(\mathbf{H}\varphi)$

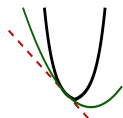
Proposition (Pascal, 2019)

$\sum_a \|\log \mathcal{L} - \log(a)\mathbf{h} - \mathbf{v}\|^2$  is  $\mu$ -strongly convex.

$a_{\min} = 2^1,$	$a_{\max}$	$2^2$	$2^3$	$2^4$	$2^5$	$2^6$
$\mu = \min \text{Sp} (2\Phi^\top \Phi)$	0.29	<b>0.72</b>	1.20	1.69	2.20	

# Accelerated algorithm based on strong-convexity

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \frac{\|\log \mathcal{L}_{a,\cdot} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}} + \lambda \frac{Q(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)}{\text{Total Variation}}$$



$\mu$ -strongly convex

nonsmooth



**Accelerated Primal-dual algorithm** (*Chambolle, 2011*)

**for**  $n = 0, 1, \dots$

$\mathbf{x} = (\mathbf{h}, \mathbf{v})$

$$\mathbf{y}^{n+1} = \text{prox}_{\sigma_n(\lambda Q)^*}(\mathbf{y}^n + \sigma_n \mathbf{D}\bar{\mathbf{x}}^n)$$

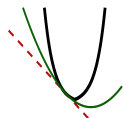
$$\mathbf{x}^{n+1} = \text{prox}_{\tau_n \|\mathcal{L} - \Phi \cdot\|_2^2}(\mathbf{x}^n - \tau_n \mathbf{D}^\top \mathbf{y}^{n+1})$$

$$\theta_n = \sqrt{1 + 2\mu\tau_n}, \quad \tau_{n+1} = \tau_n / \theta_n, \quad \sigma_{n+1} = \theta_n \sigma_n$$

$$\bar{\mathbf{x}}^{n+1} = \mathbf{x}^{n+1} + \theta_n^{-1}(\mathbf{x}^{n+1} - \mathbf{x}^n)$$

# Accelerated algorithm based on strong-convexity

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \frac{\|\log \mathcal{L}_{a,\cdot} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}} + \lambda \frac{Q(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)}{\text{Total Variation}}$$



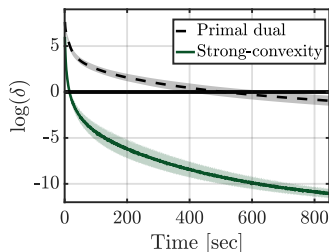
$\mu$ -strongly convex

nonsmooth



**Accelerated Primal-dual algorithm** (*Chambolle, 2011*)

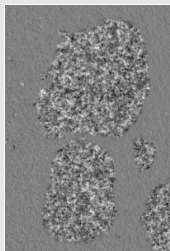
$\delta$ : duality gap,  $\delta(\mathbf{x}^n, \mathbf{y}^n) \xrightarrow{\dots} 0$



## Segmentation *via* iterated thresholding

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \frac{\|\log \mathcal{L}_{a,\cdot} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}} + \lambda \frac{Q(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)}{\text{Total Variation}}$$

Textured image    Lin. reg.  $\hat{\mathbf{h}}^{\text{LR}}$

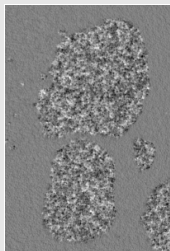




# Segmentation *via* iterated thresholding

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \frac{\|\log \mathcal{L}_{a,\cdot} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}} + \lambda \frac{Q(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)}{\text{Total Variation}}$$

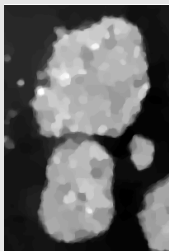
Textured image



Lin. reg.  $\hat{\mathbf{h}}^{\text{LR}}$



Co-localized  
contours  $\hat{\mathbf{h}}^{\text{C}}$



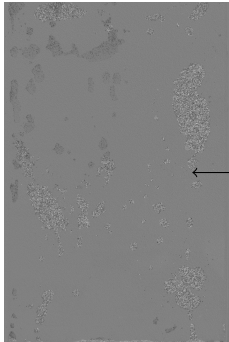
Threshold  
estimate<sup>†</sup>  $T\hat{\mathbf{h}}^{\text{C}}$



<sup>†</sup>(Cai, 2013)

# Multiphase flow through porous media

Laboratoire de Physique, ENS Lyon, V. Vidal, T. Busser, (M. Serres, IFPEN)

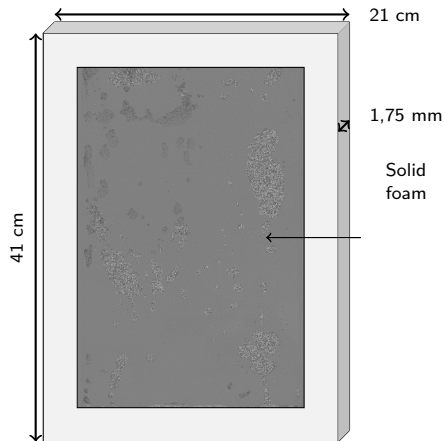


Solid  
foam



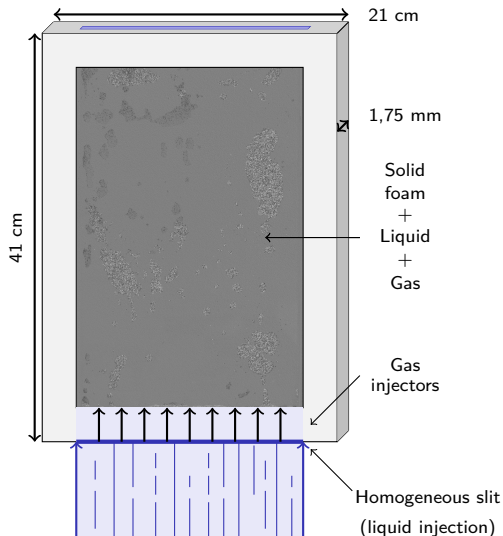
# Multiphase flow through porous media

Laboratoire de Physique, ENS Lyon, V. Vidal, T. Busser, (M. Serres, IFPEN)



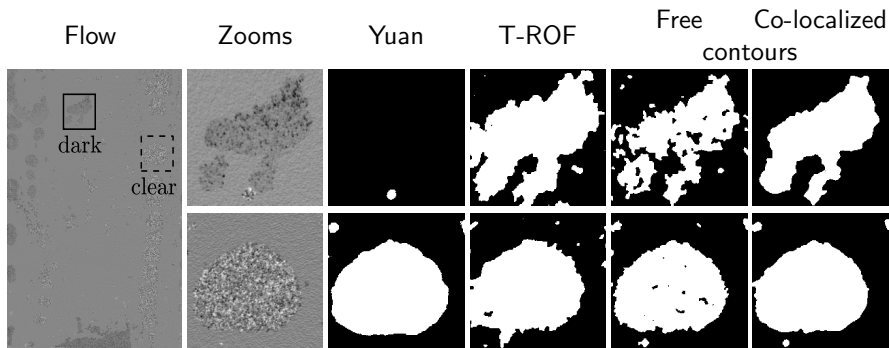
# Multiphase flow through porous media

Laboratoire de Physique, ENS Lyon, V. Vidal, T. Busser, (M. Serres, IFPEN)

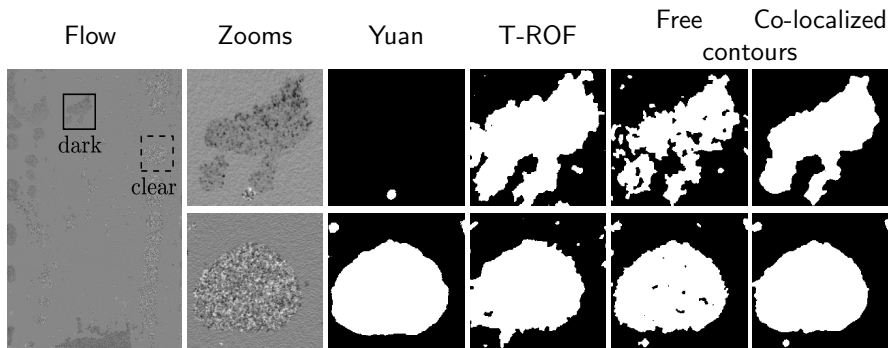


- 1600 × 1100 pixels
- video: ~ 1000 images
- phase diagram: ~ 10 flow rates

Low activity:  $Q_G = 300\text{mL}/\text{min}$  -  $Q_L = 300\text{mL}/\text{min}$



Low activity:  $Q_G = 300\text{mL}/\text{min} - Q_L = 300\text{mL}/\text{min}$



Liquid:  $h_L = 0.4$

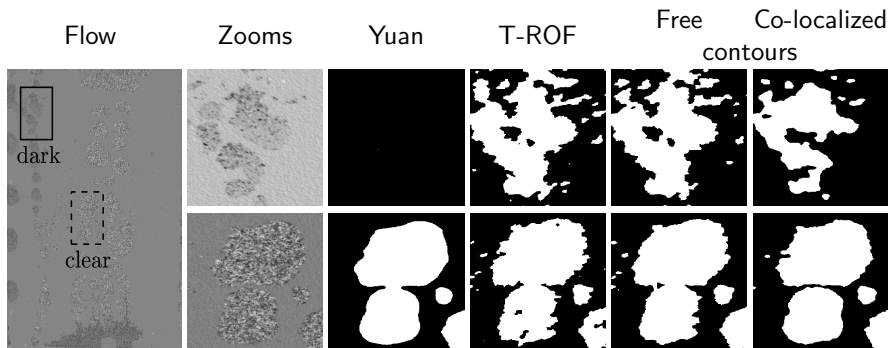
$$\sigma_{\text{dark}}^2 = 10^{-2}$$

Gas:  $h_G = 0.9$

$$\sigma_{\text{dark}}^2 = 10^{-2} \quad (\text{dark bubbles})$$

$$\sigma_{\text{clear}}^2 = 10^{-1} \quad (\text{clear bubbles})$$

Transition:  $Q_G = 400\text{mL}/\text{min} - Q_L = 700\text{mL}/\text{min}$



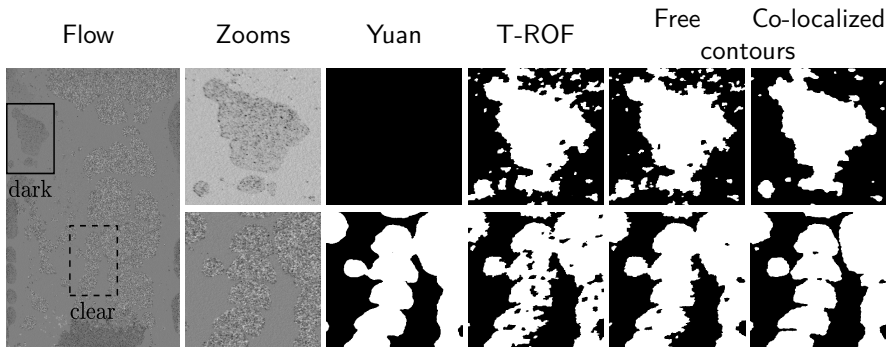
Liquid:  $h_L = 0.4$

$$\sigma_{\text{dark}}^2 = 10^{-2}$$

Gas:  $h_G = 0.9$

$$\left. \begin{array}{l} \sigma_{\text{dark}}^2 = 10^{-2} \quad (\text{dark bubbles}) \\ \sigma_{\text{clear}}^2 = 10^{-1} \quad (\text{clear bubbles}). \end{array} \right\}$$

High activity:  $Q_G = 1200\text{mL}/\text{min} - Q_L = 300\text{mL}/\text{min}$



Liquid:  $h_L = 0.4$

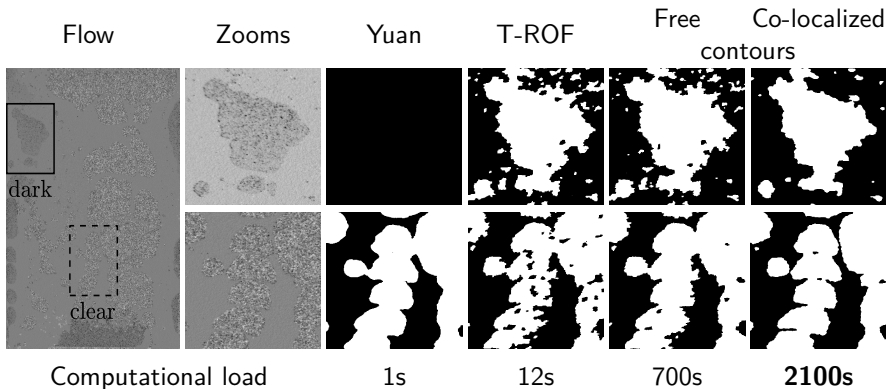
$$\sigma_{\text{dark}}^2 = 10^{-2}$$

Gas:  $h_G = 0.9$

$$\left. \begin{array}{l} \sigma_{\text{dark}}^2 = 10^{-2} \\ \sigma_{\text{clear}}^2 = 10^{-1} \end{array} \right\} \begin{array}{l} \text{(dark bubbles)} \\ \text{(clear bubbles)}. \end{array}$$



High activity:  $Q_G = 1200\text{mL}/\text{min} - Q_L = 300\text{mL}/\text{min}$



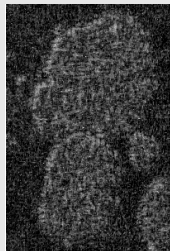
Liquid: $h_L = 0.4$	$\sigma_{\text{dark}}^2 = 10^{-2}$	
Gas: $h_G = 0.9$	$\sigma_{\text{dark}}^2 = 10^{-2}$	(dark bubbles)
	$\sigma_{\text{clear}}^2 = 10^{-1}$	(clear bubbles).

$$\left(\hat{\mathbf{h}}, \hat{\mathbf{v}}\right) (\mathcal{L}; \lambda, \alpha) = \operatorname{argmin}_{\mathbf{h}, \mathbf{v}} \sum_a \left\| \log \mathcal{L}_{a,\cdot} - \log(a) \mathbf{h} - \mathbf{v} \right\|^2 + \lambda Q(\mathbf{Dh}, \mathbf{Dv}; \alpha)$$

$$(\hat{\mathbf{h}}, \hat{\mathbf{v}}) (\mathcal{L}; \lambda, \alpha) = \operatorname{argmin}_{\mathbf{h}, \mathbf{v}} \sum_a \|\log \mathcal{L}_{a,\cdot} - \log(a)\mathbf{h} - \mathbf{v}\|^2 + \lambda Q(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)$$

Lin. reg.  $\hat{\mathbf{h}}^{\text{LR}}$

$(\lambda, \alpha) = (0, 0)$



$$(\hat{\mathbf{h}}, \hat{\mathbf{v}})(\mathcal{L}; \lambda, \alpha) = \operatorname{argmin}_{\mathbf{h}, \mathbf{v}} \sum_a \|\log \mathcal{L}_{a,\cdot} - \log(a)\mathbf{h} - \mathbf{v}\|^2 + \lambda Q(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)$$

Lin. reg.  $\hat{\mathbf{h}}^{\text{LR}}$

$(\lambda, \alpha) = (0, 0)$



Co-localized contours estimate  $\hat{\mathbf{h}}^{\text{C}}$

$(\lambda, \alpha) = (0.5, 0.5)$



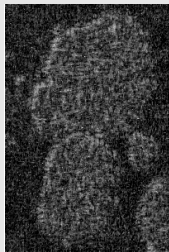
too small

## Regularization parameters selection

$$(\hat{\mathbf{h}}, \hat{\mathbf{v}})(\mathcal{L}; \lambda, \alpha) = \operatorname{argmin}_{\mathbf{h}, \mathbf{v}} \sum_a \|\log \mathcal{L}_{a,\cdot} - \log(a)\mathbf{h} - \mathbf{v}\|^2 + \lambda Q(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)$$

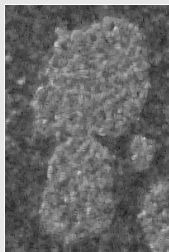
Lin. reg.  $\hat{\mathbf{h}}^{\text{LR}}$

$(\lambda, \alpha) = (0, 0)$



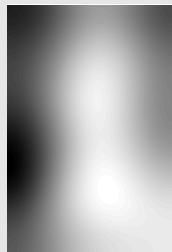
Co-localized contours estimate  $\hat{\mathbf{h}}^{\text{C}}$

$(\lambda, \alpha) = (0.5, 0.5)$



too small

$(\lambda, \alpha) = (500, 500)$



too large

## Regularization parameters selection

$$(\hat{\mathbf{h}}, \hat{\mathbf{v}}) (\mathcal{L}; \lambda, \alpha) = \operatorname{argmin}_{\mathbf{h}, \mathbf{v}} \sum_a \|\log \mathcal{L}_{a,\cdot} - \log(a)\mathbf{h} - \mathbf{v}\|^2 + \lambda Q(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)$$

Lin. reg.  $\hat{\mathbf{h}}^{\text{LR}}$

$(\lambda, \alpha) = (0, 0)$



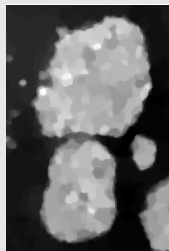
Co-localized contours estimate  $\hat{\mathbf{h}}^{\text{C}}$

$(\lambda, \alpha) = (0.5, 0.5)$



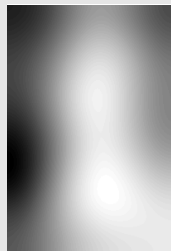
too small

$(\lambda^\dagger, \alpha^\dagger) = (11.5, 0.8)$



optimal

$(\lambda, \alpha) = (500, 500)$



too large

What *optimal* means? How to determine  $\lambda^\dagger$  and  $\alpha^\dagger$ ?

## Parameter tuning (Grid search)

$$\left(\hat{\mathbf{h}}, \hat{\mathbf{v}}\right) (\mathcal{L}; \lambda, \alpha) = \underset{\mathbf{h}, \mathbf{v}}{\operatorname{argmin}} \sum_a \|\log \mathcal{L}_{a..} - \log(a) \mathbf{h} - \mathbf{v}\|^2 + \lambda \mathcal{Q}(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)$$

*$\mathbf{h}$ : discriminant,  $\mathbf{v}$ : auxiliary*

## Parameter tuning (Grid search)

$$\left(\hat{\mathbf{h}}, \hat{\mathbf{v}}\right) (\mathcal{L}; \lambda, \alpha) = \underset{\mathbf{h}, \mathbf{v}}{\operatorname{argmin}} \sum_a \left\| \log \mathcal{L}_{a..} - \log(a) \mathbf{h} - \mathbf{v} \right\|^2 + \lambda \mathcal{Q}(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)$$

$\mathbf{h}$ : discriminant,  $\mathbf{v}$ : auxiliary

$\bar{\mathbf{h}}$ : true regularity

$$\mathcal{R}(\lambda, \alpha) = \left\| \hat{\mathbf{h}}(\mathcal{L}; \lambda, \alpha) - \bar{\mathbf{h}} \right\|^2$$



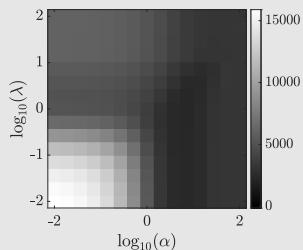
## Parameter tuning (Grid search)

$$\left(\hat{\mathbf{h}}, \hat{\mathbf{v}}\right) (\mathcal{L}; \lambda, \alpha) = \underset{\mathbf{h}, \mathbf{v}}{\operatorname{argmin}} \sum_a \left\| \log \mathcal{L}_{a..} - \log(a) \mathbf{h} - \mathbf{v} \right\|^2 + \lambda \mathcal{Q}(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)$$

$\mathbf{h}$ : discriminant,  $\mathbf{v}$ : auxiliary

$\bar{\mathbf{h}}$ : true regularity

$$\mathcal{R}(\lambda, \alpha) = \left\| \hat{\mathbf{h}}(\mathcal{L}; \lambda, \alpha) - \bar{\mathbf{h}} \right\|^2$$



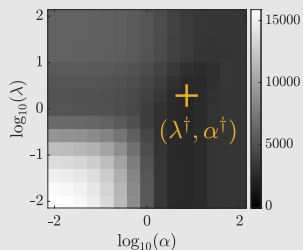
## Parameter tuning (Grid search)

$$\left(\hat{\mathbf{h}}, \hat{\mathbf{v}}\right) (\mathcal{L}; \lambda, \alpha) = \underset{\mathbf{h}, \mathbf{v}}{\operatorname{argmin}} \sum_a \left\| \log \mathcal{L}_{a..} - \log(a) \mathbf{h} - \mathbf{v} \right\|^2 + \lambda \mathcal{Q}(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)$$

$\mathbf{h}$ : discriminant,  $\mathbf{v}$ : auxiliary

$\bar{\mathbf{h}}$ : true regularity

$$\mathcal{R}(\lambda, \alpha) = \left\| \hat{\mathbf{h}}(\mathcal{L}; \lambda, \alpha) - \bar{\mathbf{h}} \right\|^2$$



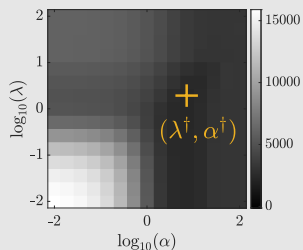
## Parameter tuning (Grid search)

$$\left(\hat{\mathbf{h}}, \hat{\mathbf{v}}\right) (\mathcal{L}; \lambda, \alpha) = \underset{\mathbf{h}, \mathbf{v}}{\operatorname{argmin}} \sum_a \left\| \log \mathcal{L}_{a..} - \log(a) \mathbf{h} - \mathbf{v} \right\|^2 + \lambda \mathcal{Q}(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)$$

$\mathbf{h}$ : discriminant,  $\mathbf{v}$ : auxiliary

$\bar{\mathbf{h}}$ : true regularity

$$\mathcal{R}(\lambda, \alpha) = \left\| \hat{\mathbf{h}}(\mathcal{L}; \lambda, \alpha) - \bar{\mathbf{h}} \right\|^2$$



$\bar{\mathbf{h}}$ : unknown!

?

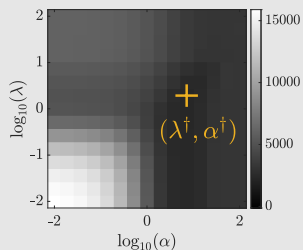
## Parameter tuning (Grid search)

$$\left(\hat{\mathbf{h}}, \hat{\mathbf{v}}\right) (\mathcal{L}; \lambda, \alpha) = \underset{\mathbf{h}, \mathbf{v}}{\operatorname{argmin}} \sum_a \left\| \log \mathcal{L}_{a..} - \log(a) \mathbf{h} - \mathbf{v} \right\|^2 + \lambda \mathcal{Q}(\mathbf{D} \mathbf{h}, \mathbf{D} \mathbf{v}; \alpha)$$

$\mathbf{h}$ : discriminant,  $\mathbf{v}$ : auxiliary

$\bar{\mathbf{h}}$ : true regularity

$$\mathcal{R}(\lambda, \alpha) = \left\| \hat{\mathbf{h}}(\mathcal{L}; \lambda, \alpha) - \bar{\mathbf{h}} \right\|^2$$



$\bar{\mathbf{h}}$ : unknown!

?

Stein Unbiased Risk Estimate  
(SURE)

## Stein Unbiased Risk Estimate (Principe)

**Observations**  $\mathbf{y} = \bar{\mathbf{x}} + \boldsymbol{\zeta} \in \mathbb{R}^P$ ,  $\bar{\mathbf{x}}$ : truth and  $\boldsymbol{\zeta} \sim \mathcal{N}(\mathbf{0}, \rho^2 \mathbf{I})$

**Observations**  $\mathbf{y} = \bar{\mathbf{x}} + \boldsymbol{\zeta} \in \mathbb{R}^P$ ,  $\bar{\mathbf{x}}$ : truth and  $\boldsymbol{\zeta} \sim \mathcal{N}(\mathbf{0}, \rho^2 \mathbf{I})$

**Parametric estimator**  $(\mathbf{y}; \lambda) \mapsto \hat{\mathbf{x}}(\mathbf{y}; \lambda)$

$$\text{Ex. } \hat{\mathbf{x}}(\mathbf{y}; \lambda) = \begin{cases} (\mathbf{I} + \lambda \mathbf{D}^\top \mathbf{D})^{-1} \mathbf{y} & \text{(linear)} \\ \underset{\mathbf{x}}{\operatorname{argmin}} \|\mathbf{y} - \mathbf{x}\|^2 + \lambda \mathcal{Q}(\mathbf{D}\mathbf{x}) & \text{(nonlinear)} \end{cases}$$

## Stein Unbiased Risk Estimate (Principe)

**Observations**  $\mathbf{y} = \bar{\mathbf{x}} + \boldsymbol{\zeta} \in \mathbb{R}^P$ ,  $\bar{\mathbf{x}}$ : truth and  $\boldsymbol{\zeta} \sim \mathcal{N}(\mathbf{0}, \rho^2 \mathbf{I})$

**Parametric estimator**  $(\mathbf{y}; \lambda) \mapsto \hat{\mathbf{x}}(\mathbf{y}; \lambda)$

$$\text{Ex. } \hat{\mathbf{x}}(\mathbf{y}; \lambda) = \begin{cases} (\mathbf{I} + \lambda \mathbf{D}^\top \mathbf{D})^{-1} \mathbf{y} & \text{(linear)} \\ \underset{\mathbf{x}}{\operatorname{argmin}} \|\mathbf{y} - \mathbf{x}\|^2 + \lambda \mathcal{Q}(\mathbf{D}\mathbf{x}) & \text{(nonlinear)} \end{cases}$$

**Quadratic error**  $R(\lambda) \triangleq \mathbb{E}_{\boldsymbol{\zeta}} \|\hat{\mathbf{x}}(\mathbf{y}; \lambda) - \bar{\mathbf{x}}\|^2 \stackrel{?}{=} \mathbb{E}_{\boldsymbol{\zeta}} \hat{R}(\mathbf{y}; \lambda)$

$\bar{\mathbf{x}}$  unknown

## Stein Unbiased Risk Estimate (Principle)

**Observations**  $\mathbf{y} = \bar{\mathbf{x}} + \boldsymbol{\zeta} \in \mathbb{R}^P$ ,  $\bar{\mathbf{x}}$ : truth and  $\boldsymbol{\zeta} \sim \mathcal{N}(\mathbf{0}, \rho^2 \mathbf{I})$

**Parametric estimator**  $(\mathbf{y}; \lambda) \mapsto \hat{\mathbf{x}}(\mathbf{y}; \lambda)$

**Ex.** 
$$\hat{\mathbf{x}}(\mathbf{y}; \lambda) = \begin{cases} (\mathbf{I} + \lambda \mathbf{D}^\top \mathbf{D})^{-1} \mathbf{y} & \text{(linear)} \\ \underset{\mathbf{x}}{\operatorname{argmin}} \|\mathbf{y} - \mathbf{x}\|^2 + \lambda \mathcal{Q}(\mathbf{D}\mathbf{x}) & \text{(nonlinear)} \end{cases}$$

**Quadratic error**  $R(\lambda) \triangleq \mathbb{E}_{\boldsymbol{\zeta}} \|\hat{\mathbf{x}}(\mathbf{y}; \lambda) - \bar{\mathbf{x}}\|^2 \stackrel{?}{=} \mathbb{E}_{\boldsymbol{\zeta}} \hat{R}(\mathbf{y}; \lambda)$   $\bar{\mathbf{x}}$  unknown

**Theorem** (Stein, 1981)

Let  $(\mathbf{y}; \lambda) \mapsto \hat{\mathbf{x}}(\mathbf{y}; \lambda)$  an estimator of  $\bar{\mathbf{x}}$

- weakly differentiable w.r.t.  $\mathbf{y}$ ,
- such that  $\boldsymbol{\zeta} \mapsto \langle \hat{\mathbf{x}}(\bar{\mathbf{x}} + \boldsymbol{\zeta}; \lambda), \boldsymbol{\zeta} \rangle$  is integrable w.r.t.  $\mathcal{N}(\mathbf{0}, \rho^2 \mathbf{I})$ .

$$\begin{aligned} \hat{R}(\mathbf{y}; \lambda) &\triangleq \|\hat{\mathbf{x}}(\mathbf{y}; \lambda) - \mathbf{y}\|^2 + 2\rho^2 \operatorname{tr}(\partial_{\mathbf{y}} \hat{\mathbf{x}}(\mathbf{y}; \lambda)) - \rho^2 P \\ &\implies R(\lambda) = \mathbb{E}_{\boldsymbol{\zeta}} [\hat{R}(\mathbf{y}; \lambda)]. \end{aligned}$$



# Generalized Stein Unbiased Risk Estimate

**Observations**  $\mathbf{y} = \Phi \bar{\mathbf{x}} + \zeta \in \mathbb{R}^P$ ,  $\bar{\mathbf{x}} \in \mathbb{R}^N$ ,  $\Phi : \mathbb{R}^{P \times N}$  and  $\zeta \sim \mathcal{N}(\mathbf{0}, \mathcal{S})$

**E.g. the estimators  $\hat{\mathbf{h}}(\mathcal{L}; \lambda, \alpha)$  with free or co-localized contours**

$$\log \mathcal{L} = \Phi(\bar{\mathbf{h}}, \bar{\mathbf{v}}) + \zeta \quad \zeta \sim \mathcal{N}(\mathbf{0}, \mathcal{S}) \quad \mathcal{R} = \|\hat{\mathbf{h}} - \bar{\mathbf{h}}\|^2$$

$$\Phi : (\mathbf{h}, \mathbf{v}) \mapsto \{\log(a)\mathbf{h} + \mathbf{v}\}_a$$


$$\Pi : (\mathbf{h}, \mathbf{v}) \mapsto (\mathbf{h}, \mathbf{0})$$

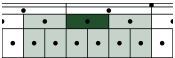
**Projected estimation error**  $R_{\Pi}(\Lambda) \triangleq \mathbb{E}_{\zeta} \|\Pi \hat{\mathbf{x}}(\mathbf{y}; \Lambda) - \Pi \bar{\mathbf{x}}\|^2$

# Generalized Stein Unbiased Risk Estimate

**Observations**  $\mathbf{y} = \Phi \bar{\mathbf{x}} + \zeta \in \mathbb{R}^P$ ,  $\bar{\mathbf{x}} \in \mathbb{R}^N$ ,  $\Phi : \mathbb{R}^{P \times N}$  and  $\zeta \sim \mathcal{N}(\mathbf{0}, \mathcal{S})$

**E.g. the estimators  $\hat{\mathbf{h}}(\mathcal{L}; \lambda, \alpha)$  with free or co-localized contours**

$$\log \mathcal{L} = \Phi(\bar{\mathbf{h}}, \bar{\mathbf{v}}) + \zeta \quad \zeta \sim \mathcal{N}(\mathbf{0}, \mathcal{S}) \quad \mathcal{R} = \|\hat{\mathbf{h}} - \bar{\mathbf{h}}\|^2$$

$$\Phi : (\mathbf{h}, \mathbf{v}) \mapsto \{\log(a)\mathbf{h} + \mathbf{v}\}_a \quad \Pi : (\mathbf{h}, \mathbf{v}) \mapsto (\mathbf{h}, \mathbf{0})$$


**Projected estimation error**  $R_{\Pi}(\Lambda) \triangleq \mathbb{E}_{\zeta} \|\Pi \hat{\mathbf{x}}(\mathbf{y}; \Lambda) - \Pi \bar{\mathbf{x}}\|^2$

**Theorem** (Pascal, 2020)

Let  $(\mathbf{y}; \Lambda) \mapsto \hat{\mathbf{x}}(\mathbf{y}; \Lambda)$  an estimator of  $\bar{\mathbf{x}}$

- weakly differentiable w.r.t.  $\mathbf{y}$ ,
- such that  $\zeta \mapsto \langle \Pi \hat{\mathbf{x}}(\bar{\mathbf{x}} + \zeta; \lambda), \mathbf{A} \zeta \rangle$  is integrable w.r.t.  $\mathcal{N}(\mathbf{0}, \mathcal{S})$ .

$$\hat{R}(\Lambda) \triangleq \|\mathbf{A}(\Phi \hat{\mathbf{x}}(\mathbf{y}; \Lambda) - \mathbf{y})\|^2 + 2 \text{tr} \left( \mathcal{S} \mathbf{A}^{\top} \Pi \partial_{\mathbf{y}} \hat{\mathbf{x}}(\mathbf{y}; \Lambda) \right) - \text{tr} \left( \mathbf{A} \mathcal{S} \mathbf{A}^{\top} \right)$$

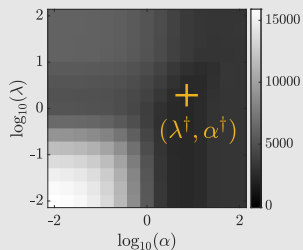
$$\implies R_{\Pi}(\Lambda) = \mathbb{E}_{\zeta} [\hat{R}(\Lambda)].$$

## Parameter tuning (Grid search)

$$\left(\hat{\mathbf{h}}, \hat{\mathbf{v}}\right) (\mathcal{L}; \lambda, \alpha) = \operatorname{argmin}_{\mathbf{h}, \mathbf{v}} \sum_a \left\| \log \mathcal{L}_{a,\cdot} - \log(a) \mathbf{h} - \mathbf{v} \right\|^2 + \lambda Q(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)$$

$\bar{\mathbf{h}}$ : true regularity

$$\mathcal{R}(\lambda, \alpha) = \left\| \hat{\mathbf{h}}(\mathcal{L}; \lambda, \alpha) - \bar{\mathbf{h}} \right\|^2$$



$\bar{\mathbf{h}}$ : unknown!

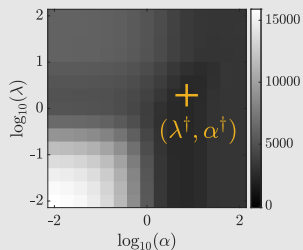
$$\hat{R}_{\nu, \epsilon}(\mathcal{L}; \lambda, \alpha | \mathcal{S})$$

# Parameter tuning (Grid search)

$$\left(\hat{\mathbf{h}}, \hat{\mathbf{v}}\right) (\mathcal{L}; \lambda, \alpha) = \operatorname{argmin}_{\mathbf{h}, \mathbf{v}} \sum_a \left\| \log \mathcal{L}_{a,\cdot} - \log(a) \mathbf{h} - \mathbf{v} \right\|^2 + \lambda Q(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)$$

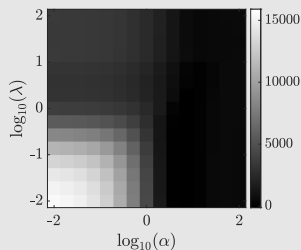
$\bar{\mathbf{h}}$ : true regularity

$$\mathcal{R}(\lambda, \alpha) = \left\| \hat{\mathbf{h}}(\mathcal{L}; \lambda, \alpha) - \bar{\mathbf{h}} \right\|^2$$



$\bar{\mathbf{h}}$ : unknown!

$$\hat{R}_{\nu, \epsilon}(\mathcal{L}; \lambda, \alpha | \mathcal{S})$$

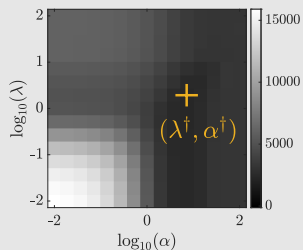


# Parameter tuning (Grid search)

$$\left(\hat{\mathbf{h}}, \hat{\mathbf{v}}\right) (\mathcal{L}; \lambda, \alpha) = \operatorname{argmin}_{\mathbf{h}, \mathbf{v}} \sum_a \left\| \log \mathcal{L}_{a,\cdot} - \log(a) \mathbf{h} - \mathbf{v} \right\|^2 + \lambda Q(\mathbf{Dh}, \mathbf{Dv}; \alpha)$$

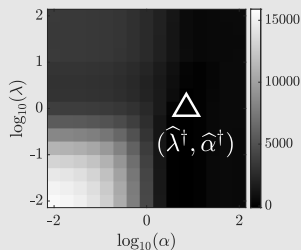
$\bar{\mathbf{h}}$ : true regularity

$$\mathcal{R}(\lambda, \alpha) = \left\| \hat{\mathbf{h}}(\mathcal{L}; \lambda, \alpha) - \bar{\mathbf{h}} \right\|^2$$



$\bar{\mathbf{h}}$ : unknown!

$$\hat{R}_{\nu, \epsilon}(\mathcal{L}; \lambda, \alpha | \mathcal{S})$$

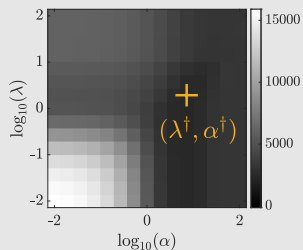


# Parameter tuning (Grid search)

$$\left(\hat{\mathbf{h}}, \hat{\mathbf{v}}\right) (\mathcal{L}; \lambda, \alpha) = \operatorname{argmin}_{\mathbf{h}, \mathbf{v}} \sum_a \left\| \log \mathcal{L}_{a,\cdot} - \log(a) \mathbf{h} - \mathbf{v} \right\|^2 + \lambda Q(\mathbf{Dh}, \mathbf{Dv}; \alpha)$$

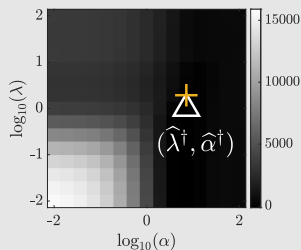
$\bar{\mathbf{h}}$ : true regularity

$$\mathcal{R}(\lambda, \alpha) = \left\| \hat{\mathbf{h}}(\mathcal{L}; \lambda, \alpha) - \bar{\mathbf{h}} \right\|^2$$



$\bar{\mathbf{h}}$ : unknown!

$$\hat{R}_{\nu, \epsilon}(\mathcal{L}; \lambda, \alpha | \mathcal{S})$$

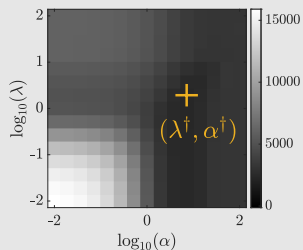


# Parameter tuning (Automatic selection)

$$\left(\widehat{\mathbf{h}}, \widehat{\mathbf{v}}\right) (\mathcal{L}; \lambda, \alpha) = \operatorname{argmin}_{\mathbf{h}, \mathbf{v}} \sum_a \left\| \log \mathcal{L}_{a,\cdot} - \log(a) \mathbf{h} - \mathbf{v} \right\|^2 + \lambda Q(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)$$

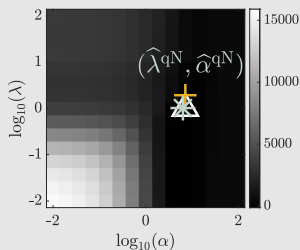
$\bar{\mathbf{h}}$ : true regularity

$$\mathcal{R}(\lambda, \alpha) = \left\| \widehat{\mathbf{h}}(\mathcal{L}; \lambda, \alpha) - \bar{\mathbf{h}} \right\|^2$$



$\bar{\mathbf{h}}$ : unknown!

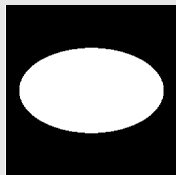
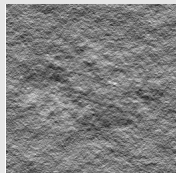
$$\widehat{R}_{\nu, \epsilon}(\mathcal{L}; \lambda, \alpha | \mathcal{S})$$



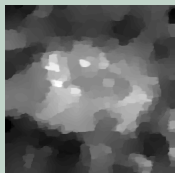
# Automated selection of regularization parameters

$$\left(\hat{\mathbf{h}}, \hat{\mathbf{v}}\right) (\mathcal{L}; \lambda, \alpha) = \operatorname{argmin}_{\mathbf{h}, \mathbf{v}} \sum_a \|\log \mathcal{L}_{a..} - \log(a)\mathbf{h} - \mathbf{v}\|^2 + \lambda \mathcal{Q}(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)$$

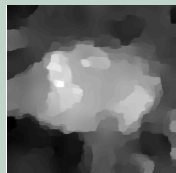
**Example**



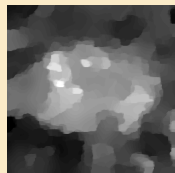
$\hat{\mathbf{h}}^F(\mathcal{L}; \lambda^\dagger, \alpha^\dagger)$   
(grid)



$\hat{\mathbf{h}}^F(\mathcal{L}; \hat{\lambda}^\dagger, \hat{\alpha}^\dagger)$   
(grid)



$\hat{\mathbf{h}}^F(\mathcal{L}; \hat{\lambda}^{\text{qN}}, \hat{\alpha}^{\text{qN}})$   
(quasi-Newton)



225 calls of the estimator over the grid v.s. 40 for quasi-Newton



### Take home messages

- ▶ Fractal texture model based on local *regularity* and *variance*
  - \* appropriate for real-world texture characterization
  - \* complementary attributes able to finely discriminate

## Take home messages

- ▶ Fractal texture model based on local *regularity* and *variance*
  - \* appropriate for real-world texture characterization
  - \* complementary attributes able to finely discriminate
  
- ▶ Simultaneous estimation and regularization
  - \* significant decrease of the estimation error
  - \* accurate and regular contours thanks to co-localized penalization

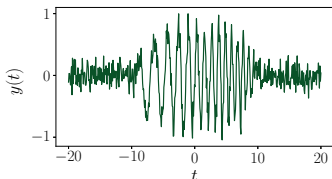
## Take home messages

- ▶ Fractal texture model based on local *regularity* and *variance*
  - \* appropriate for real-world texture characterization
  - \* complementary attributes able to finely discriminate
  
- ▶ Simultaneous estimation and regularization
  - \* significant decrease of the estimation error
  - \* accurate and regular contours thanks to co-localized penalization
  
- ▶ Fast algorithms for automated tuning of hyperparameters
  - \* possibility to manage huge amount of data
  - \* amenable to process data corrupted by *correlated* Gaussian noise
  - \* ensured objectivity and reproducibility

## **Part II:** Point processes in time-frequency analysis

# Time-frequency analysis of nonstationary signals

$y : \mathbb{R} \rightarrow \mathbb{C}$  function of time.



- electrical cardiac activity,
- audio recording,
- seismic activity,
- light intensity on a photosensor
- ...

## Information of interest:

- time events
- frequency content

e.g., an earthquake and its replica

e.g., monitoring of the heart beating rate

### **time**

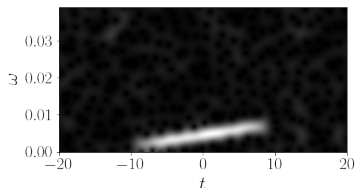
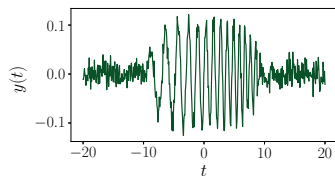
ever-changing world  
marker of events and evolutions

### **frequency**

waves, oscillations, rhythms  
intrinsic mechanisms

Short-Time Fourier Transform with window  $h$ :

$$V_h y(t, \omega) \triangleq \int_{-\infty}^{\infty} \overline{y(u)} h(u - t) \exp(-i\omega u) du$$



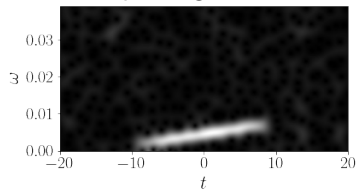
Energy density interpretation

$$\int \int_{-\infty}^{+\infty} |V_h y(t, \omega)|^2 dt \frac{d\omega}{2\pi} = \int_{-\infty}^{+\infty} |x(t)|^2 dt \quad \text{if} \quad \|h\|_2^2 = 1$$

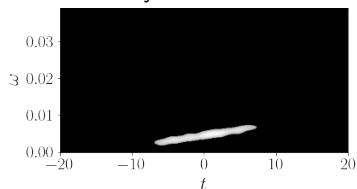
**Signal, i.e., information of interest: regions of maximal energy.**

Inversion formula 
$$y(t) = \int \int_{-\infty}^{+\infty} \overline{V_{hy}(u, \omega)} h(t - u) \exp(i\omega u) du \frac{d\omega}{2\pi}$$

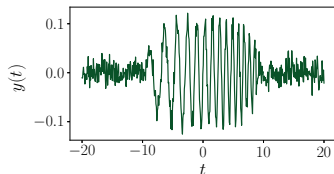
spectrogram



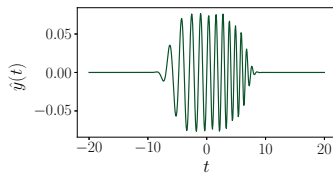
only maxima



noisy observation

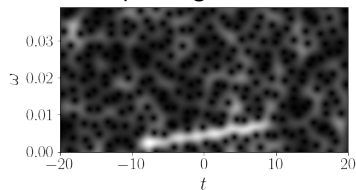


estimate

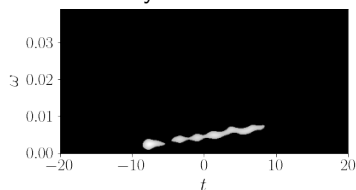


Inversion formula 
$$y(t) = \int \int_{-\infty}^{+\infty} \overline{V_h y(u, \omega)} h(t - u) \exp(i\omega u) du \frac{d\omega}{2\pi}$$

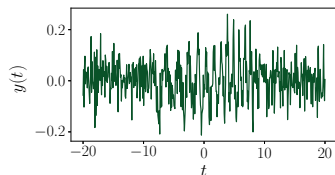
spectrogram



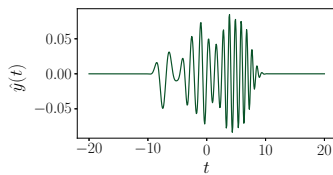
only maxima



noisy observation



estimate

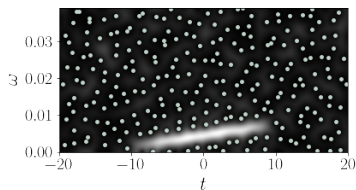
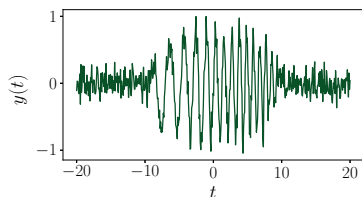


**Maxima detection:** *reassignment, synchrosqueezing, ridge extraction*



Restriction to the *circular Gaussian window*:  $g(t) = \pi^{-1/4} e^{-t^2/2}$

Look for the zeros, i.e., the points  $(t_i, \omega_i)$  such that  $|V_g y(t_i, \omega_i)|^2 = 0$ .

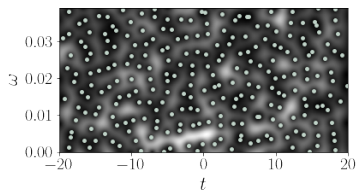
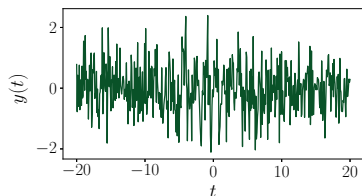


Observations: (Flandrin, 2015)

- In the noise region zeros are evenly spread.
- There exists a short-range repulsion between zeros.
- Zeros are repelled by the signal.

Restriction to the *circular Gaussian window*:  $g(t) = \pi^{-1/4} e^{-t^2/2}$

Look for the zeros, i.e., the points  $(t_i, \omega_i)$  such that  $|V_g y(t_i, \omega_i)|^2 = 0$ .

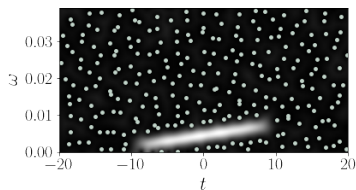
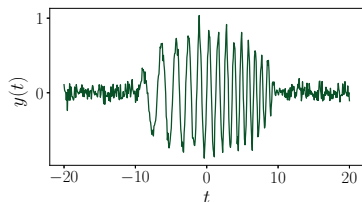


Observations: (Flandrin, 2015)

- In the noise region zeros are evenly spread.
- There exists a short-range repulsion between zeros.
- Zeros are repelled by the signal.

Restriction to the *circular Gaussian window*:  $g(t) = \pi^{-1/4} e^{-t^2/2}$

Look for the zeros, i.e., the points  $(t_i, \omega_i)$  such that  $|V_g y(t_i, \omega_i)|^2 = 0$ .

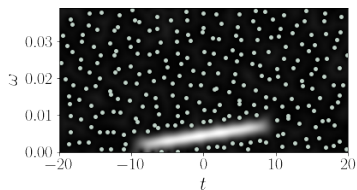
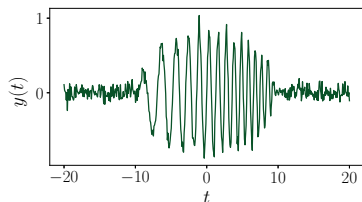


Observations: (Flandrin, 2015)

- In the noise region zeros are evenly spread.
- There exists a short-range repulsion between zeros.
- Zeros are repelled by the signal.

Restriction to the *circular Gaussian window*:  $g(t) = \pi^{-1/4} e^{-t^2/2}$

Look for the zeros, i.e., the points  $(t_i, \omega_i)$  such that  $|V_g y(t_i, \omega_i)|^2 = 0$ .



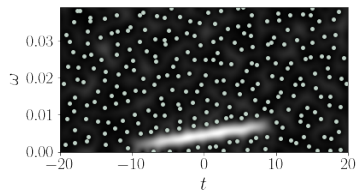
Observations: (Flandrin, 2015)

- In the noise region zeros are evenly spread.
- There exists a short-range repulsion between zeros.
- Zeros are repelled by the signal.

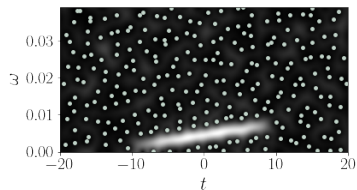
**What can be said theoretically about the zeros of the spectrogram?**

## Unorthodox time-frequency analysis: spectrogram zeros

**Idea** assimilate the time-frequency plane with  $\mathbb{C}$  through  $z = (\omega + it)/\sqrt{2}$



**Idea** assimilate the time-frequency plane with  $\mathbb{C}$  through  $z = (\omega + it)/\sqrt{2}$



Bargmann factorization

$$V_g y(t, \omega) = e^{-|z|^2/2} e^{-i\omega t/2} B_y(z)$$

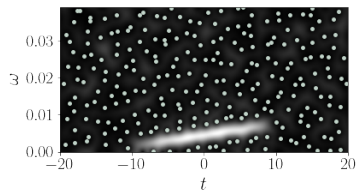
Bargmann transform of the signal  $y$

$$B_y(z) \triangleq \pi^{-1/4} e^{-z^2/2} \int_{\mathbb{R}} \overline{y(u)} \exp\left(\sqrt{2}uz - u^2/2\right) du,$$

$B_y$  is an **entire** function, almost characterized by its infinitely many zeros:

$$B_y(z) = z^m e^{C_0 + C_1 z + C_2 z^2} \prod_{n \in \mathbb{N}} \left(1 - \frac{z}{z_n}\right) \exp\left(\frac{z}{z_n} + \frac{1}{2} \left(\frac{z}{z_n}\right)^2\right).$$

**Idea** assimilate the time-frequency plane with  $\mathbb{C}$  through  $z = (\omega + it)/\sqrt{2}$



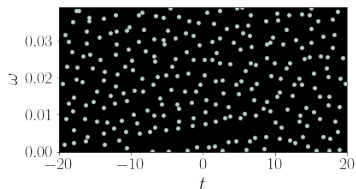
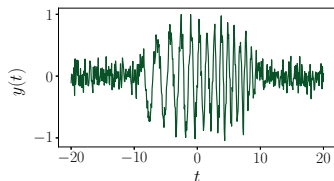
Bargmann factorization

$$V_g y(t, \omega) = e^{-|z|^2/2} e^{-i\omega t/2} B_y(z)$$

**Theorem** The zeros of the Gaussian spectrogram  $V_g y(t, \omega)$

- coincide with the zeros of the **entire** function  $B_y$ ,
- hence are **isolated** and constitute a **Point Process**,
- which almost completely **characterizes** the spectrogram.

(Flandrin, 2015)



## Advantages of working with the zeros

- Easy to find compared to relative maxima.
- Form a robust pattern in the time-frequency plane.
- Require little memory space for storage.
- Efficient tools were recently developed in **stochastic geometry**.

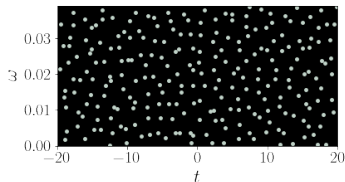
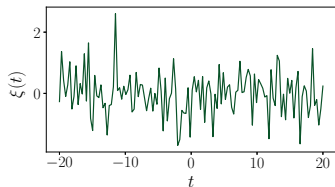
**Need for a rigorous characterization of the distribution of the zeros.**



# The zeros of the spectrogram of white noise

## Continuous complex white Gaussian noise

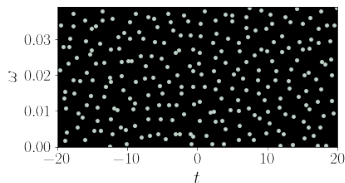
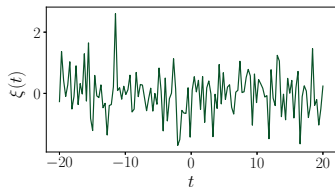
$$\xi(t) = \sum_{n=0}^{\infty} \xi[n] h_n(t), \quad \xi[n] \sim \mathcal{N}_{\mathbb{C}}(0, 1), \quad \{h_n, k = 0, 1, \dots\} \text{ Hermite functions}$$



# The zeros of the spectrogram of white noise

## Continuous complex white Gaussian noise

$$\xi(t) = \sum_{n=0}^{\infty} \xi[n] h_n(t), \quad \xi[n] \sim \mathcal{N}_{\mathbb{C}}(0, 1), \quad \{h_n, k = 0, 1, \dots\} \text{ Hermite functions}$$

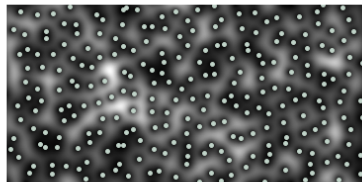


### Theorem

$$V_g \xi(t, \omega) = e^{-|z|^2/4} e^{-i\omega t/2} \text{GAF}_{\mathbb{C}}(z) \quad (\text{Bardenet \& Hardy, 2021})$$

$$\text{GAF}_{\mathbb{C}}(z) = \sum_{n=0}^{\infty} \xi[n] \frac{z^n}{\sqrt{n!}} \text{ is the planar Gaussian Analytic Function.}$$

# The zeros of the planar Gaussian Analytic Function



$$V_g \xi(t, \omega) \stackrel{\text{non-vanishing}}{\propto} \text{GAF}_{\mathbb{C}}(z)$$

$$z = (\omega + it)/\sqrt{2}$$

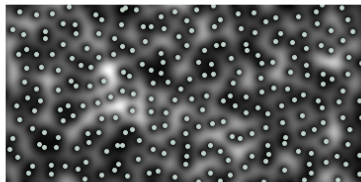
**Zeros of  $\text{GAF}_{\mathbb{C}}$ :** random set of points, i.e., a **Point Process** characterized by a probability distribution on point configurations

Properties of the Point Process of the zeros of  $\text{GAF}_{\mathbb{C}}$ :

- invariant under the isometries of  $\mathbb{C}$ , i.e., **stationary**,
- has a uniform density  $\rho^{(1)}(z) = \rho^{(1)} = 1/\pi$ ,
- explicit pair correlation function  $\rho^{(2)}(z, z') = g_0(|z - z'|)$ ,
- scaling of the *hole probability*:  $r^{-4} \log p_r \rightarrow -3e^2/4$ , as  $r \rightarrow \infty$

$$p_r = \mathbb{P}(\text{no point in the disk of center } 0 \text{ and radius } r)$$

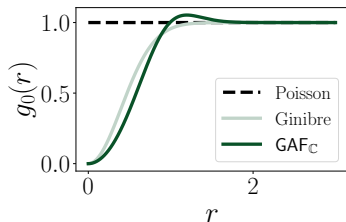
# The zeros of the planar Gaussian Analytic Function



$$V_g \xi(t, \omega) \stackrel{\text{non-vanishing}}{\propto} \text{GAF}_{\mathbb{C}}(z)$$

$$z = (\omega + it)/\sqrt{2}$$

**Zeros of  $\text{GAF}_{\mathbb{C}}$ :** random set of points, i.e., a **Point Process** characterized by a probability distribution on point configurations



Pair correlation  $\rho^{(2)}(z, z') dz dz' =$

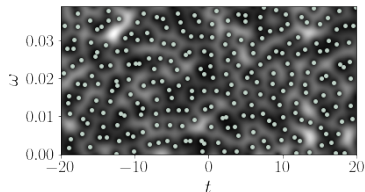
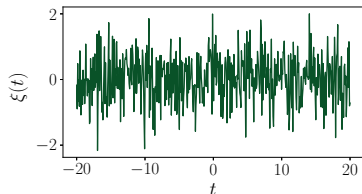
$$\mathbb{P}(\text{1 point in } B(z, dz) \text{ and 1 in } B(z', dz'))$$

The point process of the zeros of the spectrogram is not **determinantal**.

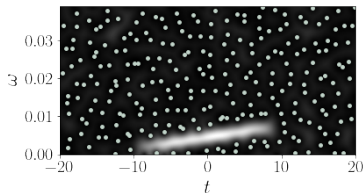
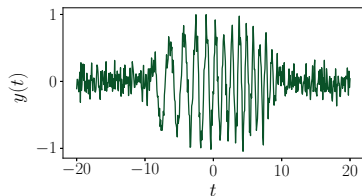
# Signal detection based on the spectrogram zeros

- $H_0$  white noisy only, i.e.,  $y(t) = \xi(t)$
- $H_1$  presence of a signal i.e.,  $y(t) = \text{snr} \times x(t) + \xi(t)$ ,  $\text{snr} > 0$

null hypothesis



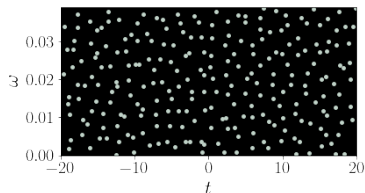
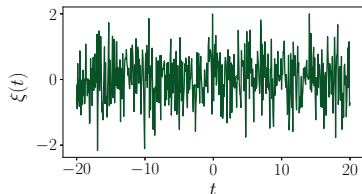
alternative hypothesis



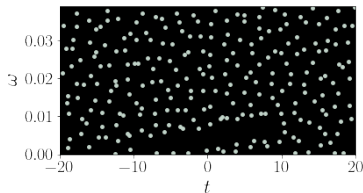
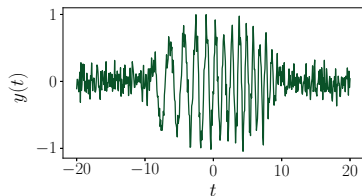
# Signal detection based on the spectrogram zeros

- $H_0$  white noisy only, i.e.,  $y(t) = \xi(t)$
- $H_1$  presence of a signal i.e.,  $y(t) = \text{snr} \times x(t) + \xi(t)$ ,  $\text{snr} > 0$

null hypothesis

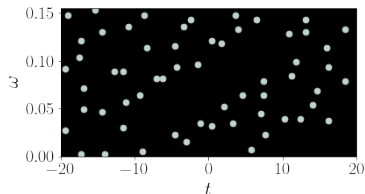
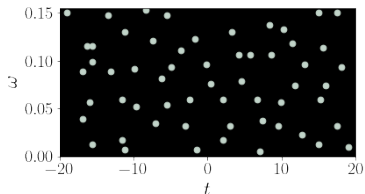


alternative hypothesis



A functional statistic: the  $F$ -function

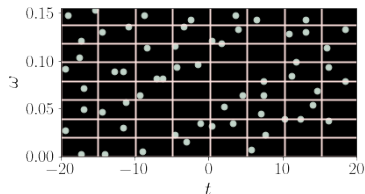
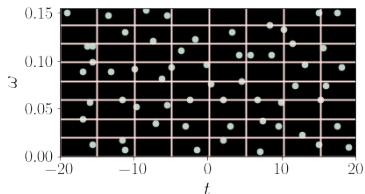
$$F(r) = \mathbb{P} \left( \inf_{z_i \in \mathcal{Z}} d(z_0, z_i) < r \right) : \text{empty space function}$$



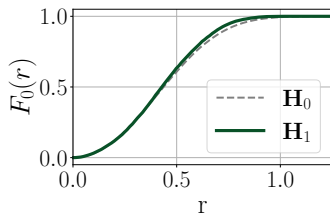
# Signal detection based on the spectrogram zeros

A functional statistic: the  $F$ -function

$$F(r) = \mathbb{P} \left( \inf_{z_i \in \mathcal{Z}} d(z_0, z_i) < r \right) : \text{empty space function}$$



$$\hat{F}(r) = \frac{1}{N_{\#}} \sum_{j=1}^{N_{\#}} \mathbf{1} \left( \inf_{z \in \text{Zeros}} d(z_j, z) < r \right)$$

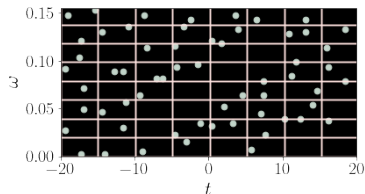
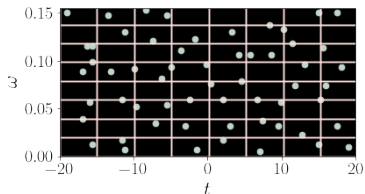




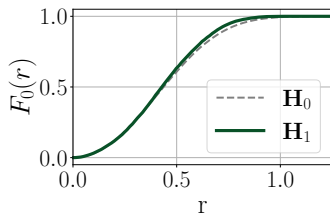
# Signal detection based on the spectrogram zeros

A functional statistic: the  $F$ -function

$$F(r) = \mathbb{P} \left( \inf_{z_i \in \mathcal{Z}} d(z_0, z_i) < r \right) : \text{empty space function}$$

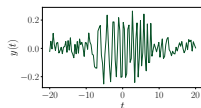


$$\hat{F}(r) = \frac{1}{N_{\#}} \sum_{j=1}^{N_{\#}} \mathbf{1} \left( \inf_{z \in \text{Zeros}} d(z_j, z) < r \right)$$

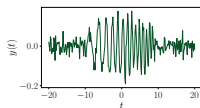


► Monte Carlo envelope test based on the discrepancy between  $\hat{F}$  and  $F_0$

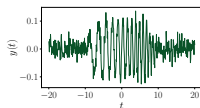
# Signal detection based on the spectrogram zeros



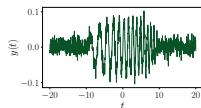
$N = 128$



$N = 256$

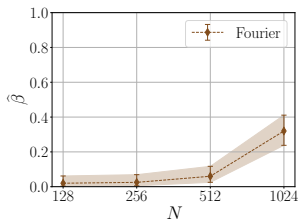


$N = 512$



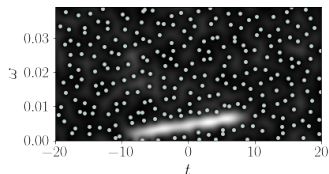
$N = 1024$

**Performance:** power of the test computed over 200 samples



✗ low detection power

✗ requires large number of samples

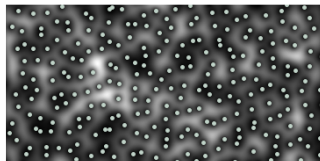


## Limitations:

- necessary discretization of the STFT: arbitrary resolution
- observe only a bounded window: edge correction to compute  $\hat{F}(r)$

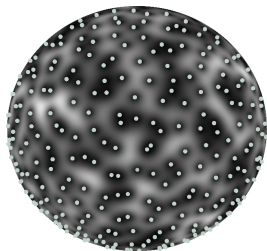
## Short-Time Fourier Transform

$$V_g \xi(t, \omega) \propto \text{GAF}_{\mathbb{C}}(z) = \sum_{n=0}^{\infty} \xi[n] \frac{z^n}{\sqrt{n!}}$$



## New transform?

$$? \propto \text{GAF}_{\mathbb{S}}(z) = \sum_{n=0}^N \xi[n] \sqrt{\binom{N}{n}} z^n$$



## Time and frequency shifts

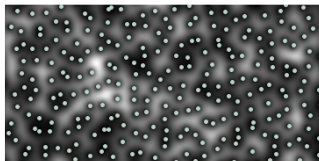
$$\mathbf{W}_{(t,\omega)}y(u) = e^{-i\omega u}y(u-t)$$

$$V_h[\mathbf{W}_{(t,\omega)}y](t', \omega') \stackrel{\text{covariance}}{=} e^{-i(\omega'-\omega)t} V_h y(t'-t, \omega'-\omega),$$

## Coherent state interpretation

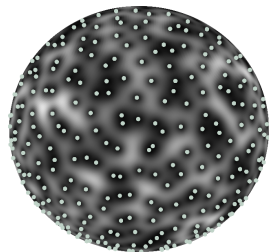
$$V_h y(t, \omega) = \langle y, \mathbf{W}_{(t,\omega)}h \rangle$$

$\{\mathbf{W}_{(t,\omega)}h, t, \omega \in \mathbb{R}\}$  covariant family



**Weyl-Heisenberg group**  $\{e^{i\gamma} \mathbf{W}_{(t,\omega)}, (\gamma, t, \omega) \in [0, 2\pi] \times \mathbb{R}^2\}$

$$\mathbf{W}_{(t',\omega')} \mathbf{W}_{(t,\omega)} = e^{i\omega t'} \mathbf{W}_{(t+t',\omega+\omega')}.$$

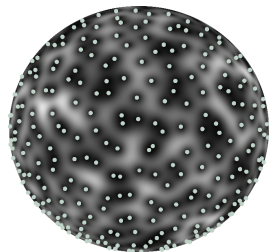


# The Kravchuk transform: covariance under $SO(3)$

Coherent state interpretation      $\mathbf{y} \in \mathbb{C}^{N+1}$

$$T\mathbf{y}(\vartheta, \varphi) = \langle \mathbf{y}, \Psi_{(\vartheta, \varphi)} \rangle$$

$$\vartheta \in [0, \pi], \varphi \in [0, 2\pi]$$

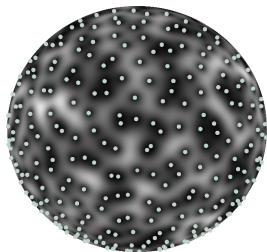


# The Kravchuk transform: covariance under SO(3)

Coherent state interpretation  $\mathbf{y} \in \mathbb{C}^{N+1}$

$$T\mathbf{y}(\vartheta, \varphi) = \langle \mathbf{y}, \Psi_{(\vartheta, \varphi)} \rangle$$

$$\vartheta \in [0, \pi], \varphi \in [0, 2\pi]$$



SO(3) **coherent states** (Gazeau, 2009)

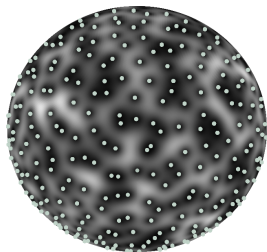
$$\Psi_{\vartheta, \varphi} = \sum_{n=0}^N \sqrt{\binom{N}{n}} \left( \cos \frac{\vartheta}{2} \right)^n \left( \sin \frac{\vartheta}{2} \right)^{N-n} e^{in\varphi} \mathbf{q}_n = \mathbf{R}_{\mathbf{u}(\vartheta, \varphi)} \Psi_{(0,0)},$$

# The Kravchuk transform: covariance under SO(3)

Coherent state interpretation  $\mathbf{y} \in \mathbb{C}^{N+1}$

$$T\mathbf{y}(\vartheta, \varphi) = \langle \mathbf{y}, \Psi_{(\vartheta, \varphi)} \rangle$$

$$\vartheta \in [0, \pi], \varphi \in [0, 2\pi]$$



SO(3) **coherent states** (Gazeau, 2009)

$$\Psi_{\vartheta, \varphi} = \sum_{n=0}^N \sqrt{\binom{N}{n}} \left(\cos \frac{\vartheta}{2}\right)^n \left(\sin \frac{\vartheta}{2}\right)^{N-n} e^{in\varphi} \mathbf{q}_n = \mathbf{R}_{\mathbf{u}(\vartheta, \varphi)} \Psi_{(0,0)},$$

**Kravchuk transform**

$\{\mathbf{q}_n, n = 0, 1, \dots, N\}$  the *Kravchuk functions*

$$T\mathbf{y}(z) = \frac{1}{\sqrt{(1+|z|^2)^N}} \sum_{n=0}^N \langle \mathbf{y}, \mathbf{q}_n \rangle \sqrt{\binom{N}{n}} z^n, \quad z = \cot(\vartheta/2)e^{i\varphi}$$

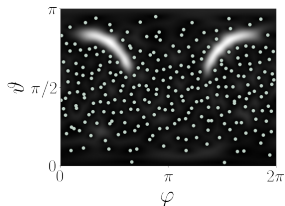
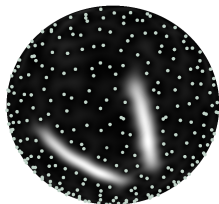
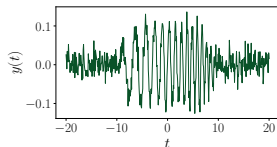


# The Kravchuk transform: covariance under SO(3)

**Kravchuk transform**

$\{\mathbf{q}_n, n = 0, 1, \dots, N\}$  the *Kravchuk functions*

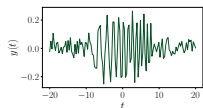
$$T\mathbf{y}(z) = \frac{1}{\sqrt{(1 + |z|^2)^N}} \sum_{n=0}^N \langle \mathbf{y}, \mathbf{q}_n \rangle \sqrt{\binom{N}{n}} z^n, \quad z = \cot(\vartheta/2) e^{i\varphi}$$



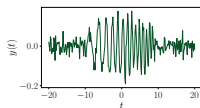
## Contributions

- rigorous link:  $T\xi(z) \stackrel{(\text{law})}{=} \sqrt{(1 + |z|^2)^{-N}} \text{GAF}_{\mathbb{S}}(z)$
- design of a robust implementation avoiding to compute  $\langle \mathbf{y}, \mathbf{q}_n \rangle$
- spatial statistics on the sphere using the chordal distance:  $\hat{F}(r)$

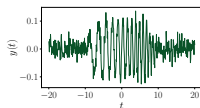
# Signal detection based on the spectrogram zeros



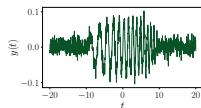
$N = 128$



$N = 256$

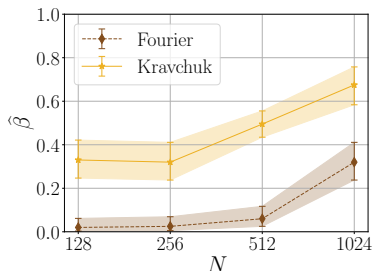


$N = 512$



$N = 1024$

**Performance:** power of the test computed over 200 samples



- ✓ higher detection power
- ✓ robust to small number of samples

- intrinsically encoded resolution: no need for prior knowledge
- compact phase space: no edge correction

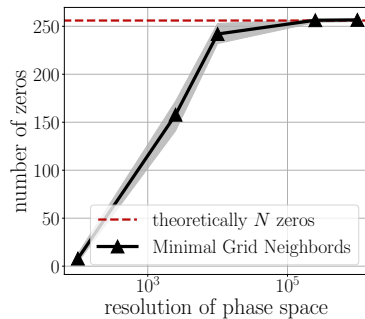
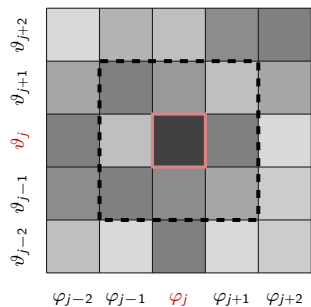
### Take home messages

- ▶ a novel covariant discrete transform
  - \* interpreted as a coherent state decomposition
  - \* zeros of the Kravchuk spectrogram of white noise fully characterized
- ▶ signal processing based on spectrogram zeros
  - \* preliminary work using the zeros of the Fourier spectrogram
  - \* significant improvement using the Kravchuk spectrogram

### Work in progress and perspectives

- ▶ convergence of the Kravchuk spectrogram toward the Fourier spectrogram
- ▶ interpretation of the action of  $SO(3)$  on  $\mathbb{C}^{N+1}$
- ▶ design of a FFT counterpart to compute the Kravchuk transform

# Detection of the zeros of the Kravchuk spectrogram



**Minimal Grid Neighbors**

**Purpose:** summary statistic  $s$ , such that  $\mathbb{E}[s(y)|\mathbf{H}_0] = 0$ ,  $\mathbb{E}[s(y)|\mathbf{H}_1] > 0$

## Test settings

- Level of significance  $\alpha$
- Number of samples under the null hypothesis  $m$
- Index  $k$ , chosen so that  $\alpha = k/(m + 1)$

## Monte Carlo strategy

- (i) generate  $m$  independent samples of complex white Gaussian noise;
- (ii) compute their summary statistics  $s_1 \geq s_2 \geq \dots \geq s_m$ ;
- (iii) compute the summary statistics of the observations  $\mathbf{y}$  under concern;
- (iv) if  $s(\mathbf{y}) \geq s_k$ , then reject the null hypothesis with confidence  $1 - \alpha$ .