



Texture segmentation based on fractal attributes using convex functional minimization with generalized Stein formalism for automated regularization parameter selection.

Barbara Pascal

November 27th 2020

Under the supervision of **Patrice Abry** and **Nelly Pustelnik**

Collaboration with **Valérie Vidal** and **Samuel Vaiter**

Image segmentation

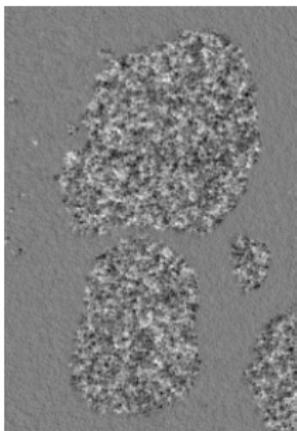
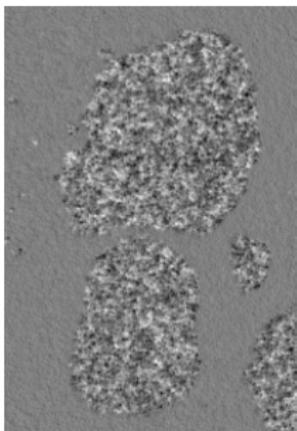


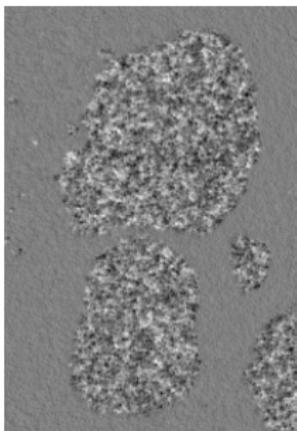
Image segmentation



Goal: obtain a partition of the image into K homogeneous regions

$$\Omega = \Omega_1 \sqcup \dots \sqcup \Omega_K$$

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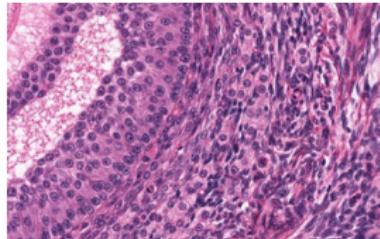
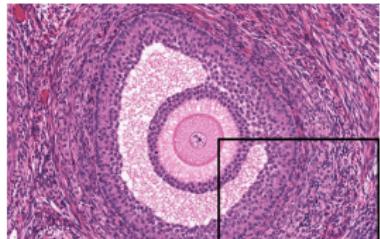
Textures



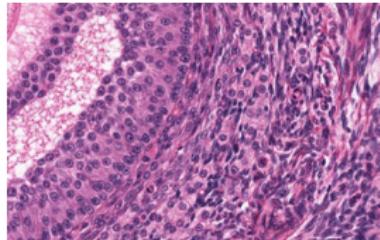
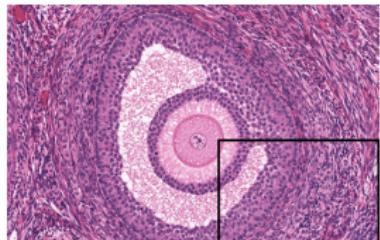
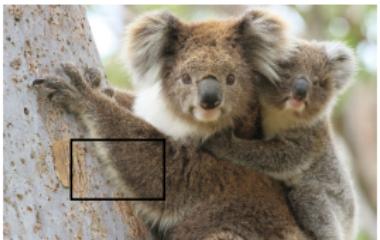
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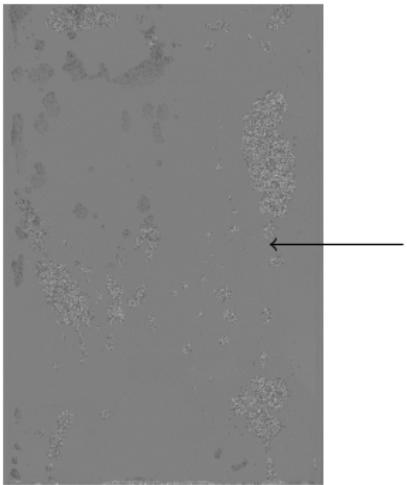
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Crucial to describe real-world images

Multiphase flow through porous media

Laboratoire de Physique, ENS Lyon, V. Vidal, T. Busser, (M. Serres, IFPEN)

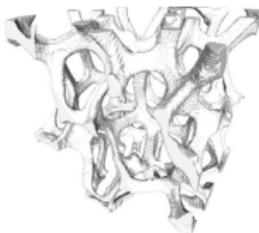
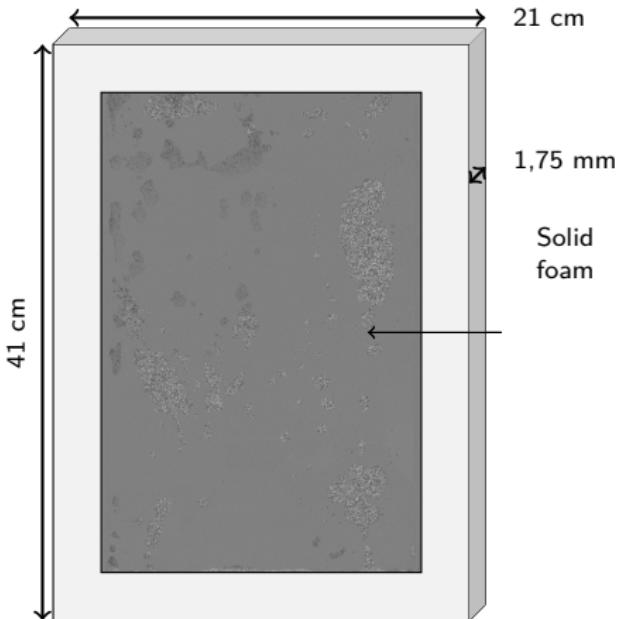


Solid
foam



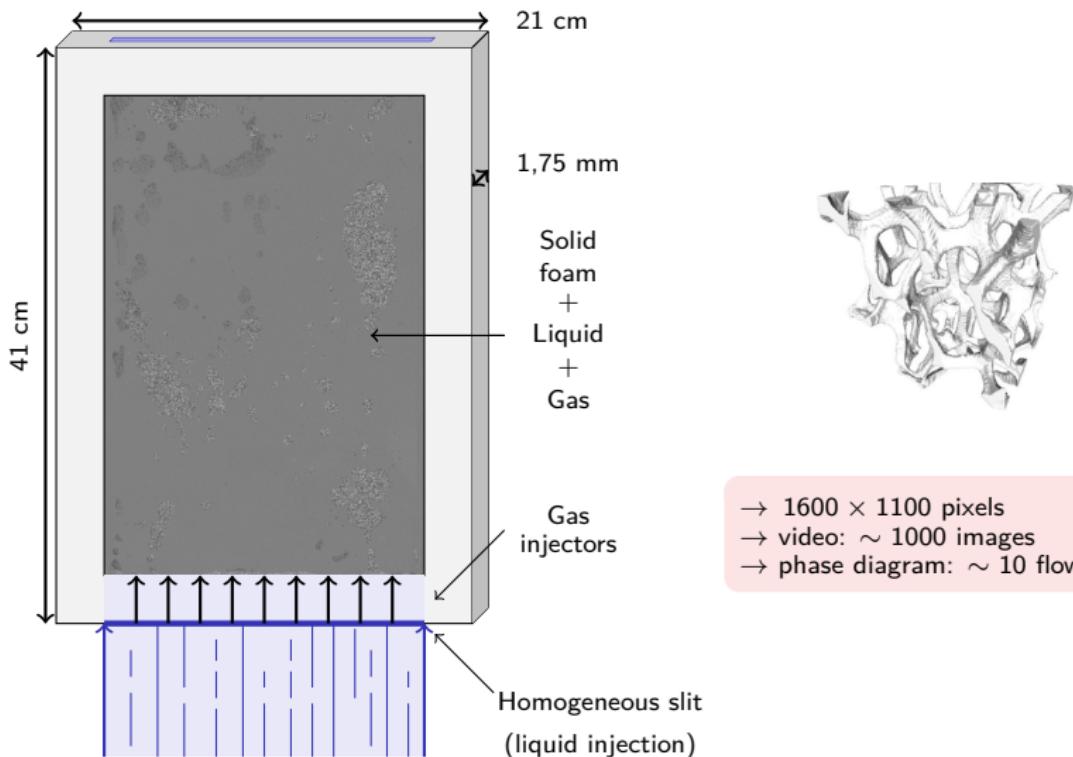
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Presentation outline

1. Texture characterization

[Gabor filters (Dunn, 1995)]

[Local amplitude and frequency (Havlicek, 1996)]

[Spectral histograms (Yuan, 2015)]

Presentation outline

1. Texture characterization

→ fractals attributes

- ▶ local variance σ^2
 - ▶ local regularity h

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2. Design of functionals

[Markov field (Geman, 1984)]

[Active contours (Chan, 2001)]

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3. Accelerated minimization algorithms

- [Forward-backward (Combettes, 2005)]
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- splitting proximal algorithms
 - ▶ computation of proximal operators [Forward-backward (Combettes, 2005)]
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4. Hyperparameters tuning

- [SURE (Stein, 1981)]
- [SURE FDMC (Ramani, 2008)]
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4. Hyperparameters tuning

- SURE under Gaussian correlated noise
 - ▶ projected estimation error
 - ▶ quasi-Newton minimization
 - ↪ Generalized SUGAR

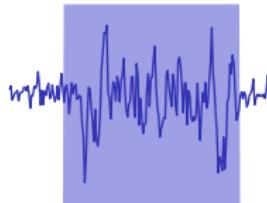
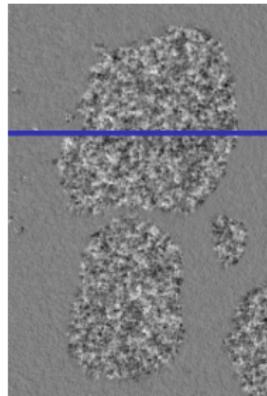
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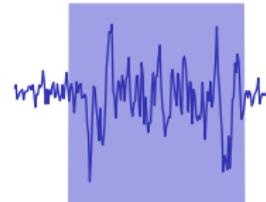
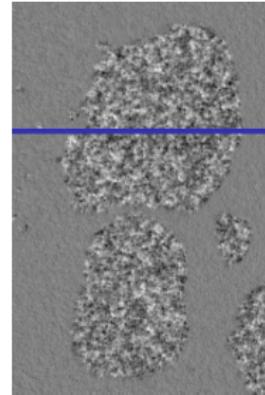
Piecewise monofractal model



Piecewise monofractal model

Fractals attributes

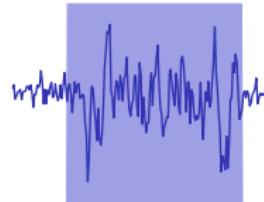
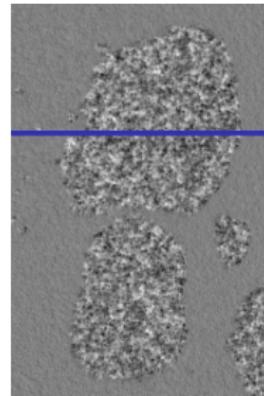
- variance σ^2 *amplitude of variations*



Piecewise monofractal model

Fractals attributes

- variance σ^2 *amplitude of variations*
 - local regularity h *scale invariance*

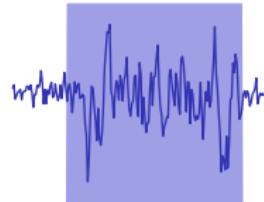
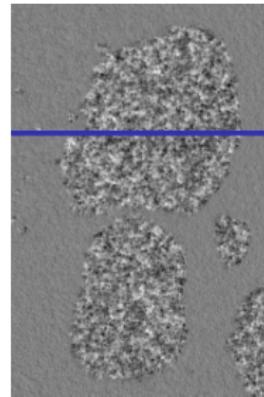


Piecewise monofractal model

Fractals attributes

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 - local regularity h *scale invariance*

$$|f(x) - f(y)| \leq \sigma(x)|x - y|^{h(x)}$$

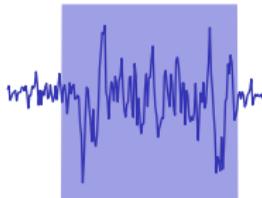
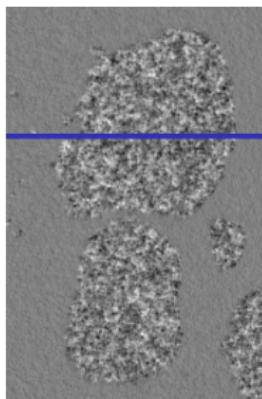
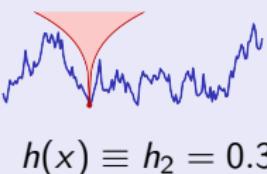
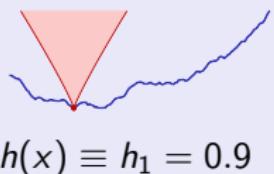


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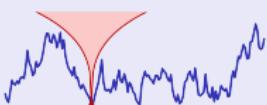
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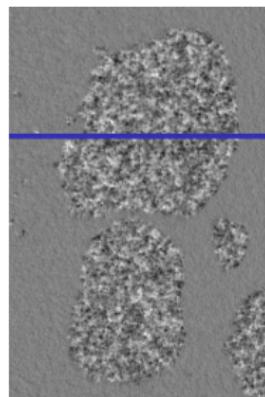
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$$h(x) \equiv h_1 = 0.9$$

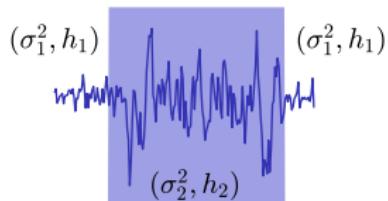


$$h(x) \equiv h_2 = 0.3$$



Segmentation

- ▶ σ^2 and h piecewise constant

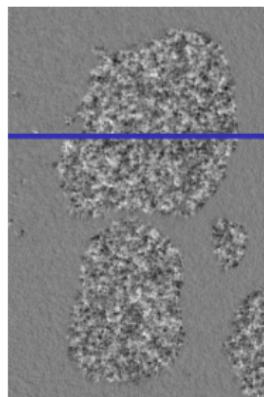
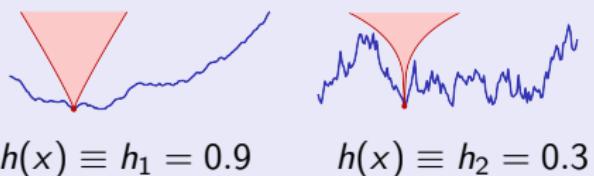


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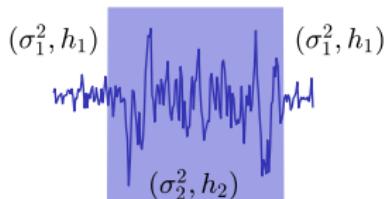
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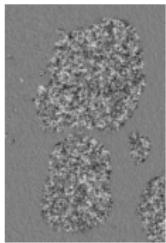
Segmentation

- ▶ σ^2 and h piecewise constant
 - ▶ region Ω_k characterized by (σ_k^2, h_k)



Multiscale analysis

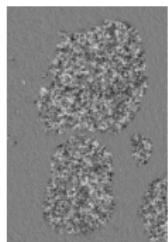
Textured image



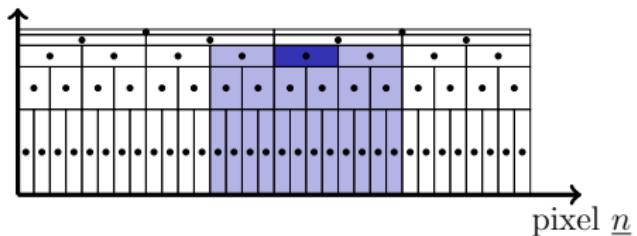
Multiscale analysis

Textured image

Local maximum of wavelet coefficients: $\mathcal{L}_{a,:}$



échelle *a*



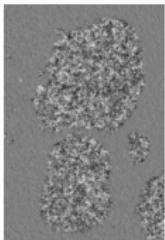
Multiscale analysis

Textured image

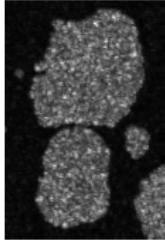
Local maximum of wavelet coefficients: $\mathcal{L}_{a,:}$

Scale

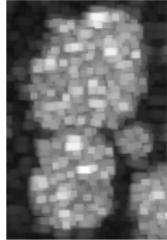
$$a = 2^1$$



$$a = 2^2$$

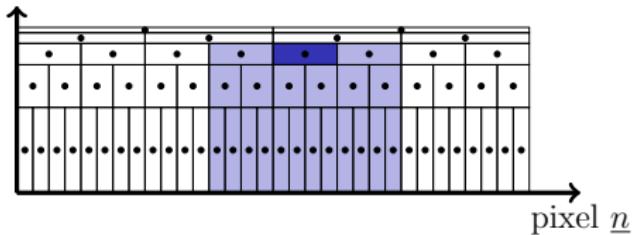


$$a = 2^5$$



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échelle *a*



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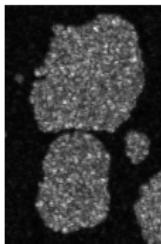
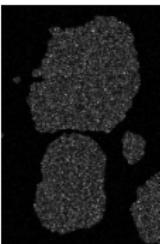
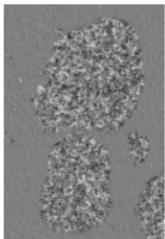
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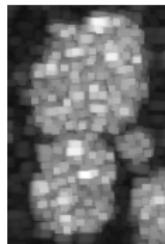
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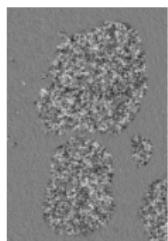
Proposition (*Jaffard, 2004*), (*Wendt, 2008*)

$$\log(\mathcal{L}_{a,\cdot}) \underset{a \rightarrow 0}{\simeq} \log(a) \frac{h}{\text{regularity}} + \frac{\nu}{\alpha \log(\sigma^2)} \text{(variance)}$$

Multiscale analysis

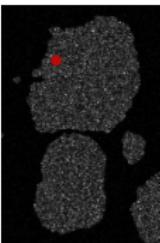
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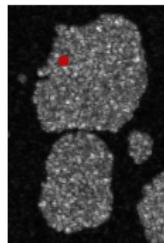


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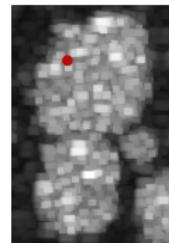


$$a = 2^2$$



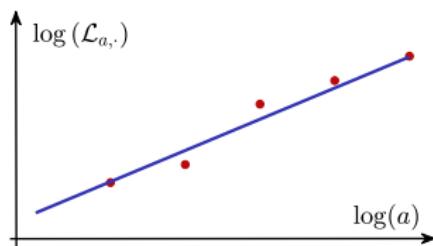
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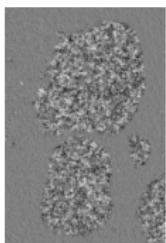
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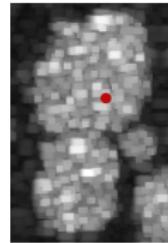


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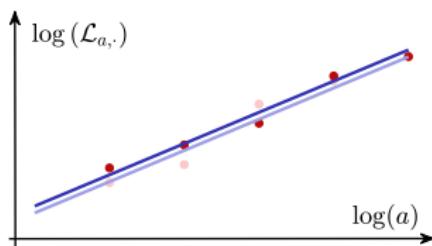
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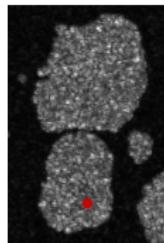
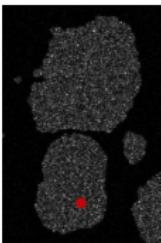
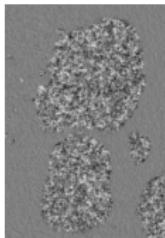
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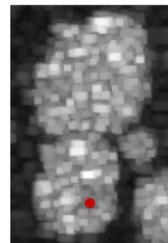
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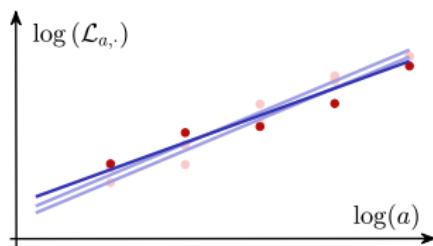


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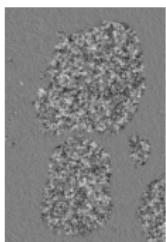
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Multiscale analysis

Textured image

Local maximum of wavelet coefficients: $\mathcal{L}_{a,:}$

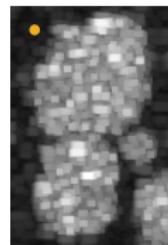


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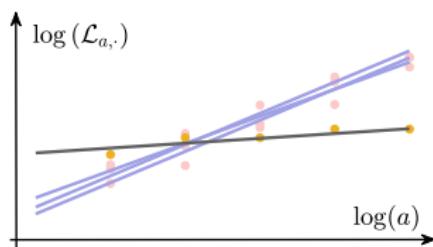
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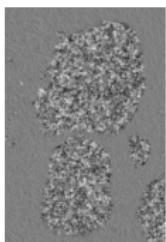
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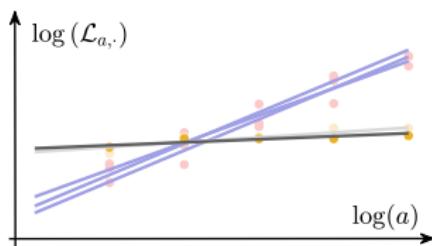
A fluorescence micrograph showing a cluster of cells with bright yellow fluorescence. The cells appear as irregular, somewhat rounded shapes with a granular internal structure. A small, distinct yellow spot is visible in the bottom left corner.

A fluorescence micrograph showing two large, granular cells in grayscale. A small yellow dot is visible in the bottom left corner.

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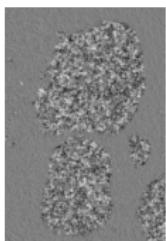
$$\log(\mathcal{L}_{a,\cdot}) \underset{a \rightarrow 0}{\simeq} \log(a) \frac{h}{\text{regularity}} + \frac{\nu}{\alpha \log(\sigma^2)} \text{(variance)}$$



Multiscale analysis

Textured image

Local maximum of wavelet coefficients: $\mathcal{L}_{a,:}$

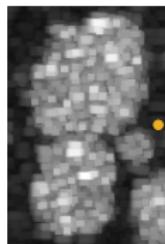


Scale

$$a = 2^1$$

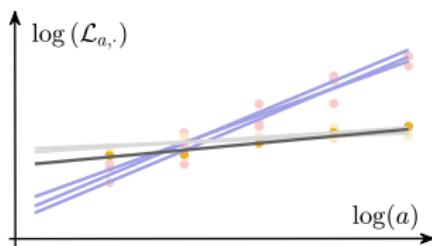
$$a = 2^2$$

• •



Proposition (*Jaffard, 2004*), (*Wendt, 2008*)

$$\log(\mathcal{L}_{a,\cdot}) \underset{a \rightarrow 0}{\simeq} \log(a) \frac{\boldsymbol{h}}{\text{regularity}} + \frac{\boldsymbol{v}}{\alpha \log(\sigma^2)}$$

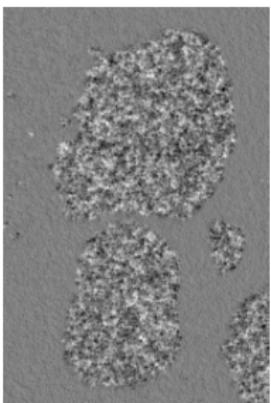


Direct punctual estimation

Linear regression

$$\log(\mathcal{L}_{a,\cdot}) \simeq \log(a) \frac{\boldsymbol{h}}{\text{regularity}} + \frac{\boldsymbol{v}}{\propto \log(\sigma^2)}$$

Textured image

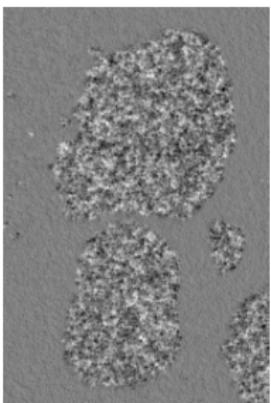


Direct punctual estimation

$$\text{Linear regression} \quad \log(\mathcal{L}_{a,\cdot}) \simeq \log(a) \frac{\mathbf{h}}{\text{regularity}} + \frac{\mathbf{v}}{\propto \log(\sigma^2)}$$

$$\left(\hat{\boldsymbol{h}}^{\text{LR}}, \hat{\boldsymbol{v}}^{\text{LR}}\right) = \underset{\boldsymbol{h}, \boldsymbol{v}}{\operatorname{argmin}} \sum_{a=a_{\min}}^{a_{\max}} \|\log(\mathcal{L}_{a,\cdot}) - \log(a)\boldsymbol{h} - \boldsymbol{v}\|^2$$

Textured image

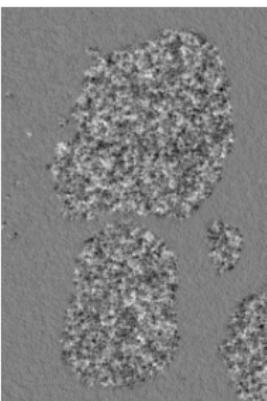


Direct punctual estimation

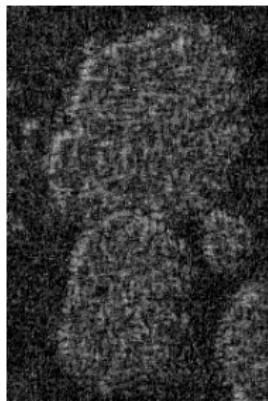
$$\text{Linear regression} \quad \log(\mathcal{L}_{a,\cdot}) \simeq \log(a) \frac{\mathbf{h}}{\text{regularity}} + \frac{\mathbf{v}}{\propto \log(\sigma^2)}$$

$$\left(\hat{\boldsymbol{h}}^{\text{LR}}, \hat{\boldsymbol{v}}^{\text{LR}}\right) = \underset{\boldsymbol{h}, \boldsymbol{v}}{\operatorname{argmin}} \sum_{a=a_{\min}}^{a_{\max}} \|\log(\mathcal{L}_{a,\cdot}) - \log(a)\boldsymbol{h} - \boldsymbol{v}\|^2$$

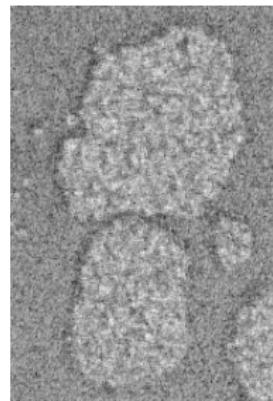
Textured image



Local regularity \hat{h}^{LR}



Local power \hat{v}^{LR}



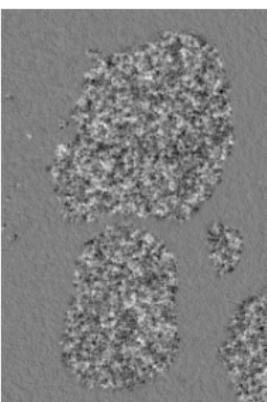
Direct punctual estimation

Linear regression

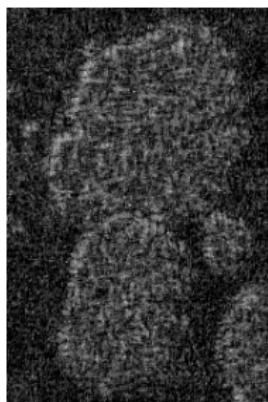
$$\mathbb{E} \log(\mathcal{L}_{a,\cdot}) = \log(a) \underbrace{\bar{h}}_{\text{regularity}} + \underbrace{\bar{v}}_{\propto \log(\sigma^2)}$$

$$\left(\hat{\boldsymbol{h}}^{\text{LR}}, \hat{\boldsymbol{v}}^{\text{LR}}\right) = \underset{\boldsymbol{h}, \boldsymbol{v}}{\operatorname{argmin}} \sum_{a=a_{\min}}^{a_{\max}} \|\log(\mathcal{L}_{a,\cdot}) - \log(a)\boldsymbol{h} - \boldsymbol{v}\|^2$$

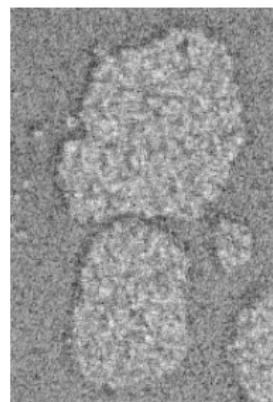
Textured image



Local regularity $\widehat{\boldsymbol{h}}^{\text{LR}}$



Local power $\hat{\nu}^{\text{LR}}$



→ large estimation variance

A posteriori regularization

Linear regression \hat{h}^{LR}



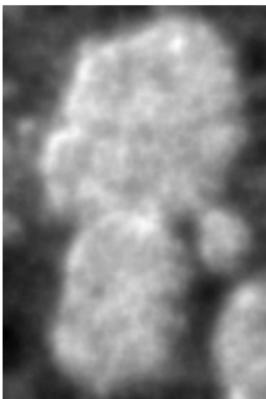
A posteriori regularization

Filter smoothing (linear)

$$\left(\mathbf{I} + \lambda \mathbf{D}^\top \mathbf{D}\right)^{-1} \hat{\mathbf{h}}^{\text{LR}}$$

Linear regression \hat{h}^{LR}

Lissage



A posteriori regularization

Filter smoothing (linear)

$$\left(\mathbf{I} + \lambda \mathbf{D}^\top \mathbf{D}\right)^{-1} \hat{\mathbf{h}}^{\text{LR}}$$

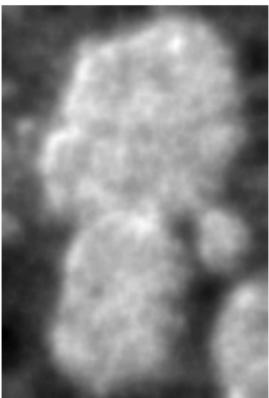
ROF denoising (nonlinear)

$$\operatorname{argmin}_{\boldsymbol{h}} \|\boldsymbol{h} - \hat{\boldsymbol{h}}^{\text{LR}}\|^2 + \lambda \|\mathbf{D}\boldsymbol{h}\|_{2,1}$$

Linear regression \hat{h}^{LR}



Lissage



ROF



A posteriori regularization

Filter smoothing (linear)

$$\left(\mathbf{I} + \lambda \mathbf{D}^\top \mathbf{D}\right)^{-1} \hat{\mathbf{h}}^{\text{LR}}$$

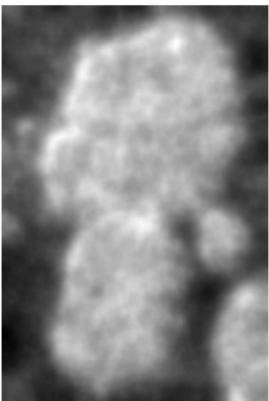
ROF denoising (nonlinear)

$$\operatorname{argmin}_{\boldsymbol{h}} \|\boldsymbol{h} - \hat{\boldsymbol{h}}^{\text{LR}}\|^2 + \lambda \|\mathbf{D}\boldsymbol{h}\|_{2,1}$$

Linear regression \hat{h}^{LR}



Lissage



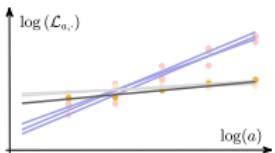
ROF



→ cumulative estimation variance and regularization bias

Functionals with either free or co-localized contours

$$\sum_a \frac{\|\log \mathcal{L}_{a,.} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}} \rightarrow \text{fidelity to the log-linear model}$$

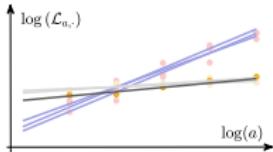


Functionals with either free or co-localized contours

$$\sum_a \frac{\|\log \mathcal{L}_{a,.} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}} + \lambda \frac{\mathcal{Q}(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)}{\text{Total Variation}}$$

\rightarrow fidelity to the log-linear model \rightarrow favors piecewise constancy

For more information about the study, please contact Dr. Michael J. Hwang at (319) 356-4530 or via email at mhwang@uiowa.edu.



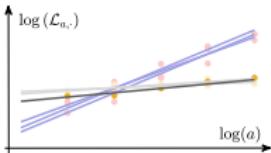
Total Variation
→ favors piecewise constancy

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Functionals with either free or co-localized contours

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \frac{\|\log \mathcal{L}_{a,.} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}} + \lambda \mathcal{Q}(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha) \quad \begin{array}{l} \text{Total Variation} \\ \rightarrow \text{favors piecewise constancy} \end{array}$$



Total Variation
→ favors piecewise constancy

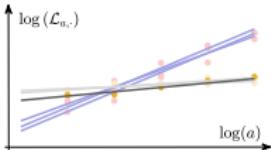
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Functionals with either free or co-localized contours

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \frac{\|\log \mathcal{L}_{a,.} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}} + \lambda \frac{\mathcal{Q}(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)}{\text{Total Variation}}$$

→ fidelity to the log-linear model → favors piecewise constancy



Total Variation
→ favors piecewise constancy

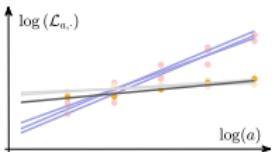
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Finite differences D_1x (horizontal), D_2x (vertical) in each pixel

Functionals with either free or co-localized contours

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \frac{\|\log \mathcal{L}_{a,.} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}} + \lambda \mathcal{Q}(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha) \quad \begin{array}{l} \text{Total Variation} \\ \rightarrow \text{favors piecewise constancy} \end{array}$$



Total Variation
→ favors piecewise constancy

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Finite differences $Dx = [D_1x, D_2x]$

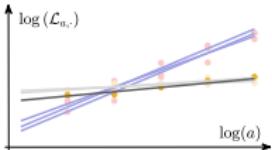
Free: h , v are **independently** piecewise constant

$$\mathcal{Q}_F(\mathbf{D}h, \mathbf{D}v; \alpha) = \alpha \|\mathbf{D}h\|_{2,1} + \|\mathbf{D}v\|_{2,1}$$

Functionals with either free or co-localized contours

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \frac{\|\log \mathcal{L}_{a,.} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}} \quad + \quad \lambda \frac{\mathcal{Q}(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)}{\text{Total Variation}}$$

→ fidelity to the log-linear model → favors piecewise constancy



Total Variation
→ favors piecewise constancy



Finite differences $Dx = [D_1x, D_2x]$

Free: h , v are **independently** piecewise constant

$$\mathcal{Q}_F(\mathbf{D}h, \mathbf{D}v; \alpha) = \alpha \|\mathbf{D}h\|_{2,1} + \|\mathbf{D}v\|_{2,1}$$

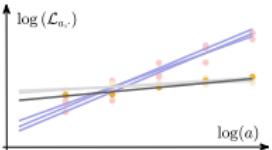
Co-localized: h , v are **concomitantly** piecewise constant

$$\mathcal{Q}_C(\mathbf{D}h, \mathbf{D}v; \alpha) = \|[\alpha \mathbf{D}h, \mathbf{D}v]\|_{2,1}$$

Functionals with either free or co-localized contours

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \frac{\|\log \mathcal{L}_{a,.} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}} + \lambda \frac{\mathcal{Q}(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)}{\text{Total Variation}}$$

→ fidelity to the log-linear model → favors piecewise constancy

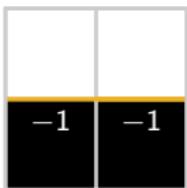


Total Variation
→ favors piecewise constancy

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Disjoint contours

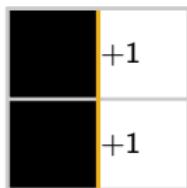


$$\boldsymbol{h} \in \mathbb{R}^{2 \times 2}$$

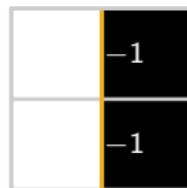


$$\mathbf{v} \in \mathbb{R}^{2 \times 2}$$

Common contours



$$\boldsymbol{h} \in \mathbb{R}^{2 \times 2}$$



$$\mathbf{v} \in \mathbb{R}^{2 \times 2}$$

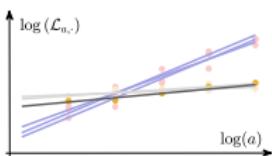
Functionals with either free or co-localized contours

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \frac{\|\log \mathcal{L}_{a,.} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}} + \lambda \frac{\mathcal{Q}(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)}{\text{Total Variation}}$$

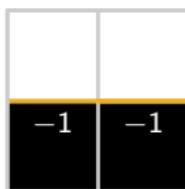
→ fidelity to the log-linear model → favors piecewise constancy

Least-Squares
fidelity to the log-linear model

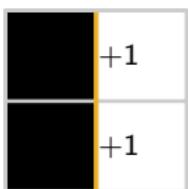
Total Variation
→ favors piecewise constancy



Disjoint contours

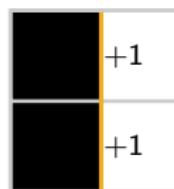


$$h \in \mathbb{R}^{2 \times 2}$$

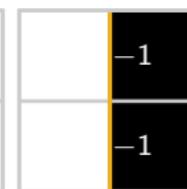


$$\mathbf{v} \in \mathbb{R}^{2 \times 2}$$

Common contours



$$h \in \mathbb{R}^{2 \times 2}$$



$$\mathbf{v} \in \mathbb{R}^{2 \times 2}$$

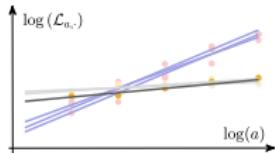
$$\mathcal{Q}_F(\mathbf{D}h, \mathbf{D}v; 1) = 4$$

$$\mathcal{Q}_F(\mathbf{D}h, \mathbf{D}v; 1) = 4$$

Functionals with either free or co-localized contours

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \frac{\|\log \mathcal{L}_{a,.} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}} \quad + \quad \lambda \frac{\mathcal{Q}(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)}{\text{Total Variation}}$$

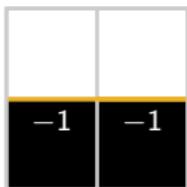
→ fidelity to the log-linear model → favors piecewise constancy



→ favors piecewise constancy



Disjoint contours

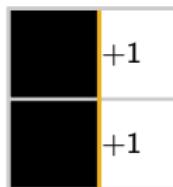


$$h \in \mathbb{R}^{2 \times 2}$$

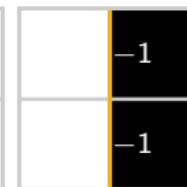


$$\mathbf{v} \in \mathbb{R}^{2 \times 2}$$

Common contours



$$h \in \mathbb{R}^{2 \times 2}$$



$$\mathbf{v} \in \mathbb{R}^{2 \times 2}$$

$$\mathcal{Q}_F(\mathbf{D}h, \mathbf{D}v; 1) = 4$$

$$\mathcal{Q}_C(\mathbf{D}h, \mathbf{D}v; 1) = 2 + \sqrt{2} \simeq 3.4$$

$$\mathcal{Q}_F(\mathbf{D}h, \mathbf{D}v; 1) = 4$$

$$\mathcal{Q}_C(\mathbf{D}h, \mathbf{D}v; 1) = 2\sqrt{2} \simeq 2.8$$

Functionals minimization

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \frac{\|\log \mathcal{L}_{a,.} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}} \quad + \quad \lambda \frac{\mathcal{Q}(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)}{\text{Total Variation}}$$



Functionals minimization

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \frac{\|\log \mathcal{L}_{a,.} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}} \quad + \quad \lambda \frac{\mathcal{Q}(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)}{\text{Total Variation}}$$



- gradient descent $\mathbf{x}^{n+1} = \mathbf{x}^n - \tau \nabla \varphi(\mathbf{x}^n)$

Functionals minimization

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \frac{\|\log \mathcal{L}_{a,.} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}} \quad + \quad \lambda \frac{\mathcal{Q}(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)}{\text{Total Variation}}$$



nonsmooth



- ▶ gradient descent $\mathbf{x}^{n+1} = \mathbf{x}^n - \tau \nabla \varphi(\mathbf{x}^n)$
 - ▶ implicit subgradient descent: proximal point algorithm

$$\mathbf{x}^{n+1} = \mathbf{x}^n - \tau \mathbf{u}^n, \quad \mathbf{u}^n \in \partial \varphi(\mathbf{x}^{n+1}) \iff \mathbf{x}^{n+1} = \text{prox}_{\tau \varphi}(\mathbf{x}^n)$$

Functionals minimization

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \frac{\|\log \mathcal{L}_{a,.} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}} + \lambda \frac{\mathcal{Q}(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)}{\text{Total Variation}}$$



nonsmooth



- ▶ gradient descent $\mathbf{x}^{n+1} = \mathbf{x}^n - \tau \nabla \varphi(\mathbf{x}^n)$
 - ▶ implicit subgradient descent: proximal point algorithm

$$\mathbf{x}^{n+1} = \mathbf{x}^n - \tau \mathbf{u}^n, \quad \mathbf{u}^n \in \partial \varphi(\mathbf{x}^{n+1}) \Leftrightarrow \mathbf{x}^{n+1} = \text{prox}_{\tau \varphi}(\mathbf{x}^n)$$
 - ▶ splitting proximal algorithm

$$\mathbf{y}^{n+1} = \text{prox}_{\sigma(\lambda \mathcal{Q})^*} (\mathbf{y}^n + \sigma \mathbf{D} \bar{\mathbf{x}}^n)$$

$$\mathbf{x}^{n+1} = \text{prox}_{\tau \|\mathcal{L} - \Phi \cdot\|_2^2} \left(\mathbf{x}^n - \tau \mathbf{D}^\top \mathbf{y}^{n+1} \right), \quad \Phi : (\mathbf{h}, \mathbf{v}) \mapsto \{\log(a) \mathbf{h} + \mathbf{v}\}_a$$

$$\bar{x}^{n+1} = 2x^{n+1} - x^n$$

Functionals minimization

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \frac{\|\log \mathcal{L}_{a,.} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}} \quad + \quad \lambda \frac{\mathcal{Q}(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)}{\text{Total Variation}}$$



nonsmooth



- ▶ gradient descent $\mathbf{x}^{n+1} = \mathbf{x}^n - \tau \nabla \varphi(\mathbf{x}^n)$
 - ▶ implicit subgradient descent: proximal point algorithm

$$\mathbf{x}^{n+1} = \mathbf{x}^n - \tau \mathbf{u}^n, \quad \mathbf{u}^n \in \partial \varphi(\mathbf{x}^{n+1}) \iff \mathbf{x}^{n+1} = \text{prox}_{\tau \varphi}(\mathbf{x}^n)$$

- splitting proximal algorithm

$$\text{prox}_{\tau\varphi}(\mathbf{x}) = \underset{\mathbf{u}}{\operatorname{argmin}} \frac{1}{2} \|\mathbf{x} - \mathbf{u}\|^2 + \tau\varphi(\mathbf{u})$$

$$\mathbf{y}^{n+1} = \text{prox}_{\sigma(\lambda\mathcal{Q})^*}(\mathbf{y}^n + \sigma\mathbf{D}\bar{\mathbf{x}}^n)$$

$$\mathbf{x}^{n+1} = \text{prox}_{\tau\|\mathcal{L}-\Phi\cdot\|_2^2} \left(\mathbf{x}^n - \tau\mathbf{D}^\top \mathbf{y}^{n+1} \right), \quad \Phi : (\mathbf{h}, \mathbf{v}) \mapsto \{\log(a)\mathbf{h} + \mathbf{v}\}_a$$

$$\bar{\mathbf{x}}^{n+1} \equiv 2\mathbf{x}^{n+1} - \mathbf{x}^n$$

Computation of proximal operators

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \frac{\|\log \mathcal{L}_{a,.} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}} \quad + \quad \lambda \frac{\mathcal{Q}(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)}{\text{Total Variation}}$$



nonsmooth



Computation of proximal operators

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \frac{\|\log \mathcal{L}_{a,.} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}} \quad + \quad \lambda \frac{\mathcal{Q}(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)}{\text{Total Variation}}$$



nonsmooth



Ex. Mixed norm: for $z = [z_1; \dots; z_l]$

$$\mathcal{Q}(\mathbf{z}) = \|\mathbf{z}\|_{2,1} = \sum_{n \in \Omega} \sqrt{\sum_{i=1}^l z_i^2(n)} = \sum_{n \in \Omega} \|\mathbf{z}(n)\|_2$$

Computation of proximal operators

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \frac{\|\log \mathcal{L}_{a,.} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}} \quad + \quad \lambda \frac{\mathcal{Q}(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)}{\text{Total Variation}}$$



nonsmooth



Ex. Mixed norm: for $z = [z_1; \dots; z_l]$

$$\mathcal{Q}(\mathbf{z}) = \|\mathbf{z}\|_{2,1} = \sum_{n \in \Omega} \sqrt{\sum_{i=1}^I z_i^2(n)} = \sum_{n \in \Omega} \|\mathbf{z}(n)\|_2$$

$$\boldsymbol{p} = \text{prox}_{\lambda \|\cdot\|_{2,1}}(\boldsymbol{z}) \quad \Leftrightarrow \quad p_i(\underline{n}) = \max \left(0, 1 - \frac{\lambda}{\|\boldsymbol{z}(\underline{n})\|_2} \right) z_i(\underline{n})$$

Computation of proximal operators

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \frac{\|\log \mathcal{L}_{a,.} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}} \quad + \quad \lambda \frac{\mathcal{Q}(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)}{\text{Total Variation}}$$



nonsmooth



Least-Squares: $\|\log \mathcal{L} - \Phi(\mathbf{h}, \mathbf{v})\|^2$, $\Phi : (\mathbf{h}, \mathbf{v}) \mapsto \{\log(a)\mathbf{h} + \mathbf{v}\}_a$

Computation of proximal operators

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \frac{\|\log \mathcal{L}_{a,.} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}} \quad + \quad \lambda \frac{\mathcal{Q}(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)}{\text{Total Variation}}$$



nonsmooth



Least-Squares: $\|\log \mathcal{L} - \Phi(\mathbf{h}, \mathbf{v})\|^2$, $\Phi : (\mathbf{h}, \mathbf{v}) \mapsto \{\log(a)\mathbf{h} + \mathbf{v}\}_a$

Proposition (*Pascal, 2019*)

$$(\tilde{\mathbf{h}}, \tilde{\mathbf{v}}) = \text{prox}_{\tau \|\mathcal{L} - \Phi \cdot\|_2^2}(\mathbf{h}, \mathbf{v}) \iff (\tilde{\mathbf{h}}, \tilde{\mathbf{v}}) = (\mathbf{I} + \tau \Phi^\top \Phi)^{-1} ((\mathbf{h}, \mathbf{v}) + \tau \Phi^\top \log \mathcal{L})$$

Computation of proximal operators

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \frac{\|\log \mathcal{L}_{a,.} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}} \quad + \quad \lambda \frac{\mathcal{Q}(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)}{\text{Total Variation}}$$



nonsmooth



Least-Squares: $\|\log \mathcal{L} - \Phi(\mathbf{h}, \mathbf{v})\|^2$, $\Phi : (\mathbf{h}, \mathbf{v}) \mapsto \{\log(a)\mathbf{h} + \mathbf{v}\}_a$

Proposition (*Pascal, 2019*)

Let $S_m = \sum_a \log^m(a)$, $\mathcal{D} = (1 + \tau S_2)(1 + \tau S_0) - \tau^2 S_1^2$,
 $\mathcal{T} = \sum_a \log \mathcal{L}_a$ and $\mathcal{G} = \sum_a \log(a) \log \mathcal{L}_a$, alors

$$\begin{aligned} (\tilde{\mathbf{h}}, \tilde{\mathbf{v}}) = & \text{prox}_{\tau \|\mathcal{L} - \Phi \cdot\|_2^2}(\mathbf{h}, \mathbf{v}) \iff (\tilde{\mathbf{h}}, \tilde{\mathbf{v}}) = (\mathbf{I} + \tau \Phi^\top \Phi)^{-1} ((\mathbf{h}, \mathbf{v}) + \tau \Phi^\top \log \mathcal{L}) \\ \iff & \begin{cases} \tilde{\mathbf{h}} = \mathcal{D}^{-1} ((1 + \tau S_0)(\tau \mathcal{G} + \mathbf{h}) - \tau S_1(\tau \mathcal{T} + \mathbf{v})) \\ \tilde{\mathbf{v}} = \mathcal{D}^{-1} ((1 + \tau S_2)(\tau \mathcal{T} + \mathbf{v}) - \tau S_1(\tau \mathcal{G} + \mathbf{h})) \end{cases} \end{aligned}$$

Accelerated algorithm based on strong-convexity

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \frac{\|\log \mathcal{L}_{a,.} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}} + \lambda \frac{\mathcal{Q}(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)}{\text{Total Variation}}$$



Primal-dual algorithm (*Chambolle, 2011*)



δ : duality gap, $\delta(\mathbf{x}^n, \mathbf{y}^n) \xrightarrow{n \rightarrow \infty} 0$

Accelerated algorithm based on strong-convexity

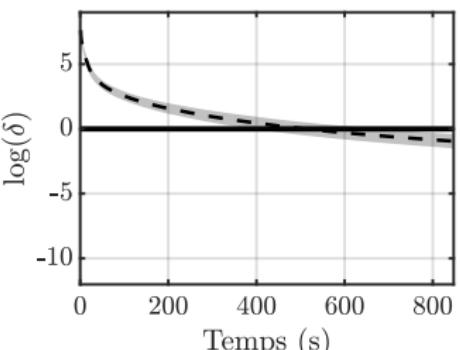
$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \frac{\|\log \mathcal{L}_{a,.} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}} \quad + \quad \lambda \frac{\mathcal{Q}(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)}{\text{Total Variation}}$$

nonsmooth



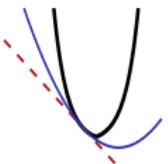
Primal-dual algorithm (*Chambolle, 2011*)

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Convexity properties

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \frac{\|\log \mathcal{L}_{a,.} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}} \quad + \quad \lambda \frac{\mathcal{Q}(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)}{\text{Total Variation}}$$



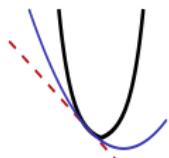
μ -strongly convex

nonsmooth



Convexity properties

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \frac{\|\log \mathcal{L}_{a,.} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}} \quad + \quad \lambda \frac{\mathcal{Q}(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)}{\text{Total Variation}}$$

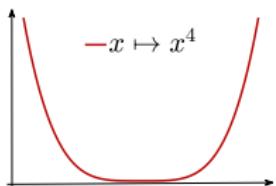


μ -strongly convex

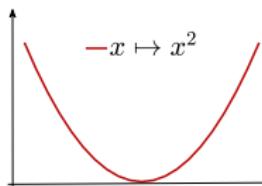


Strong-convexity

- φ μ -strongly convex iff $\varphi - \frac{\mu}{2} \|\cdot\|^2$ convex



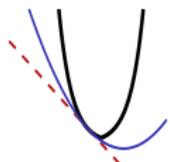
- ✓ strictly convex
- ✗ non strongly convex



- ✓ strictly convex
- ✓ 1-strongly convex

Convexity properties

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \frac{\|\log \mathcal{L}_{a,.} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}} \quad + \quad \lambda \frac{\mathcal{Q}(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)}{\text{Total Variation}}$$



μ -strongly convex

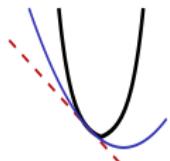


Strong-convexity

- φ μ -strongly convex iff $\varphi - \frac{\mu}{2} \|\cdot\|^2$ convex
 - $\varphi \in \mathcal{C}^2$ with Hessian matrix $H\varphi \succeq 0 \implies \mu = \min \text{Sp}(H\varphi)$

Convexity properties

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \frac{\|\log \mathcal{L}_{a,.} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}} \quad + \quad \lambda \frac{\mathcal{Q}(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)}{\text{Total Variation}}$$



μ -strongly convex

nonsmooth



Strong-convexity

- φ μ -strongly convex iff $\varphi - \frac{\mu}{2} \|\cdot\|^2$ convex
 - $\varphi \in \mathcal{C}^2$ with Hessian matrix $H\varphi \succeq 0 \implies \mu = \min \text{Sp}(H\varphi)$

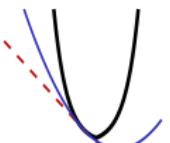
Proposition (Pascal, 2019)

$\sum_a \|\log \mathcal{L} - \log(a) \mathbf{h} - \mathbf{v}\|^2$ est μ -strongly convex.

$a_{\min} = 2^1$, a_{\max}	2 ²	2 ³	2 ⁴	2 ⁵	2 ⁶
$\mu = \min \text{Sp} (2\Phi^\top \Phi)$	0.29	0.72	1.20	1.69	2.20

Accelerated algorithm based on strong-convexity

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \frac{\|\log \mathcal{L}_{a,.} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}} \quad + \quad \lambda \frac{\mathcal{Q}(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)}{\text{Total Variation}}$$



μ -strongly convex



nonsmooth

Accelerated Primal-dual algorithm (*Chambolle, 2011*)

for $n = 0, 1, \dots$

$$x = (h, v)$$

$$\mathbf{y}^{n+1} = \text{prox}_{\sigma_n(\lambda \mathcal{Q})^*} (\mathbf{y}^n + \sigma_n \mathbf{D} \bar{\mathbf{x}}^n)$$

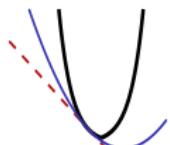
$$\mathbf{x}^{n+1} = \text{prox}_{\tau_n \|\mathcal{L} - \Phi \cdot\|_2^2} \left(\mathbf{x}^n - \tau_n \mathbf{D}^\top \mathbf{y}^{n+1} \right)$$

$$\theta_n = \sqrt{1 + 2\mu\tau_n}, \quad \tau_{n+1} = \tau_n/\theta_n, \quad \sigma_{n+1} = \theta_n \sigma_n$$

$$\bar{\mathbf{x}}^{n+1} = \mathbf{x}^{n+1} + \theta_n^{-1} (\mathbf{x}^{n+1} - \mathbf{x}^n)$$

Algorithme accéléré par forte-convexité

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \frac{\|\log \mathcal{L}_{a,.} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}} + \lambda \frac{\mathcal{Q}(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)}{\text{Total Variation}}$$



μ -strongly convex

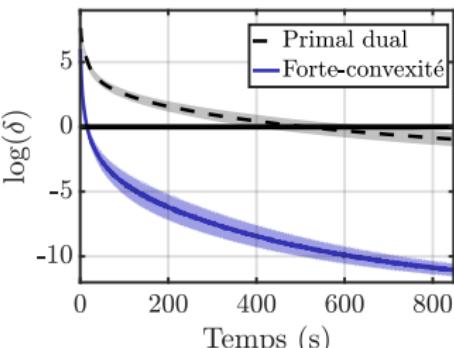


nonsmooth



Accelerated Primal-dual algorithm (*Chambolle, 2011*)

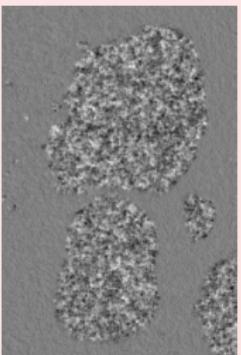
$$\delta: \text{duality gap, } \delta(\mathbf{x}^n, \mathbf{y}^n) \xrightarrow[n \rightarrow \infty]{} 0$$



Segmentation *via* iterated thresholding

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \frac{\|\log \mathcal{L}_{a,.} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}} + \lambda \frac{\mathcal{Q}(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)}{\text{Total Variation}}$$

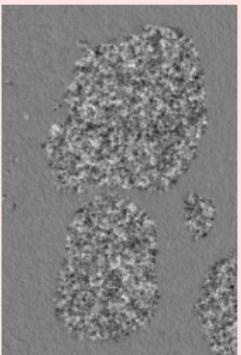
Textured image



Segmentation *via* iterated thresholding

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \frac{\|\log \mathcal{L}_{a,.} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}} + \lambda \frac{\mathcal{Q}(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)}{\text{Total Variation}}$$

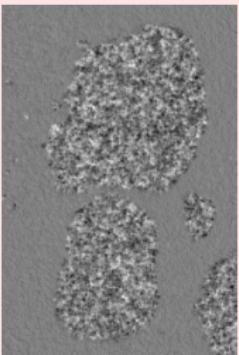
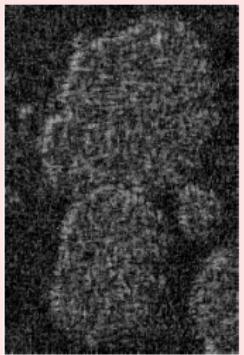
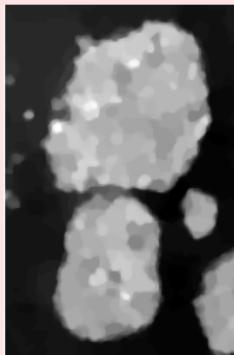
Textured image Lin. reg. $\hat{\mathbf{h}}^{\text{LR}}$



Segmentation via iterated thresholding

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \frac{\|\log \mathcal{L}_{a,.} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}} + \lambda \frac{\mathcal{Q}(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)}{\text{Total Variation}}$$

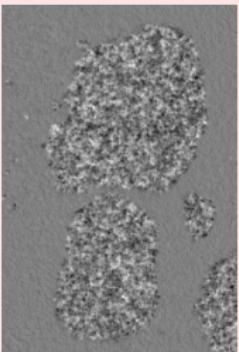
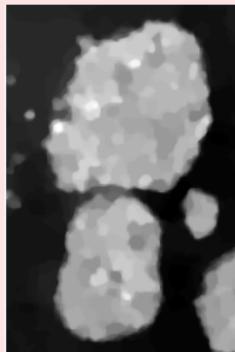
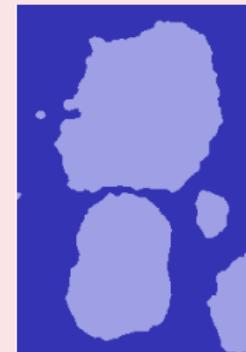
Textured image

Lin. reg. $\hat{\mathbf{h}}^{\text{LR}}$ Co-localized
contours $\hat{\mathbf{h}}^C$ 

Segmentation via iterated thresholding

$$\underset{\mathbf{h}, \mathbf{v}}{\text{minimize}} \quad \sum_a \frac{\|\log \mathcal{L}_{a,.} - \log(a)\mathbf{h} - \mathbf{v}\|^2}{\text{Least-Squares}} + \lambda \frac{\mathcal{Q}(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)}{\text{Total Variation}}$$

Textured image

Lin. reg. $\hat{\mathbf{h}}^{\text{LR}}$ Co-localized
contours $\hat{\mathbf{h}}^C$ Threshold
estimate[†] $T\hat{\mathbf{h}}^C$ [†](Cai, 2013)

State-of-the-art methods for texture segmentation

Threshold-ROF on \hat{h}^{LR}

(Naftornita, 2014), (Pustelnik, 2016)

$$\operatorname{argmin}_{\boldsymbol{h}} \|\boldsymbol{h} - \hat{\boldsymbol{h}}^{\text{LR}}\|^2 + \lambda \|\mathbf{D}\boldsymbol{h}\|_{2,1}$$

Lin. reg.



ROE



Threshold



Only based on regularity h .

State-of-the-art methods for texture segmentation

Threshold-ROF on \hat{h}^{LR}

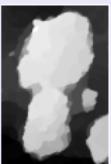
(Naftornita, 2014), (Pustelnik, 2016)

$$\operatorname{argmin}_{\boldsymbol{h}} \|\boldsymbol{h} - \hat{\boldsymbol{h}}^{\text{LR}}\|^2 + \lambda \|\mathbf{D}\boldsymbol{h}\|_{2,1}$$

Lin. reg.



ROE



Threshold



Only based on regularity h .

Factorization based segmentation[†] (Yuan, 2015)

- (i) local histograms



- (ii) matrix factorization

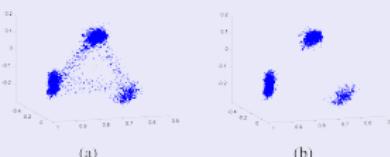


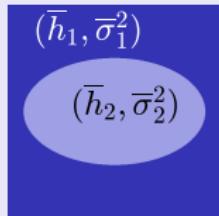
Fig. 2. Scatterplot of features in subspace. (a) Scatterplot of features projected onto the 3-d subspace. (b) Scatterplot after removing features with high edgeless.

[†]<https://sites.google.com/site/factorizationsegmentation/>

Compared segmentation performance on synthetic textures

Piecewise monofractal texture synthesis (*Pascal*, 2019)

- ▶ mask: $\Omega = \Omega_1 \sqcup \Omega_2$,
 - ▶ attributes: $(\bar{h}_k, \bar{\sigma}_k^2)_{k=1,2}$



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- ▶ mask: $\Omega = \Omega_1 \sqcup \Omega_2$,
 - ▶ attributes: $(\bar{h}_k, \bar{\sigma}_k^2)_{k=1,2}$

Ex. $\bar{h}_1 = 0.5, \bar{\sigma}_1^2 = 0.6$
 $\bar{h}_2 = 0.6, \bar{\sigma}_2^2 = 0.7$

$$(\overline{h}_1, \overline{\sigma}_1^2)$$

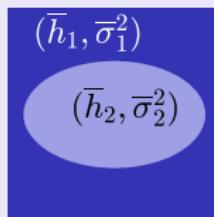


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 $\bar{h}_2 = 0.6, \bar{\sigma}_2^2 = 0.7$



Averaged segmentation performances over 5 realizations

Yuan



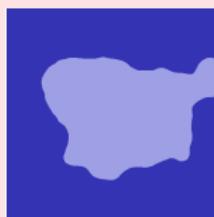
$71.1 \pm 1.3\%$

T-ROF



$78.5 \pm 1.1\%$

free



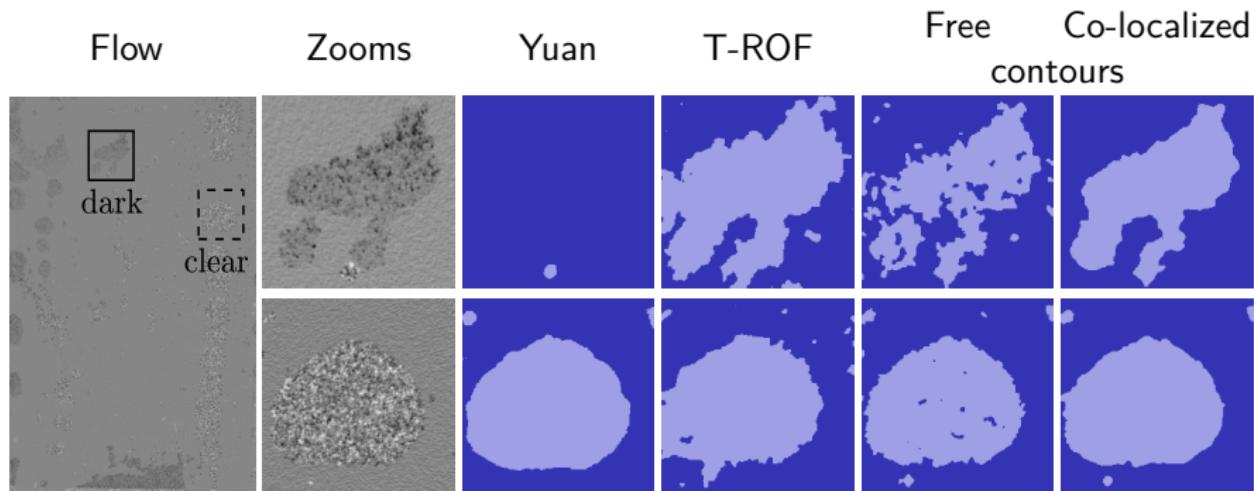
$90.2 \pm 1.9\%$

contours

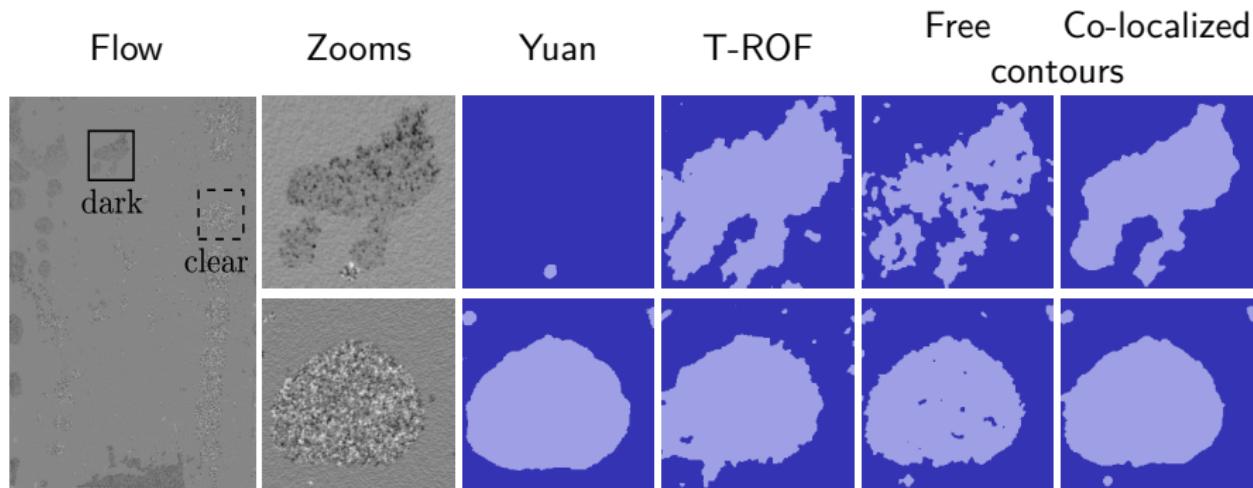


$91.1 \pm 1.5\%$

Low activity: $Q_G = 300\text{mL/min}$ - $Q_L = 300\text{mL/min}$



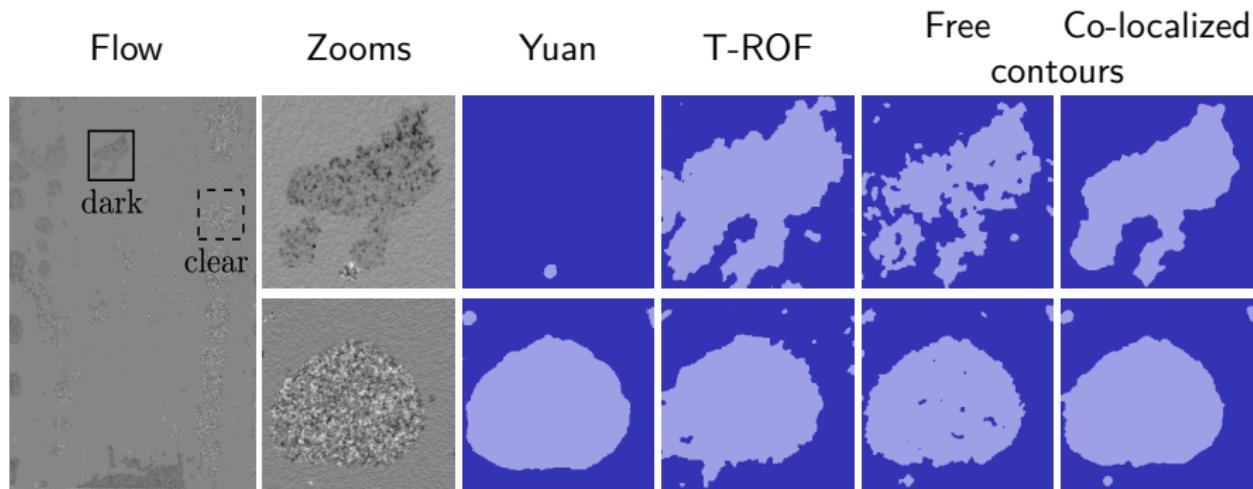
Low activity: $Q_G = 300\text{mL/min}$ - $Q_L = 300\text{mL/min}$



Liquid: $h_L = 0.4$

Gas: $h_G = 0.9$

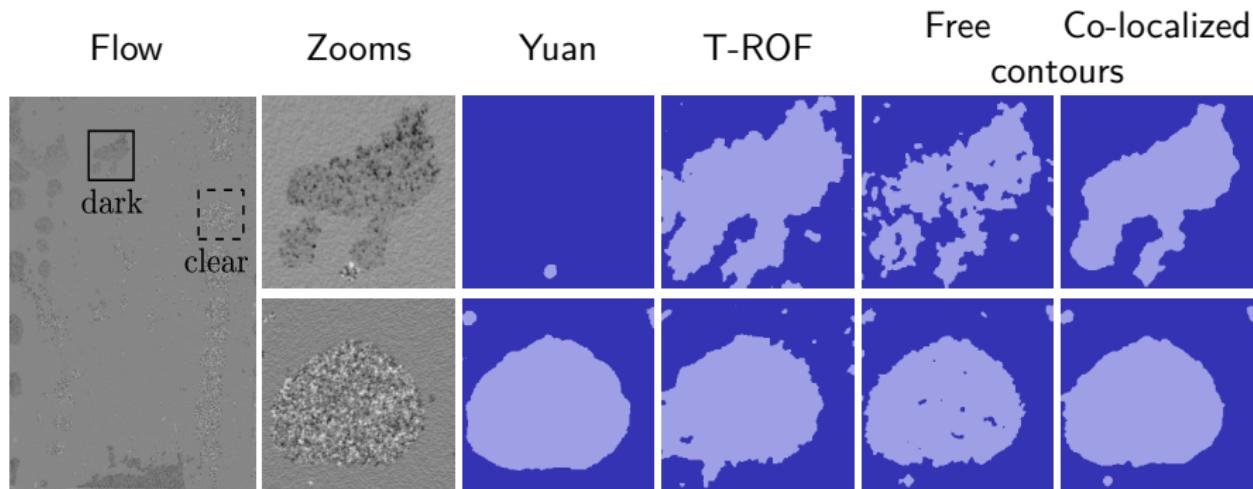
Low activity: $Q_G = 300\text{mL/min}$ - $Q_L = 300\text{mL/min}$



Liquid: $h_L = 0.4$ $\sigma_{\text{dark}}^2 = 10^{-2}$

Gas: $h_G = 0.9$

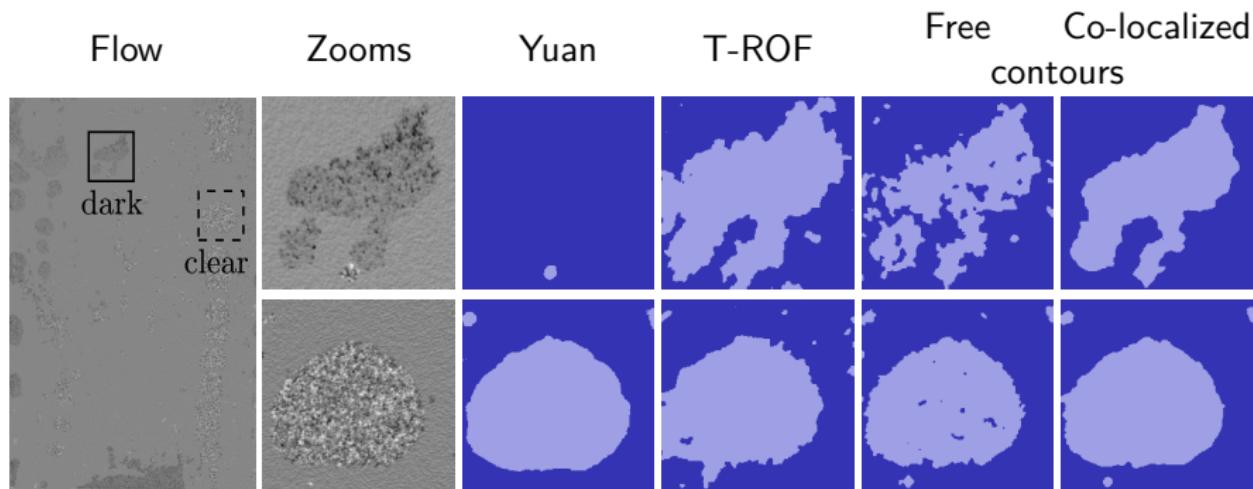
Low activity: $Q_G = 300\text{mL/min}$ - $Q_L = 300\text{mL/min}$



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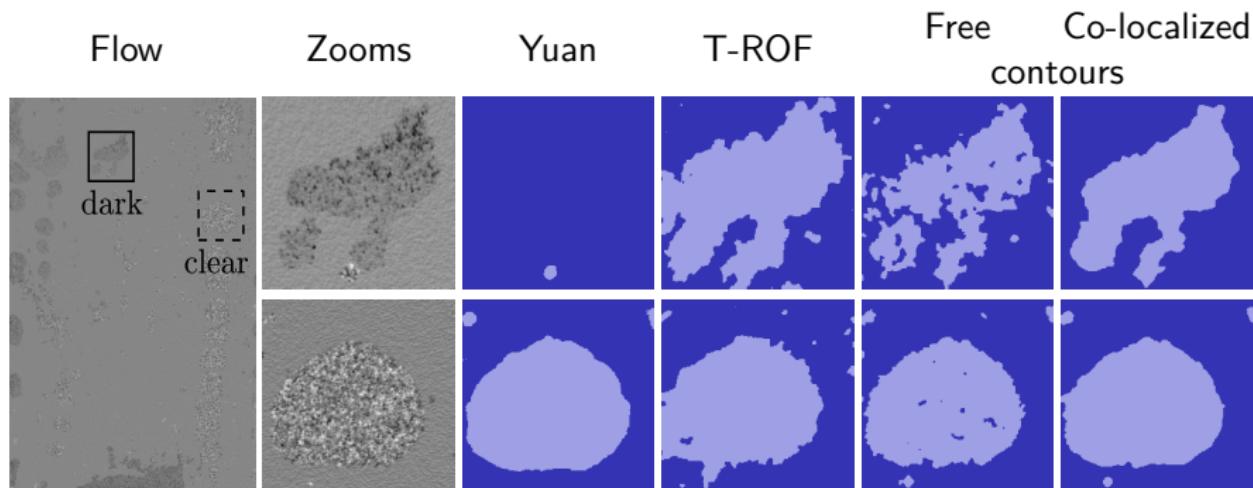
Low activity: $Q_G = 300\text{mL/min}$ - $Q_L = 300\text{mL/min}$



Liquid: $h_L = 0.4$ $\sigma_{\text{dark}}^2 = 10^{-2}$

Gas: $h_G = 0.9$ $\sigma_{\text{dark}}^2 = 10^{-2}$ (dark bubbles)

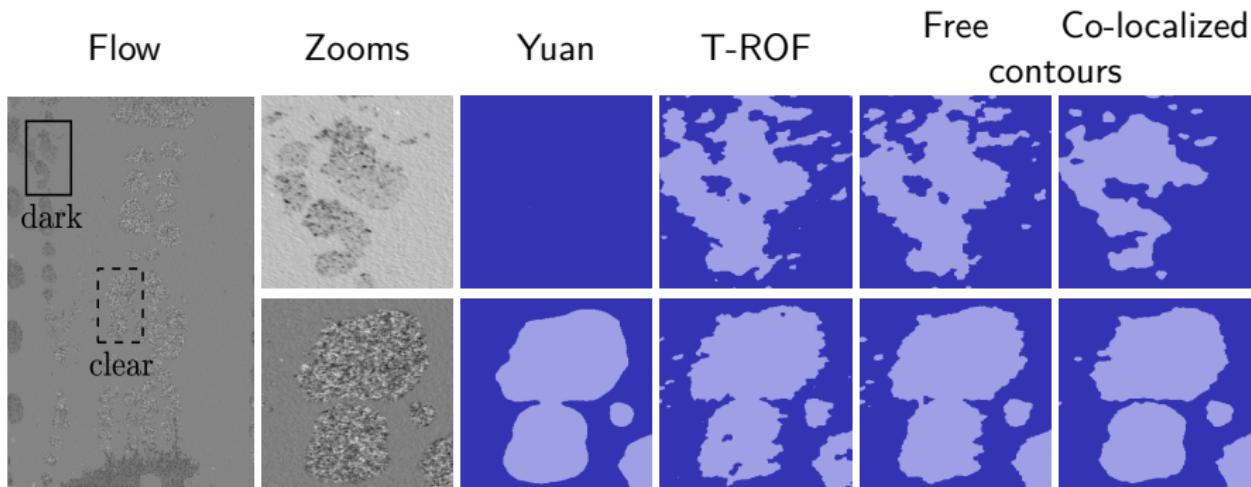
Low activity: $Q_G = 300\text{mL/min}$ - $Q_L = 300\text{mL/min}$



Liquid: $h_L = 0.4$ $\sigma_{\text{dark}}^2 = 10^{-2}$

Gas: $h_G = 0.9$ $\left| \begin{array}{l} \sigma_{\text{dark}}^2 = 10^{-2} \text{ (dark bubbles)} \\ \sigma_{\text{bright}}^2 = 10^{-1} \text{ (bright bubbles)} \end{array} \right.$

Transition: $Q_G = 400\text{mL/min}$ - $Q_L = 700\text{mL/min}$

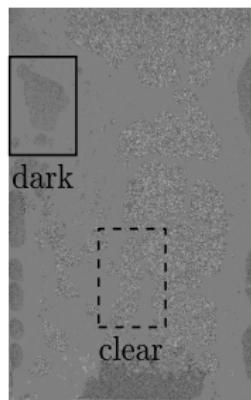


Liquid: $h_L = 0.4$ $\sigma_{\text{dark}}^2 = 10^{-2}$

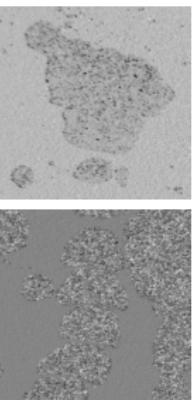
Gas: $h_G = 0.9$ $\left| \begin{array}{l} \sigma_{\text{dark}}^2 = 10^{-2} \text{ (dark bubbles)} \\ \sigma_{\text{bright}}^2 = 10^{-1} \text{ (bright bubbles).} \end{array} \right.$

High activity: $Q_G = 1200\text{mL/min}$ - $Q_L = 300\text{mL/min}$

Flow



Zooms



Yuan



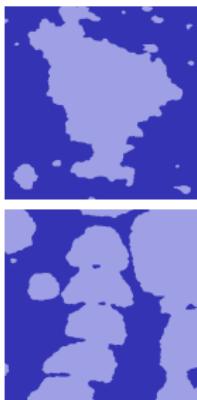
T-ROF



Free



Co-localized tours

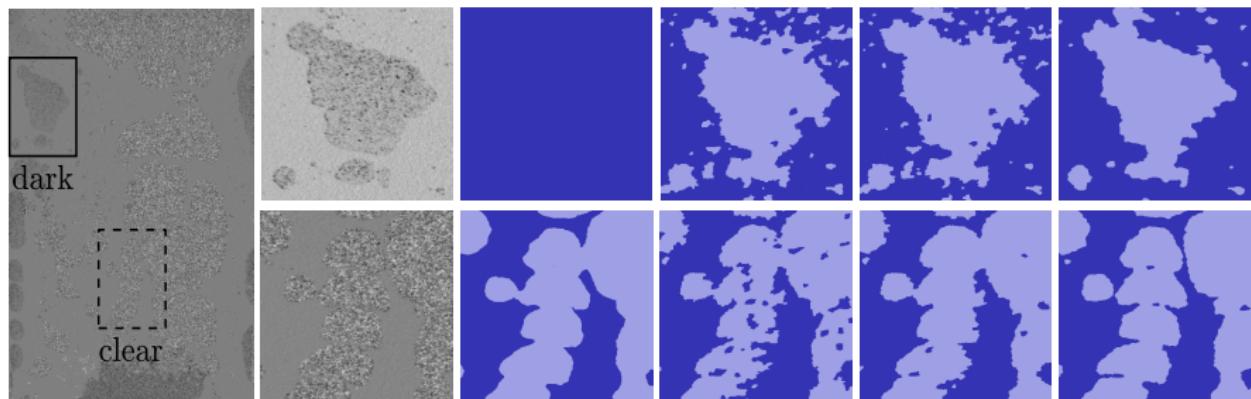


Liquid: $h_L = 0.4$ $\sigma_{\text{dark}}^2 = 10^{-2}$

$$\text{Gas: } h_G = 0.9 \quad \left| \begin{array}{l} \sigma_{\text{dark}}^2 = 10^{-2} \text{ (dark bubbles)} \\ \sigma_{\text{bright}}^2 = 10^{-1} \text{ (bright bubbles).} \end{array} \right.$$

High activity: $Q_G = 1200\text{mL/min}$ - $Q_L = 300\text{mL/min}$

Flow Zooms Yuan T-ROF Free
contours Co-localized



Computational load

1s

12s

700s

2100s

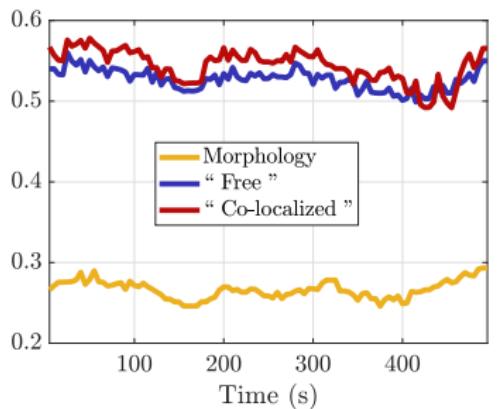
Liquid: $h_L = 0.4$ $\sigma_{\text{dark}}^2 = 10^{-2}$

$$\text{Gas: } h_G = 0.9 \quad \left| \begin{array}{l} \sigma_{\text{dark}}^2 = 10^{-2} \text{ (dark bubbles)} \\ \sigma_{\text{bright}}^2 = 10^{-1} \text{ (bright bubbles).} \end{array} \right.$$

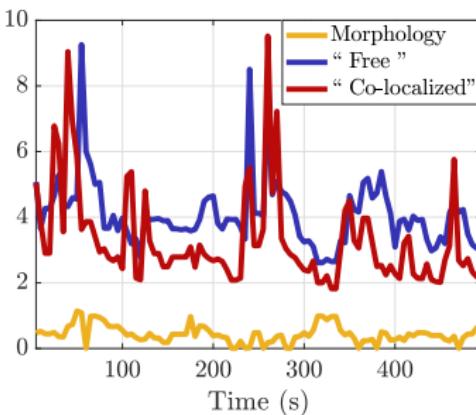
Multiphase flow through porous media

Laboratoire de Physique, ENS Lyon, V. Vidal, T. Busser, (M. Serres, IFPEN)

Gas fraction in the cell



Interface perimeter



Regularization parameters selection

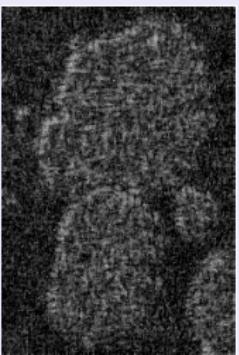
$$\left(\widehat{\boldsymbol{h}}, \widehat{\boldsymbol{v}}\right)(\mathcal{L}; \lambda, \alpha) = \operatorname{argmin}_{\boldsymbol{h}, \boldsymbol{v}} \sum_a \|\log \mathcal{L}_{a,.} - \log(a)\boldsymbol{h} - \boldsymbol{v}\|^2 + \lambda \mathcal{Q}(\mathbf{D}\boldsymbol{h}, \mathbf{D}\boldsymbol{v}; \alpha)$$

Regularization parameters selection

$$\left(\widehat{\boldsymbol{h}}, \widehat{\boldsymbol{v}}\right)(\mathcal{L}; \textcolor{brown}{\lambda}, \textcolor{brown}{\alpha}) = \operatorname{argmin}_{\boldsymbol{h}, \boldsymbol{v}} \sum_a \|\log \mathcal{L}_{a,.} - \log(a)\boldsymbol{h} - \boldsymbol{v}\|^2 + \textcolor{brown}{\lambda} \mathcal{Q}(\mathbf{D}\boldsymbol{h}, \mathbf{D}\boldsymbol{v}; \textcolor{brown}{\alpha})$$

Lin. reg. \hat{h}^{LR}

$$(\lambda, \alpha) = (0, 0)$$



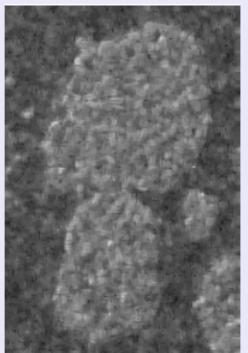
Regularization parameters selection

$$\left(\widehat{\boldsymbol{h}}, \widehat{\boldsymbol{v}}\right)(\mathcal{L}; \textcolor{brown}{\lambda}, \textcolor{brown}{\alpha}) = \operatorname{argmin}_{\boldsymbol{h}, \boldsymbol{v}} \sum_a \|\log \mathcal{L}_{a,.} - \log(a) \boldsymbol{h} - \boldsymbol{v}\|^2 + \textcolor{brown}{\lambda} \mathcal{Q}(\mathbf{D}\boldsymbol{h}, \mathbf{D}\boldsymbol{v}; \textcolor{brown}{\alpha})$$

Lin. reg. \hat{h}^{LR}

Co-localized contours estimate \hat{h}^C

$$(\lambda, \alpha) = (0, 0) \quad (\lambda, \alpha) = (0.5, 0.5)$$



too small

Regularization parameters selection

$$\left(\widehat{\boldsymbol{h}}, \widehat{\boldsymbol{v}}\right)(\mathcal{L}; \textcolor{brown}{\lambda}, \textcolor{brown}{\alpha}) = \operatorname{argmin}_{\boldsymbol{h}, \boldsymbol{v}} \sum_a \|\log \mathcal{L}_{a,.} - \log(a)\boldsymbol{h} - \boldsymbol{v}\|^2 + \textcolor{brown}{\lambda} \mathcal{Q}(\mathbf{D}\boldsymbol{h}, \mathbf{D}\boldsymbol{v}; \textcolor{brown}{\alpha})$$

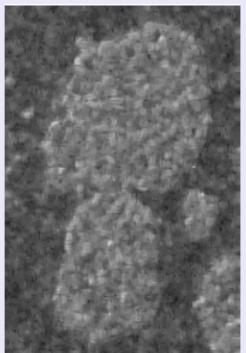
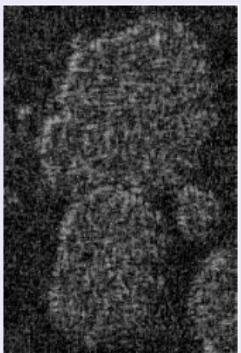
Lin. reg. \hat{h}^{LR}

Co-localized contours estimate \hat{h}^C

$$(\lambda, \alpha) = (0, 0)$$

$$(\lambda, \alpha) = (0.5, 0.5)$$

$$(\lambda, \alpha) = (500, 500)$$



too small



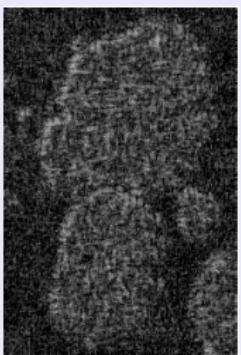
too large

Regularization parameters selection

$$\left(\widehat{\boldsymbol{h}}, \widehat{\boldsymbol{v}}\right)(\mathcal{L}; \textcolor{brown}{\lambda}, \textcolor{brown}{\alpha}) = \operatorname{argmin}_{\boldsymbol{h}, \boldsymbol{v}} \sum_a \|\log \mathcal{L}_{a,.} - \log(a)\boldsymbol{h} - \boldsymbol{v}\|^2 + \textcolor{brown}{\lambda} \mathcal{Q}(\mathbf{D}\boldsymbol{h}, \mathbf{D}\boldsymbol{v}; \textcolor{brown}{\alpha})$$

Lin. reg. \hat{h}^{LR}

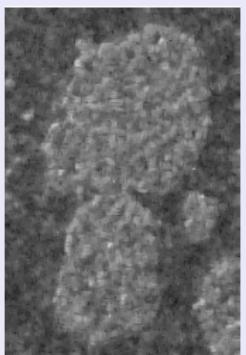
$$(\lambda, \alpha) = (0, 0)$$



too small

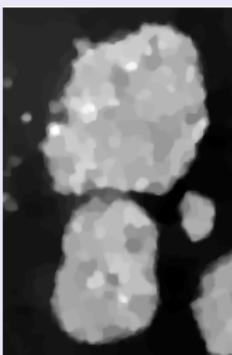
Co-localized contours estimate $\hat{\mathbf{h}}^C$

$$(\lambda, \alpha) = (0.5,$$



optimal

$$(\lambda^\dagger, \alpha^\dagger) = (11.5, 0.8) \quad (4)$$



too large

$$(\lambda, \alpha) = (500, 500)$$

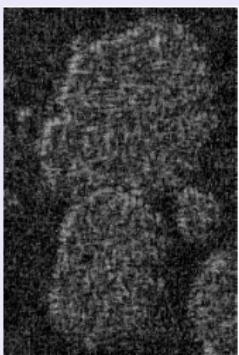


Regularization parameters selection

$$\left(\widehat{\boldsymbol{h}}, \widehat{\boldsymbol{v}}\right)(\mathcal{L}; \textcolor{brown}{\lambda}, \textcolor{brown}{\alpha}) = \operatorname{argmin}_{\boldsymbol{h}, \boldsymbol{v}} \sum_a \|\log \mathcal{L}_{a,.} - \log(a) \boldsymbol{h} - \boldsymbol{v}\|^2 + \textcolor{brown}{\lambda} \mathcal{Q}(\mathbf{D}\boldsymbol{h}, \mathbf{D}\boldsymbol{v}; \textcolor{brown}{\alpha})$$

Lin. reg. \hat{h}^{LR}

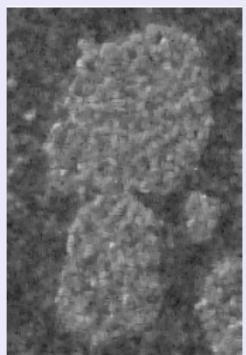
$$(\lambda, \alpha) = (0, 0)$$



too small

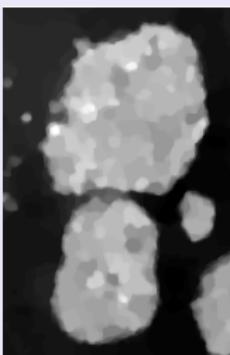
Co-localized contours estimate $\hat{\mathbf{h}}^C$

$$(\lambda, \alpha) = (0.5, 0.5) \quad ($$



optimal

$$(\lambda^\dagger, \alpha^\dagger) = (11.5, 0.8)$$



too large

What *optimal* means?

Regularization parameters selection

$$\left(\widehat{\boldsymbol{h}}, \widehat{\boldsymbol{v}}\right)(\mathcal{L}; \textcolor{brown}{\lambda}, \textcolor{brown}{\alpha}) = \operatorname{argmin}_{\boldsymbol{h}, \boldsymbol{v}} \sum_a \|\log \mathcal{L}_{a,.} - \log(a)\boldsymbol{h} - \boldsymbol{v}\|^2 + \textcolor{brown}{\lambda} \mathcal{Q}(\mathbf{D}\boldsymbol{h}, \mathbf{D}\boldsymbol{v}; \textcolor{brown}{\alpha})$$

Lin. reg. \hat{h}^{LR}

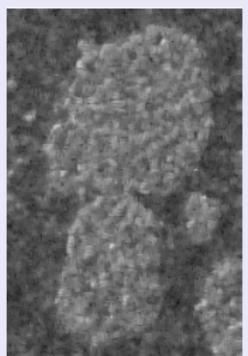
$$(\lambda, \alpha) = (0, 0)$$



too small

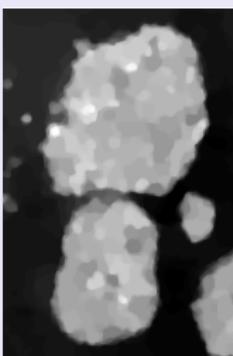
Co-localized contours estimate $\hat{\mathbf{h}}^C$

$$(\lambda, \alpha) = (0.5, 0.5)$$



optimal

$$(\lambda^\dagger, \alpha^\dagger) = (11.5, 0.8)$$



too large

What *optimal* means? How to determine λ^\dagger and α^\dagger ?

Parameter tuning (Grid search)

$$\hat{(\mathbf{h}, \mathbf{v})}(\mathcal{L}; \lambda, \alpha) = \underset{\mathbf{h}, \mathbf{v}}{\operatorname{argmin}} \sum_a \|\log \mathcal{L}_{a,.} - \log(a)\mathbf{h} - \mathbf{v}\|^2 + \lambda Q(\mathbf{Dh}, \mathbf{Dv}; \alpha)$$

\mathbf{h} : discriminant, **\mathbf{v} :** auxiliary

Parameter tuning (Grid search)

$$\hat{(\mathbf{h}, \mathbf{v})}(\mathcal{L}; \lambda, \alpha) = \underset{\mathbf{h}, \mathbf{v}}{\operatorname{argmin}} \sum_a \|\log \mathcal{L}_{a,.} - \log(a)\mathbf{h} - \mathbf{v}\|^2 + \lambda Q(\mathbf{Dh}, \mathbf{Dv}; \alpha)$$

\mathbf{h} : discriminant, **\mathbf{v} :** auxiliary

\bar{h} : true regularity

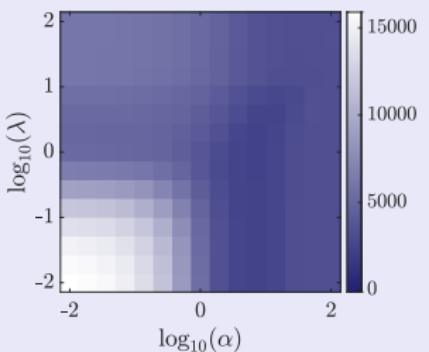
$$\mathcal{R}(\lambda, \alpha) = \left\| \widehat{\boldsymbol{h}}(\mathcal{L}; \lambda, \alpha) - \bar{\boldsymbol{h}} \right\|^2$$

Parameter tuning (Grid search)

$$\begin{aligned} \left(\hat{\mathbf{h}}, \hat{\mathbf{v}}\right)(\mathcal{L}; \lambda, \alpha) &= \underset{\mathbf{h}, \mathbf{v}}{\operatorname{argmin}} \sum_a \|\log \mathcal{L}_{a,.} - \log(a)\mathbf{h} - \mathbf{v}\|^2 + \lambda Q(\mathbf{Dh}, \mathbf{Dv}; \alpha) \\ \mathbf{h}: \text{discriminant, } \mathbf{v}: \text{auxiliary} \end{aligned}$$

\bar{h} : true regularity

$$\mathcal{R}(\lambda, \alpha) = \left\| \widehat{\boldsymbol{h}}(\mathcal{L}; \lambda, \alpha) - \bar{\boldsymbol{h}} \right\|^2$$

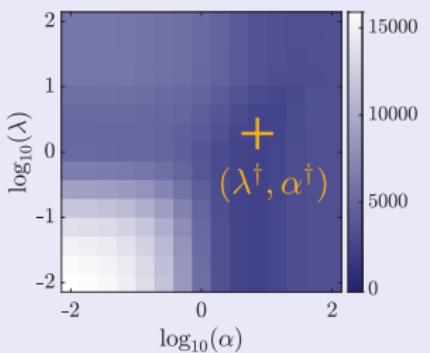


Parameter tuning (Grid search)

$$\begin{aligned} \left(\hat{\mathbf{h}}, \hat{\mathbf{v}}\right)(\mathcal{L}; \lambda, \alpha) &= \underset{\mathbf{h}, \mathbf{v}}{\operatorname{argmin}} \sum_a \|\log \mathcal{L}_{a,.} - \log(a)\mathbf{h} - \mathbf{v}\|^2 + \lambda Q(\mathbf{Dh}, \mathbf{Dv}; \alpha) \\ \mathbf{h}: \text{discriminant, } \mathbf{v}: \text{auxiliary} \end{aligned}$$

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$$\mathcal{R}(\lambda, \alpha) = \left\| \widehat{\boldsymbol{h}}(\mathcal{L}; \lambda, \alpha) - \bar{\boldsymbol{h}} \right\|^2$$

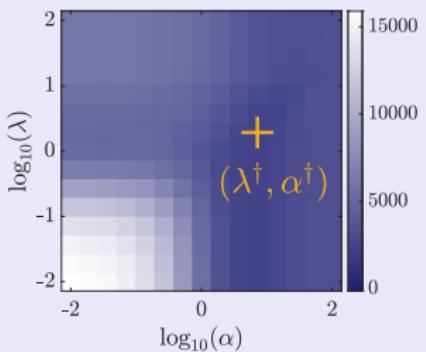


Parameter tuning (Grid search)

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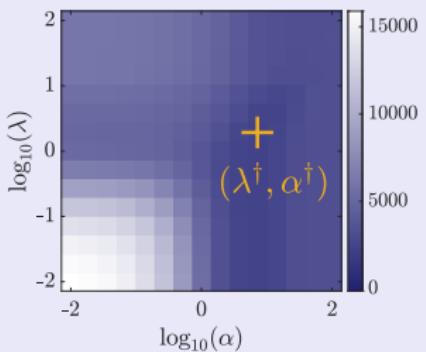


\bar{h} : unknown!

?

$$\begin{aligned} \left(\hat{\mathbf{h}}, \hat{\mathbf{v}}\right)(\mathcal{L}; \lambda, \alpha) &= \underset{\mathbf{h}, \mathbf{v}}{\operatorname{argmin}} \sum_a \|\log \mathcal{L}_{a,.} - \log(a)\mathbf{h} - \mathbf{v}\|^2 + \lambda Q(\mathbf{Dh}, \mathbf{Dv}; \alpha) \\ \mathbf{h}: \text{discriminant, } \mathbf{v}: \text{auxiliary} \end{aligned}$$

$$\mathcal{R}(\lambda, \alpha) = \left\| \widehat{\boldsymbol{h}}(\mathcal{L}; \lambda, \alpha) - \bar{\boldsymbol{h}} \right\|^2$$



\bar{h} : unknown!

?

Stein Unbiased Risk Estimate (SURE)

Stein Unbiased Risk Estimate (Principe)

Observations $y = \bar{x} + \zeta \in \mathbb{R}^P$, \bar{x} : truth and $\zeta \sim \mathcal{N}(\mathbf{0}, \rho^2 \mathbf{I})$

Stein Unbiased Risk Estimate (Principe)

Observations $y = \bar{x} + \zeta \in \mathbb{R}^P$, \bar{x} : truth and $\zeta \sim \mathcal{N}(\mathbf{0}, \rho^2 \mathbf{I})$

Parametric estimator $(y; \lambda) \mapsto \hat{x}(y; \lambda)$

$$\text{Ex. } \hat{\mathbf{x}}(\mathbf{y}; \lambda) = \begin{cases} (\mathbf{I} + \lambda \mathbf{D}^\top \mathbf{D})^{-1} \mathbf{y} & \text{(linear)} \\ \underset{\mathbf{x}}{\operatorname{argmin}} \| \mathbf{y} - \mathbf{x} \|^2 + \lambda \mathcal{Q}(\mathbf{D}\mathbf{x}) & \text{(nonlinear)} \end{cases}$$

Stein Unbiased Risk Estimate (Principe)

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Quadratic error $R(\lambda) \triangleq \mathbb{E}_\zeta \|\hat{\mathbf{x}}(\mathbf{y}; \lambda) - \bar{\mathbf{x}}\|^2 \stackrel{?}{=} \mathbb{E}_\zeta \widehat{R}(\mathbf{y}; \lambda)$ $\bar{\mathbf{x}}$ unknown

Stein Unbiased Risk Estimate (Principle)

Observations $y = \bar{x} + \zeta \in \mathbb{R}^P$, \bar{x} : truth and $\zeta \sim \mathcal{N}(\mathbf{0}, \rho^2 \mathbf{I})$

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$$\text{Ex. } \hat{\mathbf{x}}(\mathbf{y}; \lambda) = \begin{cases} \left(\mathbf{I} + \lambda \mathbf{D}^\top \mathbf{D} \right)^{-1} \mathbf{y} & (\text{linear}) \\ \underset{\mathbf{x}}{\operatorname{argmin}} \| \mathbf{y} - \mathbf{x} \|^2 + \lambda \mathcal{Q}(\mathbf{D}\mathbf{x}) & (\text{nonlinear}) \end{cases}$$

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Theorem (*Stein, 1981*)

Let $(\mathbf{y}; \lambda) \mapsto \hat{\mathbf{x}}(\mathbf{y}; \lambda)$ an estimator of $\bar{\mathbf{x}}$

- weakly differentiable w.r.t. \mathbf{y} ,
 - such that $\zeta \mapsto \langle \widehat{\mathbf{x}}(\bar{\mathbf{x}} + \zeta; \lambda), \zeta \rangle$ is integrable w.r.t. $\mathcal{N}(\mathbf{0}, \rho^2 \mathbf{I})$.

$$\begin{aligned}\widehat{R}(\mathbf{y}; \lambda) &\triangleq \|\widehat{\mathbf{x}}(\mathbf{y}; \lambda) - \mathbf{y}\|^2 + 2\rho^2 \text{tr}(\partial_{\mathbf{y}} \widehat{\mathbf{x}}(\mathbf{y}; \lambda)) - \rho^2 P \\ &\implies R(\lambda) = \mathbb{E}_{\zeta}[\widehat{R}(\mathbf{y}; \lambda)].\end{aligned}$$

Generalized Stein Unbiased Risk Estimate

Observations $y = \Phi\bar{x} + \zeta \in \mathbb{R}^P$, $\bar{x} \in \mathbb{R}^N$, $\Phi : \mathbb{R}^{P \times N}$ and $\zeta \sim \mathcal{N}(\mathbf{0}, \mathcal{S})$

Generalized Stein Unbiased Risk Estimate

Observations $\mathbf{y} = \Phi \bar{\mathbf{x}} + \zeta \in \mathbb{R}^P$, $\bar{\mathbf{x}} \in \mathbb{R}^N$, $\Phi : \mathbb{R}^{P \times N}$ and $\zeta \sim \mathcal{N}(\mathbf{0}, \mathcal{S})$

E.g. the estimators $\hat{\mathbf{h}}(\mathcal{L}; \lambda, \alpha)$ with free or co-localized contours

$$\log \mathcal{L} = \Phi(\bar{\mathbf{h}}, \bar{\mathbf{v}}) + \zeta$$

$$\Phi : (\mathbf{h}, \mathbf{v}) \mapsto \{\log(a)\mathbf{h} + \mathbf{v}\}_a$$

Generalized Stein Unbiased Risk Estimate

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$$\log \mathcal{L} = \Phi(\bar{\boldsymbol{h}}, \bar{\boldsymbol{v}}) + \zeta \quad \zeta \sim \mathcal{N}(\mathbf{0}, \mathcal{S}) \quad \mathcal{R} = \|\hat{\boldsymbol{h}} - \bar{\boldsymbol{h}}\|^2$$

$$\Phi : (\mathbf{h}, \mathbf{v}) \mapsto \{\log(a)\mathbf{h} + \mathbf{v}\}_a \quad \text{and} \quad \Pi : (\mathbf{h}, \mathbf{v}) \mapsto (\mathbf{h}, \mathbf{0})$$

Generalized Stein Unbiased Risk Estimate

Observations $y = \Phi\bar{x} + \zeta \in \mathbb{R}^P$, $\bar{x} \in \mathbb{R}^N$, $\Phi : \mathbb{R}^{P \times N}$ and $\zeta \sim \mathcal{N}(\mathbf{0}, \mathcal{S})$

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$$\text{Projected estimation error} \quad R_{\Pi}(\Lambda) \triangleq \mathbb{E}_{\zeta} \|\Pi \hat{x}(y; \Lambda) - \Pi \bar{x}\|^2$$

Introduction
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Texture characterization
○○○○

Design of functionals
○○

Accelerated minimization algorithms
○○○○○○○○○○○○○○○○

Hyperparameters tuning
○○○○●○○○○○○○○○○○

Conclusion
○

Generalized Stein Unbiased Risk Estimate (Computation)

Observations $\mathbf{y} = \Phi \bar{\mathbf{x}} + \zeta \in \mathbb{R}^P$, $\bar{\mathbf{x}} \in \mathbb{R}^N$, $\Phi : \mathbb{R}^{P \times N}$ and $\zeta \sim \mathcal{N}(\mathbf{0}, \Sigma)$

$$R_{\Pi}(\Lambda) \triangleq \mathbb{E}_{\zeta} \|\Pi \hat{\mathbf{x}}(\mathbf{y}; \Lambda) - \Pi \bar{\mathbf{x}}\|^2$$

Introduction
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Texture characterization
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Design of functionals
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Accelerated minimization algorithms
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Hyperparameters tuning
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Conclusion
○

Generalized Stein Unbiased Risk Estimate (Computation)

Observations $\mathbf{y} = \Phi \bar{\mathbf{x}} + \zeta \in \mathbb{R}^P$, $\bar{\mathbf{x}} \in \mathbb{R}^N$, $\Phi : \mathbb{R}^{P \times N}$ and $\zeta \sim \mathcal{N}(\mathbf{0}, \mathcal{S})$

$$\begin{aligned} R_{\Pi}(\Lambda) &\triangleq \mathbb{E}_{\zeta} \|\Pi \hat{\mathbf{x}}(\mathbf{y}; \Lambda) - \Pi \bar{\mathbf{x}}\|^2 \\ &= \mathbb{E}_{\zeta} \left\| \Pi (\Phi^\top \Phi)^{-1} \Phi^\top \Phi (\hat{\mathbf{x}}(\mathbf{y}; \Lambda) - \bar{\mathbf{x}}) \right\|^2 \quad \mathbf{A} \triangleq \Pi (\Phi^\top \Phi)^{-1} \Phi^\top \end{aligned}$$

Generalized Stein Unbiased Risk Estimate (Computation)

Observations $\mathbf{y} = \Phi \bar{\mathbf{x}} + \zeta \in \mathbb{R}^P$, $\bar{\mathbf{x}} \in \mathbb{R}^N$, $\Phi : \mathbb{R}^{P \times N}$ and $\zeta \sim \mathcal{N}(\mathbf{0}, \mathcal{S})$

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Generalized Stein Unbiased Risk Estimate (Computation)

Observations $\mathbf{y} = \Phi \bar{\mathbf{x}} + \zeta \in \mathbb{R}^P$, $\bar{\mathbf{x}} \in \mathbb{R}^N$, $\Phi : \mathbb{R}^{P \times N}$ and $\zeta \sim \mathcal{N}(\mathbf{0}, \Sigma)$

$$\begin{aligned} R_{\Pi}(\Lambda) &\triangleq \mathbb{E}_{\zeta} \|\Pi \hat{\mathbf{x}}(\mathbf{y}; \Lambda) - \Pi \bar{\mathbf{x}}\|^2 \\ &= \mathbb{E}_{\zeta} \left\| \Pi (\Phi^\top \Phi)^{-1} \Phi^\top \Phi (\hat{\mathbf{x}}(\mathbf{y}; \Lambda) - \bar{\mathbf{x}}) \right\|^2 \quad \mathbf{A} \triangleq \Pi (\Phi^\top \Phi)^{-1} \Phi^\top \\ &= \mathbb{E}_{\zeta} \|\mathbf{A}(\Phi \hat{\mathbf{x}}(\mathbf{y}; \Lambda) - \mathbf{y} + \mathbf{y} - \Phi \bar{\mathbf{x}})\|^2 \\ &= \mathbb{E}_{\zeta} \|\mathbf{A}(\Phi \hat{\mathbf{x}}(\mathbf{y}; \Lambda) - \mathbf{y} + \zeta)\|^2 \end{aligned}$$

Generalized Stein Unbiased Risk Estimate (Computation)

Observations $\mathbf{y} = \Phi \bar{\mathbf{x}} + \zeta \in \mathbb{R}^P$, $\bar{\mathbf{x}} \in \mathbb{R}^N$, $\Phi : \mathbb{R}^{P \times N}$ and $\zeta \sim \mathcal{N}(\mathbf{0}, \mathcal{S})$

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Generalized Stein Unbiased Risk Estimate (Computation)

Observations $\mathbf{y} = \Phi \bar{\mathbf{x}} + \zeta \in \mathbb{R}^P$, $\bar{\mathbf{x}} \in \mathbb{R}^N$, $\Phi : \mathbb{R}^{P \times N}$ and $\zeta \sim \mathcal{N}(\mathbf{0}, \mathcal{S})$

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Generalized Stein Unbiased Risk Estimate (Computation)

Observations $y = \Phi\bar{x} + \zeta \in \mathbb{R}^P$, $\bar{x} \in \mathbb{R}^N$, $\Phi : \mathbb{R}^{P \times N}$ and $\zeta \sim \mathcal{N}(\mathbf{0}, \mathcal{S})$

$$\begin{aligned}
R_\Pi(\Lambda) &\triangleq \mathbb{E}_\zeta \|\Pi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda) - \Pi \bar{\mathbf{x}}\|^2 \\
&= \mathbb{E}_\zeta \left\| \Pi(\Phi^\top \Phi)^{-1} \Phi^\top \Phi (\widehat{\mathbf{x}}(\mathbf{y}; \Lambda) - \bar{\mathbf{x}}) \right\|^2 \quad \mathbf{A} \triangleq \Pi(\Phi^\top \Phi)^{-1} \Phi^\top \\
&= \mathbb{E}_\zeta \|\mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda) - \mathbf{y}) + \mathbf{y} - \Phi \bar{\mathbf{x}}\|^2 \\
&= \mathbb{E}_\zeta \|\mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda) - \mathbf{y} + \zeta)\|^2 \\
&= \mathbb{E}_\zeta \left[\|\mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda) - \mathbf{y})\|^2 + 2 \langle \mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda) - \mathbf{y}), \mathbf{A}\zeta \rangle + \|\mathbf{A}\zeta\|^2 \right] \\
&= \mathbb{E}_\zeta \left[\|\mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda) - \mathbf{y})\|^2 + 2 \langle \mathbf{A}\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda), \mathbf{A}\zeta \rangle - 2 \langle \mathbf{A}\mathbf{y}, \mathbf{A}\zeta \rangle + \|\mathbf{A}\zeta\|^2 \right] \\
&= \mathbb{E}_\zeta \left[\|\mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda) - \mathbf{y})\|^2 + 2 \langle \mathbf{A}\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda), \mathbf{A}\zeta \rangle - 2 \langle \mathbf{A}(\Phi \bar{\mathbf{x}} + \zeta), \mathbf{A}\zeta \rangle + \|\mathbf{A}\zeta\|^2 \right]
\end{aligned}$$

Generalized Stein Unbiased Risk Estimate (Computation)

Observations $y = \Phi\bar{x} + \zeta \in \mathbb{R}^P$, $\bar{x} \in \mathbb{R}^N$, $\Phi : \mathbb{R}^{P \times N}$ and $\zeta \sim \mathcal{N}(\mathbf{0}, \mathcal{S})$

$$\begin{aligned}
R_\Pi(\Lambda) &\triangleq \mathbb{E}_\zeta \|\Pi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda) - \Pi \bar{\mathbf{x}}\|^2 \\
&= \mathbb{E}_\zeta \left\| \Pi(\Phi^\top \Phi)^{-1} \Phi^\top \Phi (\widehat{\mathbf{x}}(\mathbf{y}; \Lambda) - \bar{\mathbf{x}}) \right\|^2 \quad \mathbf{A} \triangleq \Pi(\Phi^\top \Phi)^{-1} \Phi^\top \\
&= \mathbb{E}_\zeta \|\mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda) - \mathbf{y}) + \mathbf{y} - \Phi \bar{\mathbf{x}}\|^2 \\
&= \mathbb{E}_\zeta \|\mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda) - \mathbf{y} + \zeta)\|^2 \\
&= \mathbb{E}_\zeta \left[\|\mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda) - \mathbf{y})\|^2 + 2 \langle \mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda) - \mathbf{y}), \mathbf{A}\zeta \rangle + \|\mathbf{A}\zeta\|^2 \right] \\
&= \mathbb{E}_\zeta \left[\|\mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda) - \mathbf{y})\|^2 + 2 \langle \mathbf{A}\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda), \mathbf{A}\zeta \rangle - 2 \langle \mathbf{A}\mathbf{y}, \mathbf{A}\zeta \rangle + \|\mathbf{A}\zeta\|^2 \right] \\
&= \mathbb{E}_\zeta \left[\|\mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda) - \mathbf{y})\|^2 + 2 \langle \mathbf{A}\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda), \mathbf{A}\zeta \rangle - 2 \langle \mathbf{A} \zeta, \mathbf{A}\zeta \rangle + \|\mathbf{A}\zeta\|^2 \right]
\end{aligned}$$

Generalized Stein Unbiased Risk Estimate (Computation)

Observations $\mathbf{y} = \Phi \bar{\mathbf{x}} + \zeta \in \mathbb{R}^P$, $\bar{\mathbf{x}} \in \mathbb{R}^N$, $\Phi : \mathbb{R}^{P \times N}$ and $\zeta \sim \mathcal{N}(\mathbf{0}, \Sigma)$

$$\begin{aligned}
 R_{\Pi}(\Lambda) &\triangleq \mathbb{E}_{\zeta} \|\Pi \hat{\mathbf{x}}(\mathbf{y}; \Lambda) - \Pi \bar{\mathbf{x}}\|^2 \\
 &= \mathbb{E}_{\zeta} \left\| \Pi (\Phi^\top \Phi)^{-1} \Phi^\top \Phi (\hat{\mathbf{x}}(\mathbf{y}; \Lambda) - \bar{\mathbf{x}}) \right\|^2 \quad \mathbf{A} \triangleq \Pi (\Phi^\top \Phi)^{-1} \Phi^\top \\
 &= \mathbb{E}_{\zeta} \|\mathbf{A}(\Phi \hat{\mathbf{x}}(\mathbf{y}; \Lambda)) - \mathbf{y} + \mathbf{y} - \Phi \bar{\mathbf{x}}\|^2 \\
 &= \mathbb{E}_{\zeta} \|\mathbf{A}(\Phi \hat{\mathbf{x}}(\mathbf{y}; \Lambda)) - \mathbf{y} + \zeta\|^2 \\
 &= \mathbb{E}_{\zeta} [\|\mathbf{A}(\Phi \hat{\mathbf{x}}(\mathbf{y}; \Lambda)) - \mathbf{y}\|^2 + 2 \langle \mathbf{A}(\Phi \hat{\mathbf{x}}(\mathbf{y}; \Lambda)) - \mathbf{y}, \mathbf{A} \zeta \rangle + \|\mathbf{A} \zeta\|^2] \\
 &= \mathbb{E}_{\zeta} [\|\mathbf{A}(\Phi \hat{\mathbf{x}}(\mathbf{y}; \Lambda)) - \mathbf{y}\|^2 + 2 \langle \mathbf{A} \Phi \hat{\mathbf{x}}(\mathbf{y}; \Lambda), \mathbf{A} \zeta \rangle - 2 \langle \mathbf{A} \mathbf{y}, \mathbf{A} \zeta \rangle + \|\mathbf{A} \zeta\|^2] \\
 &= \mathbb{E}_{\zeta} [\|\mathbf{A}(\Phi \hat{\mathbf{x}}(\mathbf{y}; \Lambda)) - \mathbf{y}\|^2 + 2 \langle \mathbf{A} \Phi \hat{\mathbf{x}}(\mathbf{y}; \Lambda), \mathbf{A} \zeta \rangle - 2 \langle \mathbf{A} \mathbf{y}, \mathbf{A} \zeta \rangle + \|\mathbf{A} \zeta\|^2] \\
 &= \mathbb{E}_{\zeta} \|\mathbf{A}(\Phi \hat{\mathbf{x}}(\mathbf{y}; \Lambda)) - \mathbf{y}\|^2 + 2 \mathbb{E}_{\zeta} \langle \mathbf{A} \Phi \hat{\mathbf{x}}(\mathbf{y}; \Lambda), \mathbf{A} \zeta \rangle - \mathbb{E}_{\zeta} \|\mathbf{A} \zeta\|^2
 \end{aligned}$$

Generalized Stein Unbiased Risk Estimate (Computation)

Observations $\mathbf{y} = \Phi \bar{\mathbf{x}} + \zeta \in \mathbb{R}^P$, $\bar{\mathbf{x}} \in \mathbb{R}^N$, $\Phi : \mathbb{R}^{P \times N}$ and $\zeta \sim \mathcal{N}(\mathbf{0}, \Sigma)$

$$\begin{aligned}
 R_{\Pi}(\Lambda) &\triangleq \mathbb{E}_{\zeta} \|\Pi \hat{\mathbf{x}}(\mathbf{y}; \Lambda) - \Pi \bar{\mathbf{x}}\|^2 \\
 &= \mathbb{E}_{\zeta} \left\| \Pi (\Phi^\top \Phi)^{-1} \Phi^\top \Phi (\hat{\mathbf{x}}(\mathbf{y}; \Lambda) - \bar{\mathbf{x}}) \right\|^2 \quad \mathbf{A} \triangleq \Pi (\Phi^\top \Phi)^{-1} \Phi^\top \\
 &= \mathbb{E}_{\zeta} \|\mathbf{A}(\Phi \hat{\mathbf{x}}(\mathbf{y}; \Lambda)) - \mathbf{y} + \mathbf{y} - \Phi \bar{\mathbf{x}}\|^2 \\
 &= \mathbb{E}_{\zeta} \|\mathbf{A}(\Phi \hat{\mathbf{x}}(\mathbf{y}; \Lambda)) - \mathbf{y} + \zeta\|^2 \\
 &= \mathbb{E}_{\zeta} [\|\mathbf{A}(\Phi \hat{\mathbf{x}}(\mathbf{y}; \Lambda)) - \mathbf{y}\|^2 + 2 \langle \mathbf{A}(\Phi \hat{\mathbf{x}}(\mathbf{y}; \Lambda)) - \mathbf{y}, \mathbf{A} \zeta \rangle + \|\mathbf{A} \zeta\|^2] \\
 &= \mathbb{E}_{\zeta} [\|\mathbf{A}(\Phi \hat{\mathbf{x}}(\mathbf{y}; \Lambda)) - \mathbf{y}\|^2 + 2 \langle \mathbf{A} \Phi \hat{\mathbf{x}}(\mathbf{y}; \Lambda), \mathbf{A} \zeta \rangle - 2 \langle \mathbf{A} \mathbf{y}, \mathbf{A} \zeta \rangle + \|\mathbf{A} \zeta\|^2] \\
 &= \mathbb{E}_{\zeta} [\|\mathbf{A}(\Phi \hat{\mathbf{x}}(\mathbf{y}; \Lambda)) - \mathbf{y}\|^2 + 2 \langle \mathbf{A} \Phi \hat{\mathbf{x}}(\mathbf{y}; \Lambda), \mathbf{A} \zeta \rangle - 2 \langle \mathbf{A} \mathbf{y} - \zeta, \mathbf{A} \zeta \rangle + \|\mathbf{A} \zeta\|^2] \\
 &= \mathbb{E}_{\zeta} \|\mathbf{A}(\Phi \hat{\mathbf{x}}(\mathbf{y}; \Lambda)) - \mathbf{y}\|^2 + 2 \mathbb{E}_{\zeta} \langle \mathbf{A} \Phi \hat{\mathbf{x}}(\mathbf{y}; \Lambda), \mathbf{A} \zeta \rangle - \text{tr}(\mathbf{A} \Sigma \mathbf{A}^\top)
 \end{aligned}$$

Generalized Stein Unbiased Risk Estimate (Computation)

Observations $\mathbf{y} = \Phi \bar{\mathbf{x}} + \zeta \in \mathbb{R}^P$, $\bar{\mathbf{x}} \in \mathbb{R}^N$, $\Phi : \mathbb{R}^{P \times N}$ and $\zeta \sim \mathcal{N}(\mathbf{0}, \mathcal{S})$

$$\begin{aligned}
 R_{\Pi}(\Lambda) &\triangleq \mathbb{E}_{\zeta} \|\Pi \hat{\mathbf{x}}(\mathbf{y}; \Lambda) - \Pi \bar{\mathbf{x}}\|^2 \\
 &= \mathbb{E}_{\zeta} \left\| \Pi (\Phi^\top \Phi)^{-1} \Phi^\top \Phi (\hat{\mathbf{x}}(\mathbf{y}; \Lambda) - \bar{\mathbf{x}}) \right\|^2 \quad \mathbf{A} \triangleq \Pi (\Phi^\top \Phi)^{-1} \Phi^\top \\
 &= \mathbb{E}_{\zeta} \|\mathbf{A}(\Phi \hat{\mathbf{x}}(\mathbf{y}; \Lambda)) - \mathbf{y} + \mathbf{y} - \Phi \bar{\mathbf{x}}\|^2 \\
 &= \mathbb{E}_{\zeta} \|\mathbf{A}(\Phi \hat{\mathbf{x}}(\mathbf{y}; \Lambda)) - \mathbf{y} + \zeta\|^2 \\
 &= \mathbb{E}_{\zeta} \left[\|\mathbf{A}(\Phi \hat{\mathbf{x}}(\mathbf{y}; \Lambda)) - \mathbf{y}\|^2 + 2 \langle \mathbf{A}(\Phi \hat{\mathbf{x}}(\mathbf{y}; \Lambda)) - \mathbf{y}, \mathbf{A} \zeta \rangle + \|\mathbf{A} \zeta\|^2 \right] \\
 &= \mathbb{E}_{\zeta} \left[\|\mathbf{A}(\Phi \hat{\mathbf{x}}(\mathbf{y}; \Lambda)) - \mathbf{y}\|^2 + 2 \langle \mathbf{A} \Phi \hat{\mathbf{x}}(\mathbf{y}; \Lambda), \mathbf{A} \zeta \rangle - 2 \langle \mathbf{A} \mathbf{y}, \mathbf{A} \zeta \rangle + \|\mathbf{A} \zeta\|^2 \right] \\
 &= \mathbb{E}_{\zeta} \left[\|\mathbf{A}(\Phi \hat{\mathbf{x}}(\mathbf{y}; \Lambda)) - \mathbf{y}\|^2 + 2 \langle \mathbf{A} \Phi \hat{\mathbf{x}}(\mathbf{y}; \Lambda), \mathbf{A} \zeta \rangle - 2 \langle \mathbf{A} \mathbf{y} - \zeta, \mathbf{A} \zeta \rangle + \|\mathbf{A} \zeta\|^2 \right] \\
 &= \mathbb{E}_{\zeta} \|\mathbf{A}(\Phi \hat{\mathbf{x}}(\mathbf{y}; \Lambda)) - \mathbf{y}\|^2 + 2 \mathbb{E}_{\zeta} \langle \mathbf{A} \Phi \hat{\mathbf{x}}(\mathbf{y}; \Lambda), \mathbf{A} \zeta \rangle - \underbrace{\text{tr}(\mathbf{A} \mathcal{S} \mathbf{A}^\top)}_{\text{accessible}}
 \end{aligned}$$

Generalized Stein Unbiased Risk Estimate (Computation)

Observations $y = \Phi\bar{x} + \zeta \in \mathbb{R}^P$, $\bar{x} \in \mathbb{R}^N$, $\Phi : \mathbb{R}^{P \times N}$ and $\zeta \sim \mathcal{N}(\mathbf{0}, \mathcal{S})$

$$\begin{aligned}
R_{\Pi}(\Lambda) &\triangleq \mathbb{E}_{\zeta} \|\Pi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda) - \Pi \bar{\mathbf{x}}\|^2 \\
&= \mathbb{E}_{\zeta} \left\| \Pi (\Phi^\top \Phi)^{-1} \Phi^\top \Phi (\widehat{\mathbf{x}}(\mathbf{y}; \Lambda) - \bar{\mathbf{x}}) \right\|^2 \quad \mathbf{A} \triangleq \Pi (\Phi^\top \Phi)^{-1} \Phi^\top \\
&= \mathbb{E}_{\zeta} \|\mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda)) - \mathbf{y} + \mathbf{y} - \Phi \bar{\mathbf{x}}\|^2 \\
&= \mathbb{E}_{\zeta} \|\mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda)) - \mathbf{y} + \zeta\|^2 \\
&= \mathbb{E}_{\zeta} \left[\|\mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda)) - \mathbf{y}\|^2 + 2 \langle \mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda)) - \mathbf{y}, \mathbf{A}\zeta \rangle + \|\mathbf{A}\zeta\|^2 \right] \\
&= \mathbb{E}_{\zeta} \left[\|\mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda)) - \mathbf{y}\|^2 + 2 \langle \mathbf{A}\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda), \mathbf{A}\zeta \rangle - 2 \langle \mathbf{A}\mathbf{y}, \mathbf{A}\zeta \rangle + \|\mathbf{A}\zeta\|^2 \right] \\
&= \mathbb{E}_{\zeta} \left[\|\mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda)) - \mathbf{y}\|^2 + 2 \langle \mathbf{A}\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda), \mathbf{A}\zeta \rangle - 2 \langle \mathbf{A} \underbrace{\mathbf{y}}_{\text{accessible}}, \mathbf{A}\zeta \rangle + \|\mathbf{A}\zeta\|^2 \right] \\
&= \mathbb{E}_{\zeta} \underbrace{\|\mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda)) - \mathbf{y}\|^2}_{\text{accessible}} + 2 \mathbb{E}_{\zeta} \langle \mathbf{A}\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda), \mathbf{A}\zeta \rangle - \underbrace{\text{tr}(\mathbf{A}\mathbf{S}\mathbf{A}^\top)}_{\text{accessible}}
\end{aligned}$$

Generalized Stein Unbiased Risk Estimate (Computation)

Observations $y = \Phi\bar{x} + \zeta \in \mathbb{R}^P$, $\bar{x} \in \mathbb{R}^N$, $\Phi : \mathbb{R}^{P \times N}$ and $\zeta \sim \mathcal{N}(\mathbf{0}, \mathcal{S})$

$$\begin{aligned}
R_{\Pi}(\Lambda) &\triangleq \mathbb{E}_{\zeta} \|\Pi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda) - \Pi \bar{\mathbf{x}}\|^2 \\
&= \mathbb{E}_{\zeta} \left\| \Pi (\Phi^\top \Phi)^{-1} \Phi^\top \Phi (\widehat{\mathbf{x}}(\mathbf{y}; \Lambda) - \bar{\mathbf{x}}) \right\|^2 \quad \mathbf{A} \triangleq \Pi (\Phi^\top \Phi)^{-1} \Phi^\top \\
&= \mathbb{E}_{\zeta} \|\mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda) - \mathbf{y}) + \mathbf{y} - \Phi \bar{\mathbf{x}}\|^2 \\
&= \mathbb{E}_{\zeta} \|\mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda) - \mathbf{y} + \zeta)\|^2 \\
&= \mathbb{E}_{\zeta} \left[\|\mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda) - \mathbf{y})\|^2 + 2 \langle \mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda) - \mathbf{y}), \mathbf{A}\zeta \rangle + \|\mathbf{A}\zeta\|^2 \right] \\
&= \mathbb{E}_{\zeta} \left[\|\mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda) - \mathbf{y})\|^2 + 2 \langle \mathbf{A}\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda), \mathbf{A}\zeta \rangle - 2 \langle \mathbf{A}\mathbf{y}, \mathbf{A}\zeta \rangle + \|\mathbf{A}\zeta\|^2 \right] \\
&= \mathbb{E}_{\zeta} \left[\|\mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda) - \mathbf{y})\|^2 + 2 \langle \mathbf{A}\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda), \mathbf{A}\zeta \rangle - 2 \langle \mathbf{A} \underbrace{\mathbf{y}}_{\text{accessible}}, \mathbf{A}\zeta \rangle + \|\mathbf{A}\zeta\|^2 \right] \\
&= \mathbb{E}_{\zeta} \underbrace{\|\mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda) - \mathbf{y})\|^2}_{\text{accessible}} + 2 \mathbb{E}_{\zeta} \underbrace{\langle \mathbf{A}\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda), \mathbf{A}\zeta \rangle}_{\text{accessible}} - \text{tr}(\mathbf{A} \mathbf{S} \mathbf{A}^\top)
\end{aligned}$$

Generalized Stein Unbiased Risk Estimate (Computation)

Observations $y = \Phi\bar{x} + \zeta \in \mathbb{R}^P$, $\bar{x} \in \mathbb{R}^N$, $\Phi : \mathbb{R}^{P \times N}$ and $\zeta \sim \mathcal{N}(\mathbf{0}, \mathcal{S})$

$$\begin{aligned}
R_\Pi(\Lambda) &\triangleq \mathbb{E}_\zeta \|\Pi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda) - \Pi \bar{\mathbf{x}}\|^2 \\
&= \mathbb{E}_\zeta \left\| \Pi (\Phi^\top \Phi)^{-1} \Phi^\top \Phi (\widehat{\mathbf{x}}(\mathbf{y}; \Lambda) - \bar{\mathbf{x}}) \right\|^2 \quad \mathbf{A} \triangleq \Pi (\Phi^\top \Phi)^{-1} \Phi^\top \\
&= \mathbb{E}_\zeta \|\mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda)) - \mathbf{y} + \mathbf{y} - \Phi \bar{\mathbf{x}}\|^2 \\
&= \mathbb{E}_\zeta \|\mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda)) - \mathbf{y} + \zeta\|^2 \\
&= \mathbb{E}_\zeta \left[\|\mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda)) - \mathbf{y}\|^2 + 2 \langle \mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda)) - \mathbf{y}, \mathbf{A}\zeta \rangle + \|\mathbf{A}\zeta\|^2 \right] \\
&= \mathbb{E}_\zeta \left[\|\mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda)) - \mathbf{y}\|^2 + 2 \langle \mathbf{A}\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda), \mathbf{A}\zeta \rangle - 2 \langle \mathbf{A}\mathbf{y}, \mathbf{A}\zeta \rangle + \|\mathbf{A}\zeta\|^2 \right] \\
&= \mathbb{E}_\zeta \left[\|\mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda)) - \mathbf{y}\|^2 + 2 \langle \mathbf{A}\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda), \mathbf{A}\zeta \rangle - 2 \langle \mathbf{A} - \zeta, \mathbf{A}\zeta \rangle + \|\mathbf{A}\zeta\|^2 \right] \\
&= \mathbb{E}_\zeta \underline{\|\mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda)) - \mathbf{y}\|^2} + 2 \mathbb{E}_\zeta \underline{\langle \mathbf{A}\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda), \mathbf{A}\zeta \rangle} - \underline{\text{tr}(\mathbf{A}\mathcal{S}\mathbf{A}^\top)}
\end{aligned}$$

$$\mathbb{E}_C \langle \mathbf{A} \Phi \widehat{\mathbf{x}}(\mathbf{v}; \Lambda), \mathbf{A} \zeta \rangle \equiv \mathbb{E}_C \langle \Pi \widehat{\mathbf{x}}(\mathbf{v}; \Lambda), \mathbf{A} \zeta \rangle$$

Generalized Stein Unbiased Risk Estimate (Computation)

Observations $y = \Phi\bar{x} + \zeta \in \mathbb{R}^P$, $\bar{x} \in \mathbb{R}^N$, $\Phi : \mathbb{R}^{P \times N}$ and $\zeta \sim \mathcal{N}(\mathbf{0}, \mathcal{S})$

$$\begin{aligned}
R_{\Pi}(\Lambda) &\triangleq \mathbb{E}_{\zeta} \|\Pi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda) - \Pi \bar{\mathbf{x}}\|^2 \\
&= \mathbb{E}_{\zeta} \left\| \Pi (\Phi^\top \Phi)^{-1} \Phi^\top \Phi (\widehat{\mathbf{x}}(\mathbf{y}; \Lambda) - \bar{\mathbf{x}}) \right\|^2 \quad \mathbf{A} \triangleq \Pi (\Phi^\top \Phi)^{-1} \Phi^\top \\
&= \mathbb{E}_{\zeta} \|\mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda) - \mathbf{y}) + \mathbf{y} - \Phi \bar{\mathbf{x}}\|^2 \\
&= \mathbb{E}_{\zeta} \|\mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda) - \mathbf{y} + \zeta)\|^2 \\
&= \mathbb{E}_{\zeta} [\|\mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda) - \mathbf{y})\|^2 + 2 \langle \mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda) - \mathbf{y}), \mathbf{A}\zeta \rangle + \|\mathbf{A}\zeta\|^2] \\
&= \mathbb{E}_{\zeta} [\|\mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda) - \mathbf{y})\|^2 + 2 \langle \mathbf{A}\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda), \mathbf{A}\zeta \rangle - 2 \langle \mathbf{A}\mathbf{y}, \mathbf{A}\zeta \rangle + \|\mathbf{A}\zeta\|^2] \\
&= \mathbb{E}_{\zeta} [\|\mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda) - \mathbf{y})\|^2 + 2 \langle \mathbf{A}\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda), \mathbf{A}\zeta \rangle - 2 \langle \mathbf{A}\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda), \mathbf{A}\zeta \rangle + \|\mathbf{A}\zeta\|^2] \\
&= \mathbb{E}_{\zeta} \underbrace{\|\mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda) - \mathbf{y})\|^2}_{\text{accessible}} + \underbrace{2\mathbb{E}_{\zeta} \langle \mathbf{A}\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda), \mathbf{A}\zeta \rangle - \text{tr}(\mathbf{A}\mathcal{S}\mathbf{A}^\top)}_{\text{accessible}} \\
\mathbb{E}_{\zeta} \langle \mathbf{A}\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda), \mathbf{A}\zeta \rangle &= \mathbb{E}_{\zeta} \langle \Pi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda), \mathbf{A}\zeta \rangle \\
&= \int \langle \Pi \widehat{\mathbf{x}}(\Phi \bar{\mathbf{x}} + \zeta; \lambda), \mathbf{A}\zeta \rangle \exp\left(-\frac{\zeta^\top \mathcal{S}^{-1} \zeta}{2}\right) d\zeta
\end{aligned}$$

Generalized Stein Unbiased Risk Estimate (Computation)

Observations $\mathbf{y} = \Phi \bar{\mathbf{x}} + \zeta \in \mathbb{R}^P$, $\bar{\mathbf{x}} \in \mathbb{R}^N$, $\Phi : \mathbb{R}^{P \times N}$ and $\zeta \sim \mathcal{N}(\mathbf{0}, \mathcal{S})$

$$\begin{aligned}
R_{\Pi}(\Lambda) &\triangleq \mathbb{E}_{\zeta} \|\Pi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda) - \Pi \bar{\mathbf{x}}\|^2 \\
&= \mathbb{E}_{\zeta} \left\| \Pi (\Phi^\top \Phi)^{-1} \Phi^\top \Phi (\widehat{\mathbf{x}}(\mathbf{y}; \Lambda) - \bar{\mathbf{x}}) \right\|^2 \quad \mathbf{A} \triangleq \Pi (\Phi^\top \Phi)^{-1} \Phi^\top \\
&= \mathbb{E}_{\zeta} \|\mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda)) - \mathbf{y} + \mathbf{y} - \Phi \bar{\mathbf{x}}\|^2 \\
&= \mathbb{E}_{\zeta} \|\mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda)) - \mathbf{y} + \zeta\|^2 \\
&= \mathbb{E}_{\zeta} [\|\mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda)) - \mathbf{y}\|^2 + 2 \langle \mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda)) - \mathbf{y}, \mathbf{A}\zeta \rangle + \|\mathbf{A}\zeta\|^2] \\
&= \mathbb{E}_{\zeta} [\|\mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda)) - \mathbf{y}\|^2 + 2 \langle \mathbf{A}\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda), \mathbf{A}\zeta \rangle - 2 \langle \mathbf{A}\mathbf{y}, \mathbf{A}\zeta \rangle + \|\mathbf{A}\zeta\|^2] \\
&= \mathbb{E}_{\zeta} [\|\mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda)) - \mathbf{y}\|^2 + 2 \langle \mathbf{A}\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda), \mathbf{A}\zeta \rangle - 2 \langle \mathbf{A} \quad \zeta, \mathbf{A}\zeta \rangle + \|\mathbf{A}\zeta\|^2] \\
&= \mathbb{E}_{\zeta} \underbrace{\|\mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda)) - \mathbf{y}\|^2}_{\text{accessible}} + 2 \underbrace{\mathbb{E}_{\zeta} \langle \mathbf{A}\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda), \mathbf{A}\zeta \rangle}_{\text{accessible}} - \text{tr}(\mathbf{A} \mathcal{S} \mathbf{A}^\top) \\
\mathbb{E}_{\zeta} \langle \mathbf{A}\Phi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda), \mathbf{A}\zeta \rangle &= \mathbb{E}_{\zeta} \langle \Pi \widehat{\mathbf{x}}(\mathbf{y}; \Lambda), \mathbf{A}\zeta \rangle \\
&= \int \langle \Pi \widehat{\mathbf{x}}(\Phi \bar{\mathbf{x}} + \zeta; \lambda), \mathbf{A}\zeta \rangle \exp\left(-\frac{\zeta^\top \mathcal{S}^{-1} \zeta}{2}\right) d\zeta \\
(\text{Gen. I.b.P.}) &= \mathbb{E}_{\zeta} \text{tr}(\mathcal{S} \mathbf{A}^\top \Pi \partial_y \widehat{\mathbf{x}}(\mathbf{y}; \Lambda))
\end{aligned}$$

Stein Unbiased Risk Estimate généralisé

Observations $\mathbf{y} = \Phi \bar{\mathbf{x}} + \zeta \in \mathbb{R}^P$, $\bar{\mathbf{x}} \in \mathbb{R}^N$, $\Phi : \mathbb{R}^{P \times N}$ and $\zeta \sim \mathcal{N}(\mathbf{0}, \mathcal{S})$

E.g. the estimators $\hat{\mathbf{h}}(\mathcal{L}; \lambda, \alpha)$ with free or co-localized contours

$$\log \mathcal{L} = \Phi(\bar{\mathbf{h}}, \bar{\mathbf{v}}) + \zeta \quad \zeta \sim \mathcal{N}(\mathbf{0}, \mathcal{S}) \quad \mathcal{R} = \|\hat{\mathbf{h}} - \bar{\mathbf{h}}\|^2$$

$$\Phi : (\mathbf{h}, \mathbf{v}) \mapsto \{\log(a)\mathbf{h} + \mathbf{v}\}_a \quad \begin{array}{|c|c|c|c|c|c|c|c|} \hline & \cdot \\ \hline \cdot & \cdot \\ \hline \end{array} \quad \Pi : (\mathbf{h}, \mathbf{v}) \mapsto (\mathbf{h}, \mathbf{0})$$

Projected estimation error $R_\Pi(\Lambda) \triangleq \mathbb{E}_\zeta \|\Pi \hat{\mathbf{x}}(\mathbf{y}; \Lambda) - \Pi \bar{\mathbf{x}}\|^2$

Theorem (Pascal, 2020)

Let $(\mathbf{y}; \Lambda) \mapsto \hat{\mathbf{x}}(\mathbf{y}; \Lambda)$ an estimator of $\bar{\mathbf{x}}$

- weakly differentiable w.r.t. \mathbf{y} ,
- such that $\zeta \mapsto \langle \Pi \hat{\mathbf{x}}(\bar{\mathbf{x}} + \zeta; \lambda), \mathbf{A} \zeta \rangle$ is integrable w.r.t. $\mathcal{N}(\mathbf{0}, \mathcal{S})$.

$$\begin{aligned} \widehat{R}(\Lambda) &\triangleq \|\mathbf{A}(\Phi \hat{\mathbf{x}}(\mathbf{y}; \Lambda) - \mathbf{y})\|^2 + 2\text{tr} \left(\mathcal{S} \mathbf{A}^\top \Pi \partial_{\mathbf{y}} \hat{\mathbf{x}}(\mathbf{y}; \Lambda) \right) - \text{tr} \left(\mathbf{A} \mathcal{S} \mathbf{A}^\top \right) \\ &\implies R_\Pi(\Lambda) = \mathbb{E}_\zeta [\widehat{R}(\Lambda)]. \end{aligned}$$

Stein Unbiased Risk Estimate généralisé

Observations $\mathbf{y} = \Phi \bar{\mathbf{x}} + \zeta \in \mathbb{R}^P$, $\bar{\mathbf{x}} \in \mathbb{R}^N$, $\Phi : \mathbb{R}^{P \times N}$ and $\zeta \sim \mathcal{N}(\mathbf{0}, \mathcal{S})$

E.g. the estimators $\hat{\mathbf{h}}(\mathcal{L}; \lambda, \alpha)$ with free or co-localized contours

$$\log \mathcal{L} = \Phi(\bar{\mathbf{h}}, \bar{\mathbf{v}}) + \zeta \quad \zeta \sim \mathcal{N}(\mathbf{0}, \mathcal{S}) \quad \mathcal{R} = \|\hat{\mathbf{h}} - \bar{\mathbf{h}}\|^2$$

$$\Phi : (\mathbf{h}, \mathbf{v}) \mapsto \{\log(a)\mathbf{h} + \mathbf{v}\}_a \quad \begin{array}{|c|c|c|c|c|c|c|c|c|c|c|} \hline & \cdot \\ \hline \cdot & \cdot \\ \hline \end{array} \quad \Pi : (\mathbf{h}, \mathbf{v}) \mapsto (\mathbf{h}, \mathbf{0})$$

Projected estimation error $R_\Pi(\Lambda) \triangleq \mathbb{E}_\zeta \|\Pi \hat{\mathbf{x}}(\mathbf{y}; \Lambda) - \Pi \bar{\mathbf{x}}\|^2$

Theorem (Pascal, 2020)

Let $(\mathbf{y}; \Lambda) \mapsto \hat{\mathbf{x}}(\mathbf{y}; \Lambda)$ an estimator of $\bar{\mathbf{x}}$

- weakly differentiable w.r.t. \mathbf{y} ,
- such that $\zeta \mapsto \langle \Pi \hat{\mathbf{x}}(\bar{\mathbf{x}} + \zeta; \lambda), \mathbf{A} \zeta \rangle$ is integrable w.r.t. $\mathcal{N}(\mathbf{0}, \mathcal{S})$.

$$\begin{aligned} \widehat{R}(\Lambda) &\triangleq \|\mathbf{A}(\Phi \hat{\mathbf{x}}(\mathbf{y}; \Lambda) - \mathbf{y})\|^2 + 2\text{tr} \left(\mathcal{S} \mathbf{A}^\top \Pi \partial_{\mathbf{y}} \hat{\mathbf{x}}(\mathbf{y}; \Lambda) \right) - \text{tr} \left(\mathbf{A} \mathcal{S} \mathbf{A}^\top \right) \\ &\implies R_\Pi(\Lambda) = \mathbb{E}_\zeta [\widehat{R}(\Lambda)]. \end{aligned}$$

Computation of the degrees of freedom

Degrees of freedom

$$\text{dof} \triangleq \text{tr} \left(\boldsymbol{\mathcal{S}} \mathbf{A}^\top \Pi \partial_{\mathbf{y}} \hat{\mathbf{x}}(\mathbf{y}; \boldsymbol{\Lambda}) \right)$$

Computation of the degrees of freedom

Degrees of freedom

$$\text{dof} \triangleq \text{tr} \left(\mathcal{S} \mathbf{A}^\top \boldsymbol{\Pi} \partial_{\mathbf{y}} \hat{\mathbf{x}}(\mathbf{y}; \boldsymbol{\Lambda}) \right)$$

- Monte Carlo strategy (MC) Large size matrix $\mathbf{M} \in \mathbb{R}^{P \times P}$
$$\text{tr}(\mathbf{M}) = \mathbb{E}_{\boldsymbol{\varepsilon}} \langle \mathbf{M} \boldsymbol{\varepsilon}, \boldsymbol{\varepsilon} \rangle, \quad \boldsymbol{\varepsilon} \sim \mathcal{N}(\mathbf{0}, \mathbf{I}_P)$$

Computation of the degrees of freedom

Degrees of freedom $\text{dof} \triangleq \text{tr} \left(\mathcal{S} \mathbf{A}^\top \boldsymbol{\Pi} \partial_{\mathbf{y}} \hat{\mathbf{x}}(\mathbf{y}; \boldsymbol{\Lambda}) \right)$

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- Finite Differences (FD) Inaccessible Jacobian matrix
$$\partial_{\mathbf{y}} \hat{\mathbf{x}} [\boldsymbol{\varepsilon}] \underset{\nu \rightarrow 0}{\simeq} \frac{1}{\nu} (\hat{\mathbf{x}}(\mathbf{y} + \nu \boldsymbol{\varepsilon}; \boldsymbol{\Lambda}) - \hat{\mathbf{x}}(\mathbf{y}; \boldsymbol{\Lambda}))$$

Computation of the degrees of freedom

Degrees of freedom $\text{dof} \triangleq \text{tr} \left(\mathcal{S} \mathbf{A}^\top \boldsymbol{\Pi} \partial_{\mathbf{y}} \hat{\mathbf{x}}(\mathbf{y}; \boldsymbol{\Lambda}) \right)$

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Proposition (Pascal, 2020)

Let $(\mathbf{y}; \boldsymbol{\Lambda}) \mapsto \hat{\mathbf{x}}(\mathbf{y}; \boldsymbol{\Lambda})$ an estimator of $\bar{\mathbf{x}}$

- uniformly Lipschitz continuous w.r.t. \mathbf{y} ,
- such that $\forall \boldsymbol{\Lambda} \in \mathbb{R}^L, \hat{\mathbf{x}}(\mathbf{0}_P; \boldsymbol{\Lambda}) = \mathbf{0}_N$. Then

$$\mathbb{E}_\zeta [\text{dof}] = \lim_{\nu \rightarrow 0} \mathbb{E}_{\zeta, \varepsilon} \left[\frac{1}{\nu} \left\langle \mathcal{S} \mathbf{A}^\top \boldsymbol{\Pi} (\hat{\mathbf{x}}(\mathbf{y} + \nu \varepsilon; \boldsymbol{\Lambda}) - \hat{\mathbf{x}}(\mathbf{y}; \boldsymbol{\Lambda})), \varepsilon \right\rangle \right]$$

Stein Unbiased Risk Estimate (Computation)

Observations $\mathbf{y} = \Phi \bar{\mathbf{x}} + \zeta \in \mathbb{R}^P$, $\bar{\mathbf{x}} \in \mathbb{R}^N$, $\Phi : \mathbb{R}^{P \times N}$ and $\zeta \sim \mathcal{N}(\mathbf{0}, \sigma^2 \mathbf{I})$

Projected estimation error $R_{\Pi}(\Lambda) \triangleq \mathbb{E}_{\zeta} \|\Pi \hat{\mathbf{x}}(\mathbf{y}; \Lambda) - \Pi \bar{\mathbf{x}}\|^2$

Generalized Finite Difference Monte Carlo SURE

$$\begin{aligned} \widehat{R}_{\nu, \epsilon}(\mathbf{y}; \Lambda | \mathcal{S}) &\triangleq \|\mathbf{A}(\Phi \hat{\mathbf{x}}(\mathbf{y}; \Lambda) - \mathbf{y})\|^2 + \\ &\frac{2}{\nu} \left\langle \mathcal{S} \mathbf{A}^\top \Pi(\hat{\mathbf{x}}(\mathbf{y} + \nu \epsilon; \Lambda) - \hat{\mathbf{x}}(\mathbf{y}; \Lambda)), \epsilon \right\rangle - \text{tr}(\mathbf{A} \mathcal{S} \mathbf{A}^\top) \end{aligned}$$

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Let $(\mathbf{y}; \Lambda) \mapsto \hat{\mathbf{x}}(\mathbf{y}; \Lambda)$ an estimator of $\bar{\mathbf{x}}$

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- such that $\forall \Lambda \in \mathbb{R}^L$, $\hat{\mathbf{x}}(\mathbf{0}_P; \Lambda) = \mathbf{0}_N$. Then

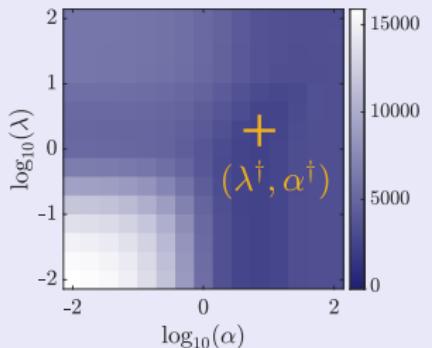
$$R_{\Pi}(\Lambda) = \lim_{\nu \rightarrow 0} \mathbb{E}_{\zeta, \epsilon} [\widehat{R}_{\nu, \epsilon}(\mathbf{y}; \Lambda | \mathcal{S})]$$

Parameter tuning (Grid search)

$$\left(\hat{\mathbf{h}}, \hat{\mathbf{v}}\right)(\mathcal{L}; \lambda, \alpha) = \operatorname{argmin}_{\mathbf{h}, \mathbf{v}} \sum_a \|\log \mathcal{L}_{a,.} - \log(a)\mathbf{h} - \mathbf{v}\|^2 + \lambda \mathcal{Q}(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)$$

$\bar{\mathbf{h}}$: true regularity

$$\mathcal{R}(\lambda, \alpha) = \left\| \hat{\mathbf{h}}(\mathcal{L}; \lambda, \alpha) - \bar{\mathbf{h}} \right\|^2$$



$\bar{\mathbf{h}}$: unknown!

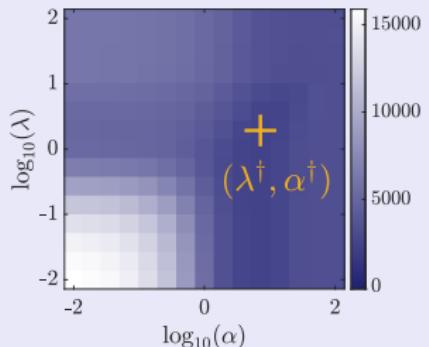
$$\hat{R}_{\nu, \epsilon}(\mathcal{L}; \lambda, \alpha | \mathcal{S})$$

Parameter tuning (Grid search)

$$\left(\hat{\mathbf{h}}, \hat{\mathbf{v}} \right) (\mathcal{L}; \lambda, \alpha) = \underset{\mathbf{h}, \mathbf{v}}{\operatorname{argmin}} \sum_a \|\log \mathcal{L}_{a,.} - \log(a) \mathbf{h} - \mathbf{v}\|^2 + \lambda \mathcal{Q}(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)$$

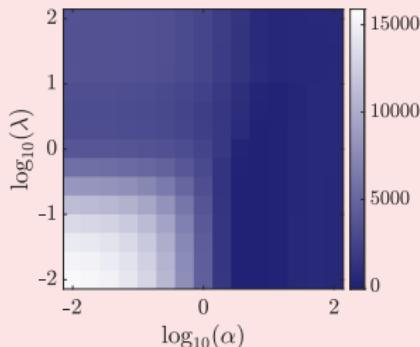
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$\bar{\mathbf{h}}$: unknown!

$$\hat{R}_{\nu, \epsilon}(\mathcal{L}; \lambda, \alpha | \mathcal{S})$$

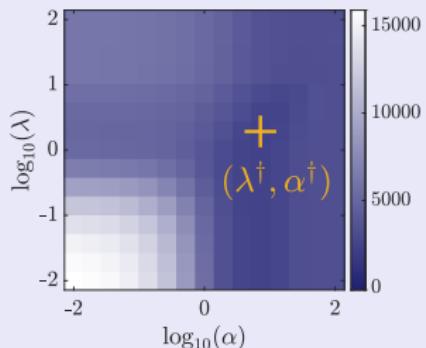


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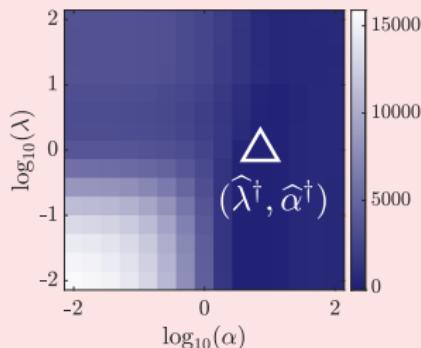
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$$\mathcal{R}(\lambda, \alpha) = \left\| \hat{\mathbf{h}}(\mathcal{L}; \lambda, \alpha) - \bar{\mathbf{h}} \right\|^2$$



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$$\hat{\mathcal{R}}_{\nu, \epsilon}(\mathcal{L}; \lambda, \alpha | \mathcal{S})$$

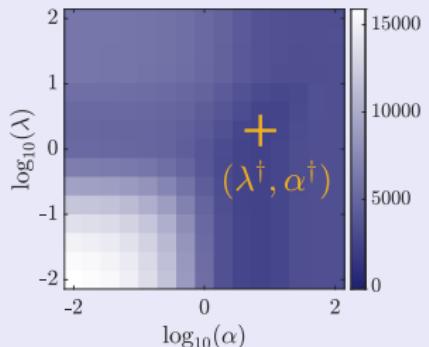


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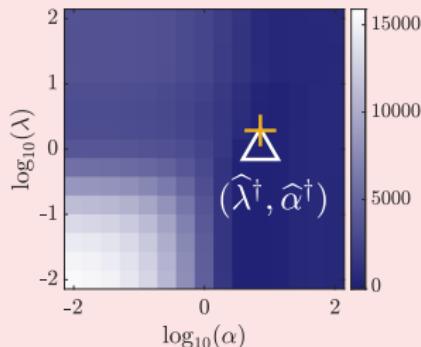
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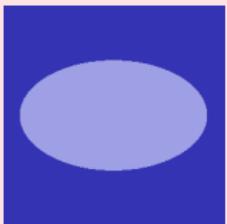
$$\hat{\mathcal{R}}_{\nu, \epsilon}(\mathcal{L}; \lambda, \alpha | \mathcal{S})$$



Grid search selection of regularization parameters

$$\left(\hat{\mathbf{h}}^F, \hat{\mathbf{v}}^F \right) (\mathcal{L}; \boldsymbol{\Lambda}) = \operatorname{argmin}_{\mathbf{h}, \mathbf{v}} \sum_a \|\log \mathcal{L}_{a,.} - \log(a) \mathbf{h} - \mathbf{v}\|^2 + \lambda Q_F(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)$$

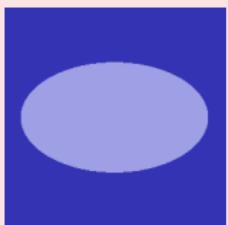
Example



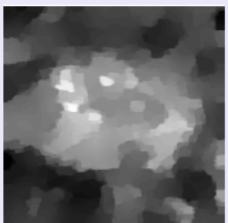
Grid search selection of regularization parameters

$$\left(\hat{\mathbf{h}}^F, \hat{\mathbf{v}}^F \right) (\mathcal{L}; \boldsymbol{\Lambda}) = \operatorname{argmin}_{\mathbf{h}, \mathbf{v}} \sum_a \| \log \mathcal{L}_{a,.} - \log(a) \mathbf{h} - \mathbf{v} \|^2 + \lambda Q_F(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)$$

Example



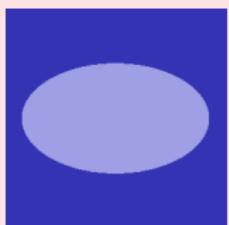
$\hat{\mathbf{h}}^F(\mathcal{L}; \lambda^\dagger, \alpha^\dagger)$
(grid)



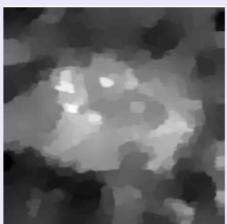
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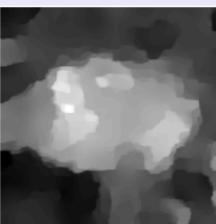
Example



$\hat{\mathbf{h}}^F(\mathcal{L}; \lambda^\dagger, \alpha^\dagger)$
(grid)



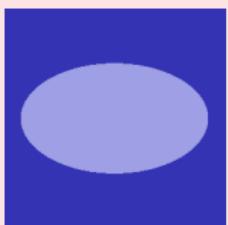
$\hat{\mathbf{h}}^F(\mathcal{L}; \hat{\lambda}^\dagger, \hat{\alpha}^\dagger)$
(grid)



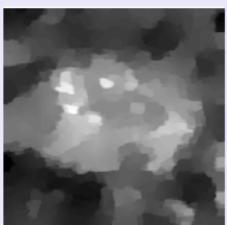
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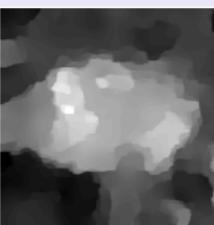
Example



$\hat{\mathbf{h}}^F(\mathcal{L}; \lambda^\dagger, \alpha^\dagger)$
(grid)



$\hat{\mathbf{h}}^F(\mathcal{L}; \hat{\lambda}^\dagger, \hat{\alpha}^\dagger)$
(grid)



$15 \times 15 = 225$ parameters \rightarrow grid search is very costly!

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Automatic minimization of Generalized SURE

Observations $\mathbf{y} = \Phi \bar{\mathbf{x}} + \zeta \in \mathbb{R}^P$, $\bar{\mathbf{x}} \in \mathbb{R}^N$, $\Phi : \mathbb{R}^{P \times N}$ and $\zeta \sim \mathcal{N}(\mathbf{0}, \mathcal{S})$

Generalized FDMC SURE $\lim_{\nu \rightarrow 0} \mathbb{E}_{\zeta, \epsilon} \hat{R}_{\nu, \epsilon}(\mathbf{y}; \Lambda | \mathcal{S}) = R_{\Pi}(\Lambda)$

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But : minimize $\hat{R}_{\nu, \epsilon}(\mathbf{y}; \Lambda | \mathcal{S})$ for given \mathbf{y}, \mathcal{S}

Automatic minimization of Generalized SURE

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Broyden-Fletcher-Goldfarb-Shanno Quasi-Newton (*Nocedal, 2006*)

for $t = 0, 1, \dots$

$$\mathbf{d}^{[t]} = -\mathbf{H}^{[t]} \partial_{\Lambda} \widehat{R}(\Lambda^{[t]}) \quad \text{descent direction}$$

$$\alpha^{[t]} \in \underset{\alpha \in \mathbb{R}}{\operatorname{Argmin}} \widehat{R}(\Lambda^{[t]} + \alpha \mathbf{d}^{[t]}) \quad \text{line search}$$

$$\Lambda^{[t+1]} = \Lambda^{[t]} + \alpha^{[t]} \mathbf{d}^{[t]}$$

$$\mathbf{u}^{[t]} = \partial_{\Lambda} \widehat{R}(\Lambda^{[t+1]}) - \partial_{\Lambda} \widehat{R}(\Lambda^{[t]}) \quad \text{gradient increment}$$

$$\mathbf{H}^{[t+1]} = \text{BFGS}(\mathbf{H}^{[t]}, \mathbf{d}^{[t]}, \mathbf{u}^{[t]}) \quad \text{"inverse Hessian" update}$$

Automatic minimization of Generalized SURE

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Stein Unbiased GrAdient Risk estimate

Generalized FDMC SURE

$$\begin{aligned}\widehat{R}_{\nu,\varepsilon}(\mathbf{y}; \boldsymbol{\Lambda} | \mathcal{S}) &= \|\mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \boldsymbol{\Lambda}) - \mathbf{y})\|^2 + \\ \frac{2}{\nu} \left\langle \mathcal{S} \mathbf{A}^\top \Pi(\widehat{\mathbf{x}}(\mathbf{y} + \nu\varepsilon; \boldsymbol{\Lambda}) - \widehat{\mathbf{x}}(\mathbf{y}; \boldsymbol{\Lambda})), \varepsilon \right\rangle - \text{tr}(\mathbf{A} \mathcal{S} \mathbf{A}^\top)\end{aligned}$$

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Generalized FDMC SURE

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Generalized Finite Difference Monte Carlo SUGAR

$$\begin{aligned}\partial_{\boldsymbol{\Lambda}} \widehat{R}_{\nu,\varepsilon}(\mathbf{y}; \boldsymbol{\Lambda} | \mathcal{S}) = & 2(\mathbf{A} \Phi \partial_{\boldsymbol{\Lambda}} \widehat{\mathbf{x}}(\mathbf{y}; \boldsymbol{\Lambda}))^\top \mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \boldsymbol{\Lambda}) - \mathbf{y}) \\ & + \frac{2}{\nu} \left\langle \mathcal{S} \mathbf{A}^\top \boldsymbol{\Pi} (\partial_{\boldsymbol{\Lambda}} \widehat{\mathbf{x}}(\mathbf{y} + \nu\varepsilon; \boldsymbol{\Lambda}) - \partial_{\boldsymbol{\Lambda}} \widehat{\mathbf{x}}(\mathbf{y}; \boldsymbol{\Lambda})) , \varepsilon \right\rangle\end{aligned}$$

Stein Unbiased GrAdient Risk estimate

Generalized FDMC SURE

$$\widehat{R}_{\nu,\varepsilon}(\mathbf{y}; \boldsymbol{\Lambda} | \mathcal{S}) = \|\mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \boldsymbol{\Lambda}) - \mathbf{y})\|^2 + \frac{2}{\nu} \left\langle \mathcal{S} \mathbf{A}^\top \boldsymbol{\Pi} (\widehat{\mathbf{x}}(\mathbf{y} + \nu \varepsilon; \boldsymbol{\Lambda}) - \widehat{\mathbf{x}}(\mathbf{y}; \boldsymbol{\Lambda})), \varepsilon \right\rangle - \text{tr}(\mathbf{A} \mathcal{S} \mathbf{A}^\top)$$

Generalized Finite Difference Monte Carlo SUGAR

$$\begin{aligned} \partial_{\boldsymbol{\Lambda}} \widehat{R}_{\nu,\varepsilon}(\mathbf{y}; \boldsymbol{\Lambda} | \mathcal{S}) &= 2(\mathbf{A} \Phi \partial_{\boldsymbol{\Lambda}} \widehat{\mathbf{x}}(\mathbf{y}; \boldsymbol{\Lambda}))^\top \mathbf{A}(\Phi \widehat{\mathbf{x}}(\mathbf{y}; \boldsymbol{\Lambda}) - \mathbf{y}) \\ &\quad + \frac{2}{\nu} \left\langle \mathcal{S} \mathbf{A}^\top \boldsymbol{\Pi} (\partial_{\boldsymbol{\Lambda}} \widehat{\mathbf{x}}(\mathbf{y} + \nu \varepsilon; \boldsymbol{\Lambda}) - \partial_{\boldsymbol{\Lambda}} \widehat{\mathbf{x}}(\mathbf{y}; \boldsymbol{\Lambda})), \varepsilon \right\rangle \end{aligned}$$

Theorem (Pascal, 2020)

Let $(\mathbf{y}; \boldsymbol{\Lambda}) \mapsto \widehat{\mathbf{x}}(\mathbf{y}; \boldsymbol{\Lambda})$ an estimator of $\bar{\mathbf{x}}$

- uniformly Lipschitz continuous w.r.t. \mathbf{y}
- such that $\forall \boldsymbol{\Lambda} \in \mathbb{R}^L$, $\widehat{\mathbf{x}}(\mathbf{0}_P; \boldsymbol{\Lambda}) = \mathbf{0}_N$,
- uniformly L -Lipschitz continuous w.r.t. $\boldsymbol{\Lambda}$, L ind. of \mathbf{y} . Then

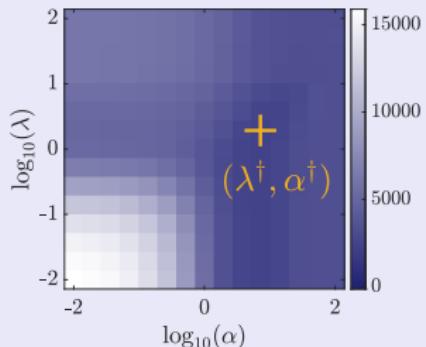
$$\partial_{\boldsymbol{\Lambda}} R_{\boldsymbol{\Pi}}(\boldsymbol{\Lambda}) = \lim_{\nu \rightarrow 0} \mathbb{E}_{\zeta, \varepsilon} \left[\partial_{\boldsymbol{\Lambda}} \widehat{R}_{\nu, \varepsilon}(\mathbf{y}; \boldsymbol{\Lambda} | \mathcal{S}) \right]$$

Parameter tuning (Automatic selection)

$$\left(\hat{\mathbf{h}}, \hat{\mathbf{v}} \right) (\mathcal{L}; \lambda, \alpha) = \underset{\mathbf{h}, \mathbf{v}}{\operatorname{argmin}} \sum_a \| \log \mathcal{L}_{a,.} - \log(a) \mathbf{h} - \mathbf{v} \|^2 + \lambda \mathcal{Q}(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)$$

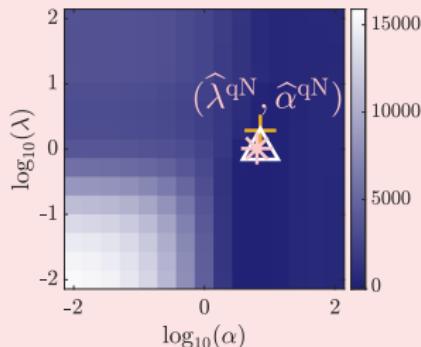
$\bar{\mathbf{h}}$: true regularity

$$\mathcal{R}(\lambda, \alpha) = \left\| \hat{\mathbf{h}}(\mathcal{L}; \lambda, \alpha) - \bar{\mathbf{h}} \right\|^2$$



$\bar{\mathbf{h}}$: unknown!

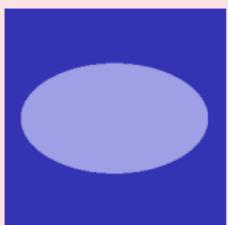
$$\widehat{R}_{\nu, \epsilon}(\mathcal{L}; \lambda, \alpha | \mathcal{S})$$



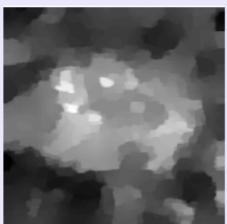
Automated selection of regularization parameters

$$\left(\hat{\mathbf{h}}^F, \hat{\mathbf{v}}^F \right) (\mathcal{L}; \boldsymbol{\Lambda}) = \operatorname{argmin}_{\mathbf{h}, \mathbf{v}} \sum_a \| \log \mathcal{L}_{a,.} - \log(a) \mathbf{h} - \mathbf{v} \|^2 + \lambda Q_F(\mathbf{D}\mathbf{h}, \mathbf{D}\mathbf{v}; \alpha)$$

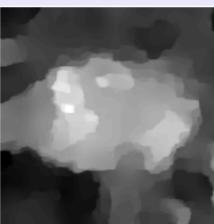
Example



$\hat{\mathbf{h}}^F(\mathcal{L}; \lambda^\dagger, \alpha^\dagger)$
(grid)



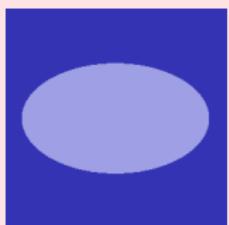
$\hat{\mathbf{h}}^F(\mathcal{L}; \hat{\lambda}^\dagger, \hat{\alpha}^\dagger)$
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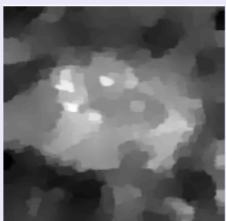
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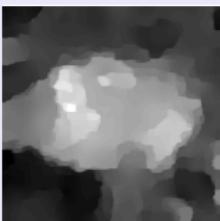
Example



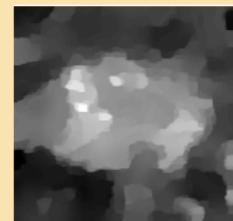
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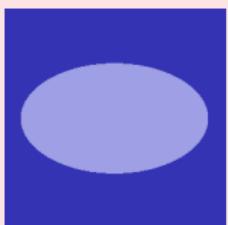
$\hat{\mathbf{h}}^F(\mathcal{L}; \hat{\lambda}^{qN}, \hat{\alpha}^{qN})$
(quasi-Newton)



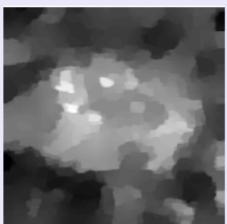
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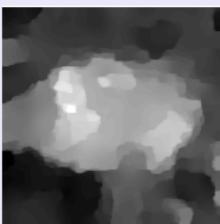
Example



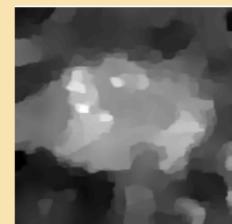
$\hat{\mathbf{h}}^F(\mathcal{L}; \lambda^\dagger, \alpha^\dagger)$
(grid)



$\hat{\mathbf{h}}^F(\mathcal{L}; \hat{\lambda}^\dagger, \hat{\alpha}^\dagger)$
(grid)



$\hat{\mathbf{h}}^F(\mathcal{L}; \hat{\lambda}^{qN}, \hat{\alpha}^{qN})$
(quasi-Newton)



40 calls of the estimator v.s. 225 over a grid

Introduction
○○○○

Texture characterization
○○○○

Design of functionals
○○

Accelerated minimization algorithms
○○○○○○○○○○○○○○○○

Hyperparameters tuning
○○○○○○○○○○○○○○○○

Conclusion
●

Take home message

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 - ▶ ensured objectivity and reproducibility
- In progress: automated analysis of temporal series of images from multiphase flow experiments [Ann. Telecom, 2020]

